

## CANADA'S ALCOHOL AND OTHER DRUGS SURVEY 1994 : A DISCUSSION OF THE FINDINGS



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A nalyse des résultats

# CANADA'S ALCOHOL AND OTHER DRUGS SURVEY 1994: <br> A DISCUSSION OF THE FINDINGS 

Prepared for the<br>Office of Alcohol, Drugs and Dependency Issues<br>Health Canada<br>by<br>Patricia MacNeil<br>and<br>Ikuko Webster

## PREFACE

Canada's Alcohol and Other Drugs Survey (CADS) 1994: A Discussion of the Findings is a detailed look at Canadians' behaviours and attitudes around alcohol and other drugs as revealed in the second and most recent national survey conducted under the research arm of Phase II of Canada's Drug Strategy, a collaborative endeavour of federal, provincial and territorial governments, and many non-governmental organizations. The publication both updates and expands on data gathered in the 1989 National Alcohol and Other Drugs Survey, and also reflects the increased emphasis in CADS on applied research, an issue of particular importance as financial resources decline. It will be useful to people in the field of alcohol and other drugs, and related health and social fields, and of particular significance to policy makers, scientists, and treatment and program specialists.

Canada's Alcohol and Other Drugs Survey: Preview 1995, a point-form review of findings in CADS published in the fall of 1995, supplements and supports material in this current publication. A comprehensive review of available knowledge on alcohol and other drug use, and related health and social problems in Canada, with special attention to emerging developments in all provinces and territories, appears in another earlier research publication, Horizons 1994: Alcohol and Other Drug Use in Canada.

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The research conducted on behalf of Canada's Drug Strategy (CDS) between 1993 and 1996 benefited from the co-operation and hard work of many dedicated people. Special thanks are extended to our federal partners in Canada's Drug Strategy, our colleagues in the provincial and territorial departments of health and addictions agencies, the regional Health Promotion Research Centres, and the many non-governmental organizations that were involved in the research activities.

Canada's Alcohol and Other Drugs Survey 1994: A Discussion of the Findings is a final product of a three-year research effort. It would not have been possible without the assistance of many people, including the staff of the Office of Alcohol, Drugs and Dependency Issues, who were extremely patient and supportive of the process. Indeed, the spirit of "trying new methods and applications" endorsed by CDS has helped to make the research more responsive and meaningful to many people.

Several groups made significant contributions to this project: Dalhousie University had a leading role in the questionnaire development and consultation process; the Canadian Centre on Substance Abuse, Carleton University, and the Addiction Research Foundation also contributed a great deal, not only to this project, but to various other research activities conducted between 1993 and 1996. To
assist in the technical development of research projects, a national research advisory team provided insight and guidance in specific areas.

Finally, we acknowledge some individuals: Anne MacLennan for her editing style, Diane Jacovella and Ziggie Malyniwsky for their faith in and support of research, Jim Anderson and Pete Conley for reviewing the draft material, and so many others who lent their support and encouragement - thank you.

## H I G H L I G H T S

Research is critical to the development of health policies and programs. However, as budgets shrink and the need for accountability increases, there is growing concern that demand - and support - for research will diminish, unless it is responsive, accessible and comprehensible to those who need it.

Each stage of the development of Canada's Alcohol and Other Drugs Survey (CADS), and of the presentation of its findings, has reflected this emerging reality, building bridges between research and development, and the practice that must be based on it. Consultation - and compromise - have been paramount.

Preview 1995, an overview of CADS findings, responded to requests from a wide range of people working in health and social fields for timely, easy-to-read updates on national and provincial attitudes and behaviours around alcohol and other drugs.

This publication, in which a statistical technique called multivariate analysis is brought to bear on the findings, allows for more rigorous analyses of the many independent variables that may influence particular alcohol and other drug-use behaviours, and indeed of the interrelatedness of some of the variables themselves. The clearer explication of the role of various predictors in alcohol and other drug use allows, in turn, for more informed decision-making and policy formulation.

Since 1979, there has been an overall and continuing downward trend in the proportion of Canadians reporting alcohol consumption. At the same time, drinking patterns in the population vary. Multivariate analyses identified socio-demographic characteristics that are independently associated with six drinking patterns.

- Lifetime Abstainers: Women and people 55 years and older are more likely to be lifetime abstainers, as are people living in Atlantic Canada, those speaking a language other than French or English at home, and/or people reporting lower income.
- Former Drinkers: Women and people 45 years and older are more likely to be former drinkers, as are people who have not completed secondary school education. Compared to the Canada-wide rate, Quebec has a lower rate and British Columbia a higher rate of former drinkers.
- Light/Infrequent Drinkers: Light/infrequent drinkers have alcohol less than once a week, and average fewer than five drinks on days they drink. Characteristics associated with this drinking pattern are: being female, being in the youngest age group (15 to 17), residing in the Prairie provinces, speaking a Ianguage other than English or French, being married, and/or reporting lower income.
- Light/Frequent Drinkers: Light/frequent drinkers have alcohol once a week or more, and fewer than five drinks on days they drink. Characteristics associated with a light/frequent drinking pattern are: being male, being 35 and older, living in Quebec or Ontario, having a university degree, and/or reporting a high income.
- Heavy/Infrequent Drinkers: Heavy/infrequent drinkers use alcohol less than once a week, but have five or more drinks, when they do drink. Characteristics associated with this pattern are: being male, being between 15 and 34 (and especially between 15 and 19), living in Atlantic Canada and/or not having completed secondary school.
- Heavy/Frequent Drinkers: Heavy/frequent drinkers have alcohol once a week or more, and five or more drinks on days they drink. Characteristics associated with this pattern of drinking are: being male, belonging to age groups between 18 and 44 (and particularly between 18 and 24), living in Atlantic Canada, being single or never married, having less than secondary school education, and/or a lower income.

In the past 30 years, there has been a decreasing trend in the proportions of current

Tobacco mines or heroin are small and appear to have changed little from 1989 to 1994 However, the prevalence of the use of cannabis has fluctuated over the past five years. The proportion of Canadians reporting the use of cannabis was $6.5 \%$ in 1989, $5 \%$ in 1990, $4.2 \%$ in 1993 and $7.4 \%$ in 1994.

Illicit drug-use consists primarily of the use of cannabis. Whereas about $7.4 \%$ of Canadians report having used cannabis in the 12 months prior to the survey, less than
one percent of the general population report use of cocaine/crack, LSD, amphetamines or heroin. Illicit drug-use is a behaviour found primarily among youth and twice as often among males.


## CHAPTER 1

# OBJECTIVES AND METHODOLOGY 

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Canada's Alcohol and Other Drugs Survey 1994, (CADS) is the second national survey to focus on alcohol and other drug-use in Canada. It was conducted under the research arm of Canada's Drug Strategy, Phase II. CADS updates and expands on data collected in the first National Alcohol and Other Drugs Survey (NADS) conducted in 1989.

Survey questions were developed by research, policy and program representatives from all of the provinces and territories in association with their federal partners in Canada's Drug Strategy. Although the 1989 survey instrument forms its base, CADS was expanded to accommodate emerging policy, program and research needs among partners; to improve data quality; and to better reflect the current Canadian population. Questions were added to capture data on new issues and changes in drug types since 1989, and others were added, revised or deleted to make the survey more relevant and sensitive to gender and cultural issues. The collaborative nature of the development process is reflected also in the new emphasis on policy.

A point-form review of some of the major findings in CADS, which supports and supplements this current publication, appears in Canada's Alcohol and other Drugs Survey: Preview 1995, published in the fall of 1995.

As agreed by the working groups, the main survey objectives were to:

- measure the prevalence and patterns of alcohol and other drug use in Canada;
- assess related harm;
- evaluate trends;
- measure demographic, contextual or proximal risk factors of use; and
- assess the range of responses to problems, including attitudes towards users and problem behaviours

The survey was carried out by Statistics Canada during the period of September 7 to November 5, 1994, using a Random Digit Dialling (RDD) telephone samplingmethod. ${ }^{1}$

The target population was people 15 years of age and older in Canada, excluding residents of the Yukon and Northwest Territories and full-time residents of institutions. A second phase of CADS was conducted in the two territories in 1995 and 1996.

Objectives

Survey M ethod

Population Coverage

With RDD, households without telephones were also excluded. However, as this group represents less than $2 \%$ of the target population, survey estimates have been adjusted (weighted) to represent them.

To carry out sampling, each of the 10 provinces was divided into strata or geographic areas. Generally, for each province, one stratum represented the Census Metropolitan Areas (CMAs) of the province, and another represented the non-CMAs. Exceptions were Prince Edward Island, with no CMA and hence no CMA stratum, and Montreal and Toronto, each with separate strata.

The method known as Elimination of Non-Working Banks (ENWB) was used in which an attempt is made to identify all working banks for an area (i.e. to identify all banks with at least one household). Thus, all telephone numbers within nonworking banks are eliminated from the sampling frame.

For each province, lists of telephone numbers in use were purchased from the telephone companies, and lists of working banks were extracted. Each bank was assigned to a stratum within its province. A random sample of telephone numbers was generated in each stratum. On contact with a household, all household members were listed, and basic demographic information was collected on age, sex and marital status. One person 15 years or older was randomly selected in each household, and his/her relationship to all other household members was assessed. For each selected person, a CADS interview was executed.

Responses were obtained from 12,155 of 16,082 selected households, yielding a $75.6 \%$ response rate. Sample sizes and response rates for each province are listed in Appendix B.

Of the 2,939 (18.3\%) non-responding households, 1,666 were refusals, 789 could not be reached for the entire survey period ("ring-no-answer" households), and 484 were cases of language difficulties or other problems. From a total of 13,143 responding households, 12,155 useable responses were obtained.

Non-respondents are more likely to be males and younger ${ }^{1}$. In the responding sample, $3.2 \%$ were males between the ages of 15 and 19, while, in the overall population, approximately $4.4 \%$ are males in this age group. Thus, the sample counts cannot be considered representative of the target population, unless appropriate weights are applied.

CATI, the survey questions appear on a computer monitor, and the interviewer asks the questions and enters the responses into the computer as the interview progresses. Built-in edits and fewer processing steps provide for more efficient and accurate data collection. CATI methodology also eliminates the need for paper questionnaires. A paper document has been produced, but has not been included in this report because of its excessive length. For copies of the complete instrument, contact Health Canada's Office of Alcohol, Drugs and Dependency Issues.

A field test of the questionnaire was completed in J une 1994 in the Statistics Canada Halifax and Montreal regional offices. Data collection for the main survey began the second week of September 1994 and continued through the first week of November 1994. All interviewing took place using CATI telephone facilities in five regional offices of Statistics Canada: Halifax, Montreal, Toronto, Edmonton and Vancouver. All interviewers, most of whom had previous experience, were trained by Statistics Canada staff in telephone interviewing-techniques, and survey concepts and procedures. Interviews were conducted between 9:00 and 21:30 hours from Monday through Saturday for the survey period.

Responses to the survey questions were entered directly into the CATI mini-computers in Statistics Canada's regional offices and transmitted to Ottawa. The data-capture program allowed for a valid range of codes for each question and automatically followed the flow of the questionnaire for CADS.

The expansion weight (FINWGHT) included in Statistics Canada's microdata file was used to obtain population estimates of counts and percentages. The value of FINWGHT indicates how many individuals in the population are represented by each respondent in the sample, and weighting by this variable ensures that the estimates are representative of the Canadian population at the time of the survey September 1994. This variable was derived from the inclusion probability of telephone numbers in each stratum, adjusted for non-response, multiple telephones per household, the number of eligible persons per household, and for the population projection in each of the provinces by age and sex groupings.

Weighting by FINWGHT expands the total count from a sample size of 12,155 to the estimated population of $23,029,739$. The average FINWGHT is: $23,029,739$ / $12,155=1,894.7$ ). This means that, on average, one respondent represents almost 1,900 people.

Multivariate Analysis Using Logistic Regression

Weighting by FINWGHT cannot be used for analyses that involve statistical inference because variance estimates based on the population count are meaningless. Instead, FINWGHT was rescaled by dividing by the average FINWGHT $(1,894.7)$. This produced a second weighting variable (WT12155) that retains the sample size of 12,155 and ensures the representativeness of the results at the same time ${ }^{\text {a }}$.

CADS used a complex sampling-design involving stratification and multistage selection, rather than simple random-sampling, which is assumed in the statistical procedures in standard packages involving significance testing. A complex sampling design has indeed resulted in loss of precision (i.e. a higher variance estimate) in CADS, as indicated by the fact that the average design effect presented in Appendix B is greater than one in every geographic unit. Therefore, use of standard packages with CADS data weighted by WT12155 will result in underestimation of the variance, leading to an inflated probability of declaring effects or differences significant, when they are not.

Experts in the field ${ }^{2,3}$ strongly recommend the use of specialized statistical packages that have been programmed to incorporate complex sampling-designs, but often the requisite information on sampling design such as the stratum-identification variable is suppressed to protect confidentiality. CADS is no exception, and the only alternative ${ }^{4}$ is to further scale down WT12155 by dividing by the average design effect for each province. The total count $(10,530)$ produced through the use of this new weighting variable, ESSPROV ${ }^{\text {b }}$ is called the effective sample size, which is substantially smaller than the original sample size. This down-sizing is intended to compensate for the underestimated variance.

In assessing the association between a dependent variable and two or more independent variables (predictors), it is imperative to go beyond looking at one independent variable in isolation; i.e. a series of two-way cross-tabulation results. This is because independent variables themselves are very often interrelated to varying degrees.

[^0]The effects of region and language on tobacco use provide an example (see Chapter 3). Compared to the Canada-wide statistic ( $27.4 \%$ ), the prevalence of current smokers is higher in Quebec (33.6\%) and also among francophones (33.3\%). Given that the majority of francophones are in Quebec, are region and language both contributing factors or is one of them the primary factor with the other an artifact?

A three-way cross-tabulation, including the dependent and both of the independent variables is sometimes helpful, as it turns out to be in this case (see Table T2). It illustrates that, in four of the five regions, including Quebec, prevalence of current smoking among francophones appears to be somewhat lower than among anglophones. Residency in Quebec is the main contributing factor involved, and the high prevalence found among francophones is its artifact.

Having established that language is not a factor contributing to a higher prevalence in Quebec, the next question might be whether education and income level are involved. Rather than proceeding with more three-way or possibly fourway cross-tabulations, a multivariate technique allows assessment of the unique or independent association or effect of each predictor, controlling for all of the other predictors included in the analysis.

As the dependent variable is dichotomous, the method of choice is logistic regression, the name of which comes from the use of "logit" or logarithmically transformed "odds" as the dependent variable. The Canada-wide prevalence of current smokers $(27.4 \%$ ) is translated to odds of $.274 /(1-274)=.377$, while Quebec's $33.6 \%$ yields odds of $.336 /(1-.336)=506$. Quebec's odds ratio relative to the Canadawide ratio is given by $.506 / .377=1.34$. This is presented in Table T3 under "Unadjusted odds ratio" i.e. observed odds ratio without taking the other predictors into consideration. An odds ratio greater than 1.0 indicates a greater than average odds, while an odds ratio less than 1.0 indicates a smaller than average odds for being a current smoker.

What logistic regression does is build an equation that explains or predicts the "logit" by combining all the independent variables, each of which is weighted by the best possible coefficient so that the composite or predicted values of "logit" for all the respondents will agree most closely with the observed data. If a coefficient is significant (reliably different from 0 ), it can be interpreted in terms of the direction and the size of the independent association or effect the predictor has on "logit".
© Throughout this report, small discrepancies will be noted when percentages in a table presenting multivariate results weighted by ESSPROV (e.g. Table T3) are compared to the percentages obtained with FINWGHT as the weighting variable (e.g. Table T1). They are due oo the difference in weighting and exclusion of the "not stated" categories of the dependent variable and the predictors from the multivariate analysis. The "not stated" categories of the predictors did not exceed 2.1\% overall, weighted by FINWGHT.

Sampling Variability and
Data-Release Criteria

For ease of interpretation, however, the coefficient is often translated back to the form that applies to the odds ratio. For example, the coefficient for Quebec is .3385, and its exponential transformation 1.403 is presented in Table T3 under "Adjusted odds ratio". "Adjusted" refers to the fact that the effects of all of the other predictors are controlled. With every factor being equal, the odds ratio for Quebec are estima-ted to be 1.4 times as high as the Canada-wide odds. This difference is statistically reliable as indicated by **. The francophone effect, on the other hand, is not significant after adjustment for all other predictors is made. In fact, anglophone odds are significantly higher than the overall odds, the odds ratio being 1.2. Multivariate analysis can sometimes unmask effects hidden by certain associations between the predictors, in this case the association between region and language.

As evident in Table T3, at least one category of each predictor is making a significant independent contribution to explaining/predicting the dependent variable when adjusted for all the other predictors. The strength of that significant contribution can be judged by its adjusted odds ratio. For odds ratios greater than one, the higher the ratio, the stronger the contribution, whereas the opposite holds for odds ratios that are smaller than one. The highest significant odds ratio is associated with 'age 20-24', and the lowest with 'age $75+$ '.

Although they are not included in the tables of this report, logistic regression output from standard statistical packages such as SPSS, BMDP and SAS provide statistics that enable a similar evaluation of each predictor comprised of several categories. For example, 'educational level' is judged to be the strongest among the demographic predictors included in the analysis of current smoking, despite the fact that categories with the highest and lowest odds ratios both belong to the 'age' variable. This means that removing 'educational level' from the logistic regression equation will amount to the greatest loss in terms of predicting or explaining the dependent variable. The discrepancy may be due to a fairly uniform distribution of individuals across four major categories of 'educational level', in contrast to an uneven distribution of individuals across age categories. ${ }^{5}$

The difference between the population estimates obtained from the sample and the results from a complete population count (census) is called the "sampling error" of the estimates. Although the exact sampling error cannot be measured from the sample results alone, it is possible to estimate it from the sample data; this is the standard error of the estimates. Because of the large number of estimates that can be derived from a survey of this size, the standard error is usually expressed
relative to the estimate to which it pertains. This is known as the coefficient of variation of the estimate (CV) and is expressed as a percentage of the estimate. The CADS Microdata User's Guide includes a CV table for each of the 13 geographic units within Canada. (Canada, 10 provinces, and two territories). These tables include the design effect resulting from the complex sampling-design.

Statistics Canada's data-release guidelines recommend that any estimate with a CV exceeding $33.3 \%$ not be released. This guideline has been followed with the exception of the use of " $<x \%$ " to substitute for suppressed estimates, where $x \%$ is the upper limit of the $99 \%$ confidence interval.

Significance testing involving $t$-tests for comparisons of two estimated percentages (e.g. current smokers among men and women) followed the methods outlined in Statistics Canada's "User's Guide." A confidence level of $99 \%$, or equivalently a significance criterion of .01 , was used in each analysis to avoid significant, yet trivial effects.

1 Statistics Canada. (1994). Canada's Alcohol and Other Drugs Survey: Microdata User's Guide. Ottawa: Statistics Canada

2 Särndal, C-E., Swensson, B., and Wretman, J. (1992). M odel Assisted Survey Sampling. New York: Wiley.

3 Skinner, C.J., Holt, D., and Smith, T.M .F. (eds.) (1989). Analysis of Complex Surveys. New York: Wiley.

4 Lee, E.S., Forthofer, R.N., and Lorimor, R.J. Analyzing Complex Survey Data (Series: Quantitative Applications in the Social Sciences, No. 71). Beverley Hills and London: Sage Publications.

5 Hosmer, D.W. Jr. and Lemeshow, S. (1989). Applied Logistic Regression. New York: Wiley.

## CHAPTER 2

## ALCOHOL

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To distinguish high- from low-risk modes of alcohol use, it is useful to identify patterns of use. The following discussion presents two patterns of abstinence and four patterns of alcohol consumption. $1^{1,2}$ A bstainer categories differentiate between respondents who have never used alcohol and those who are former drinkers. Table A 1 defines the various drinking patterns and compares the proportions in CADS with those in the 1989 National Alcohol and Other Drugs Survey. ${ }^{\text {a }}$ The classifications of drinkers are based on responses to questions about usual frequency of alcohol use and average number of drinks on days they drank.

While the lighter/heavier and infrequent/frequent drinker divisions tend to be arbitrary, they nevertheless provide indications of the volume of alcohol consumed, and thus of the potential for intoxication and likelihood of harm as a result of high intake, either to others through disturbing behaviour or to self. The discussion in this section will verify that abstention and use patterns are differentially distributed in the population, allowing, in turn, for programs and policies to be designed specifically for those at risk of harm.

Comparisons of percentages of the drinker types for the two survey years reveal little change overall in patterns of alcohol use. There are, however, some substantial differences between the extreme types. The percentage of current nondrinkers (lifetime abstainers and former drinkers) in the population has increased (to $26.3 \%$ from $22.3 \%$ ) between 1989 and 1994. The apparent increase in the percentage of lifetime abstainers (to $12.8 \%$ from $6.6 \%$ ) is difficult to interpret due to the changes in the questions concerning whether the respondent had ever used any alcohol. ${ }^{\text {b }}$ There is also a decrease in the rate of heavy/frequent drinkers, (to $5.4 \%$ from $6.7 \%$ ) an encouraging finding for this group at high risk for negative medical and social consequences of drinking.

More than a quarter of the population 15 years and older (26.3\%) does not use alcohol, and most current drinkers usually have fewer than five drinks per occasion. Among lighter drinkers, the rate of infrequent users is higher than frequent users ( $33.6 \%$ vs. $29.2 \%$ ). For the heavier drinkers, or those who report usually having five or more drinks, when they use alcohol, the situation is reversed: the proportion of frequent/heavy drinkers is higher than the proportion of infrequent/heavy drinkers (5.4\% vs. 3.3\% of the population).

Table A2 presents the drinking pattern categories in terms of socio-demographic predictors. The rate of current abstention is higher for females than for males ( $32.3 \%$ vs. $20.1 \%$ ).

[^1]Table A1
Drinking patterns of Canadians: NADS, 1989 and CADS, 1994

Drinking patterns by sex, a language, marital status, education and income

| Pattern |  | Percent |  |
| :---: | :---: | :---: | :---: |
|  | Definition | $\begin{gathered} \text { NADS } \\ 1989 \end{gathered}$ | $\begin{aligned} & \text { CADS } \\ & 1994 \end{aligned}$ |
| Lifetime abstainers | Never had alcohol beyond sips or tastes. | 6.6 | 12.8 |
| Former drinkers | Drank sometime during their lives, but not during the past 12 months preceding the survey. | 15.7 | 13.5 |
| Light/infrequent drinkers | Drink less often than once a week, usually fewer than five drinks, when alcohol is used. | 35.5 | 33.6 |
| Light/frequent drinkers | Drink once a week or more, usually fewer than five drinks, when alcohol is used. | 31.3 | 29.2 |
| Heavy/infrequent drinkers | Drink less often than once a week, usually five or more drinks, when alcohol is used. | 3.6 | 3.3 |
| Heavy/frequent drinkers | Drink once a week or more, usually five or more drinks, when alcohol is used. | 6.7 | 5.4 |
| Not stated |  | . 6 | 2.1 |

Note: Weighted by FINWGHT in CADS, and WEIGT in NADS.

A mong current drinkers, the most prominent category for women is light/infrequent, while, for men, it is light/frequent drinkers. Men are far more likely than women to be heavier drinkers

|  | Pop. est. (000s) | Drinking pattern (percent) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Lifetime abstainers | Former drinkers | Light infrequent | Light frequent | Heavy infrequen | Heavy frequent |
| Overail | 23,030 | 12.8 | 13.5 | 33.6 | 29.2 | 3.3 | 5.4 |
| Sex |  |  |  |  |  |  |  |
| Male | 11,337 | 8.9 | 11.2 | 27.3 | 36.4 | 4.5 | 9.1 |
| Female | 11,692 | 16.7 | 15.6 | 39.8 | 22.2 | 2.2 | 1.9 |
| Age |  |  |  |  |  |  |  |
| 15-17 | 1,247 | 21.8 | 12.8 | 33.8 | $7.1{ }^{\text {Q }}$ | 12.5 | 10.0 |
| 18-19 | 711 | 11.50 | 9.00 | 38.6 | 16.1 | 10.30 | 14.0 |
| 20-24 | 2,051 | 8.0 | 7.1 | 37.1 | 26.8 | 6.4 | 12.9 |
| 25-34 | 4,952 | 8.6 | 10.4 | 39.1 | 29.9 | 4.2 | 6.3 |
| 35-44 | 4,802 | 9.2 | 11.7 | 34.7 | 34.8 | 2.4 | 5.0 |
| 45-54 | 3,531 | 10.8 | 14.7 | 33.7 | 33.6 | 1.40 | 3.7 |
| 55-64 | 2,470 | 16.7 | 16.8 | 27.7 | 32.7 | -- | 2.30 |
| 65-74 | 2,195 | 21.6 | 21.2 | 25.5 | 27.6 | -- | -- |
| 75+ | 1,071 | 28.4 | 24.0 | 23.8 | 20.5 | -- | -- |
| Region |  |  |  |  |  |  |  |
| Atlantic | 1,907 | 15.5 | 14.1 | 35.4 | 18.9 | 7.0 | 8.4 |
| Quebec | 5,796 | 14.8 | 11.0 | 32.1 | 33.3 | 2.5 | 5.8 |
| Ontario | 8,673 | 14.3 | 13.1 | 31.7 | 29.5 | 2.7 | 4.4 |
| Prairies | 3,715 | 9.6 | 15.1 | 38.3 | 25.8 | 4.2 | 6.2 |
| B.C. | 2,939 | 7.0 | 16.8 | 35.4 | 31.0 | 3.1 | 5.1 |
| Language |  |  |  |  |  |  |  |
| English | 15,006 | 9.9 | 14.5 | 35.4 | 29.5 | 3.8 | 5.9 |
| French | 5,170 | 13.8 | 12.0 | 32.9 | 32.8 | 2.6 | 5.6 |
| Other | 1,452 | 31.8 | 11.3 | 30.4 | 21.0 | 2.50 | 2.00 |
| Not stated | 1,402 | 20.9 | 9.8 | 20.3 | 21.2 | -- | 3.2 |
| M arital status |  |  |  |  |  |  |  |
| M arried/common-law | 13,564 | 11.3 | 13.9 | 34.8 | 32.5 | 2.2 | 3.2 |
| Single/never married | 6,317 | 12.3 | 10.6 | 34.1 | 23.9 | 6.5 | 11.1 |
| Widowed | 1,316 | 29.7 | 19.8 | 27.9 | 19.2 | -- | -- |
| Divorced/separated | 1,587 | 12.5 | 15.9 | 29.4 | 31.9 | 2.30 | 6.3 |
| Not stated | 246 | 23.50 | -- | 16.50 | 15.30 | -- | -- |


|  | Pop. est. (000s) | Drinking pattern (percent) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Lifetime abstainers | Former drinkers | Light infrequent | Light frequent | Heavy infrequent | Heavy frequent |
| Educational level |  |  |  |  |  |  |  |
| Less than secondary | 5,936 | 18.6 | 20.4 | 29.8 | 18.5 | 5.0 | 6.7 |
| Secondary | 5,415 | 11.0 | 12.9 | 36.7 | 28.6 | 3.7 | 6.2 |
| Some post-secondary | 6,455 | 9.2 | 10.8 | 37.1 | 33.5 | 2.9 | 6.1 |
| University degree | 3,610 | 8.6 | 9.0 | 34.8 | 43.4 | 1.50 | 2.30 |
| Not stated | 1,614 | 21.7 | 10.5 | 20.9 | 21.2 | -- | 3.00 |
| Income |  |  |  |  |  |  |  |
| Low | 3,612 | 16.4 | 17.8 | 35.4 | 17.4 | 4.8 | 7.4 |
| Middle | 7,742 | 8.6 | 13.1 | 37.5 | 32.0 | 3.1 | 5.3 |
| High | 2,778 | 5.0 | 8.9 | 30.8 | 47.2 | 2.70 | 5.0 |
| Not stated | 8,898 | 17.5 | 13.4 | 30.4 | 25.9 | 3.0 | 4.9 |

Note: Frequent drinking = once a week or more. Heavy drinking = usual number is 5 or more. 'Not stated' category of drinking pattern ( $2.1 \%$ overall) is not presented in the table, although it was included in the computation of the percentages.
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability
Lifetime abstainers tend to be concentrated at the youngest and oldest ends of the age spectrum, while rates of former drinkers increase with age. Rates for light/frequent drinking plateaus at 35 to 64 years of age, while heavier drinking rates are concentrated among those 24 and younger.

Drinking patterns vary among regions of the country. The Atlantic provinces have the highest rates of both lifetime abstainers and heavy drinkers, suggesting that alcohol use may be more of a contentious issue there than in the rest of Canada. The Atlantic provinces also have a much lower rate of light/frequent drinkers. Quebec has the lowest proportion of former drinkers.

The highest rates of lifetime abstention occur among those whose language at home is neither English nor French, indicating some resistance to alcohol use among recent immigrants. Heavier drinking rates are highest among those whose language at home is English.

Those who are widowed have relatively high rates of lifetime abstention, while single people tend to have higher rates of heavier drinking. The rate for heavy/frequent drinking among those who are separated or divorced is twice as high as for those who have partners (married/common-law).

Both extremes - abstention and heavier drinking - occur with higher rates than the overall among those with the lowest level of education, while the rate of light/frequent drinking increases with level of education. High abstention rates also occur among those with lower incomes, and relatively high heavy drinking rates appear in low income categories, as well.

To identify those members of the Canadian population who may be at risk of harm associated with alcohol, it is useful to profile drinker types by their dominant characteristics. Tables A3 to A8 show the relationship between alcohol-use patterns and socio-demographic predictors. The column labeled "unadjusted odds ratio" provides much the same information as the percentage tables, showing (for example) that women (with odds ratio of 1.354 ) are more than twice as likely as men (.668) to report lifetime abstention. Thus, the unadjusted odds ratio is an indication of the extent to which people with a given characteristic will report abstention or the drinking pattern in question as compared to the overall odds.

The adjusted odds ratio is the result of a multivariate analysis (logistic regression). This analysis assesses the independent contribution of each predictor category to the outcome, controlling for all of the other variables. Using lifetime abstention as an example again, the unadjusted odds among those reporting being widowed is high (2.949) and significant, indicating that widows are almost three times more likely than others to abstain from alcohol. Controlling for all other factors but widowhood, however, reduces the odds ratio to almost unity (1.167), and the statistic is not significant. This outcome suggests that widowhood itself does not impact on abstention once the effects of other characteristics associated with widowhood (e.g. being female and being older) have been removed.

Note that certain factors are associated with decreases in the likelihood of an outcome. For example, the odds of both greater frequency and greater volume of alcohol use are significantly decreased, when the respondent is a woman (Tables A6, A7, and A8).

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overall | 10235 | 12.9 |  |  |
| Sex |  |  |  |  |
| Male | 5030 | 9.0 | .668** | .710** |
| Female | 5206 | 16.7 | $1.354 * *$ | 1.409** |
| Age |  |  |  |  |
| 15-17 | 559 | 22.2 | 1.927** | 1.349** |
| 18-19 | 322 | 11.2 | . 852 | . 711 |
| 20-24 | 921 | 8.1 | .595** | .557** |
| 25-34 | 2220 | 8.5 | . 627 7* | . $642^{* *}$ |
| 35-44 | 2126 | 9.1 | .676** | .718** |
| 45-54 | 1568 | 11.1 | . 843 | . 867 |
| 55-64 | 1083 | 16.8 | 1.363** | 1.308* |
| $65-74$ | 963 | 22.1 | $1.916 * *$ | 1.535** |
| $75+$ | 473 | 29.4 | 2.812** | $2.334 * *$ |
| Region |  |  |  |  |
| Atlantic | 895 | 15.6 | 1.248 | 1.596** |
| Quebec | 2711 | 14.9 | 1.182 | 1.140 |


| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Ontario | 3595 | 14.6 | $1.154 * *$ | 1.128 |
| Prairies | 1701 | 9.6 | .717** | . 830 |
| B.C. | 1333 | 7.1 | . 516 ** | . 587 ** |
| Language |  |  |  |  |
| English | 6714 | 9.9 | .742** | .589** |
| French | 2418 | 13.9 | 1.090 | . 779 |
| Other | 650 | 32.0 | $3.177 \times *$ | $3.021 * *$ |
| Not stated | 454 | 25.4 | 2.299** | . 721 |
| M arital status |  |  |  |  |
| M arried/common-law | 6086 | 11.5 | .877** | . 874 |
| Single/never married | 2845 | 12.4 | . 956 | . 998 |
| Widowed | 591 | 30.4 | $2.949 * *$ | 1.167 |
| Divorced/separated | 714 | 12.7 | . 982 | .983 |
| Educational level |  |  |  |  |
| Less than secondary | 2684 | 18.8 | 1.563** | 1.100 |
| Secondary | 2451 | 11.0 | . 835 | . 862 |
| Some post-secondary | 2937 | 9.2 | . $684 \times$ | .785* |
| University degree | 1632 | 8.5 | . 627 7* | .753* |
| Not stated | 532 | 26.1 | 2.385** | 1.784* |
|  |  |  |  |  |
| Income |  |  |  |  |
| Low | 1647 | 16.5 | 1.334** | 1.332** |
| Middle | 3546 | 8.6 | . $635 *$ | .848* |
| High | 1258 | 5.0 | . 355 ** | . 626 ** |
| Not stated | 3784 | 18.1 | 1.492** | 1.413** |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size.
*p<.01; ** $p<.001$

Women and older people have higher odds of being lifetime abstainers. Living in Atlantic Canada, speaking a language other than French or English at home, and reporting lower income are all independently and significantly associated with increased odds of never having used alcohol. Associated with decreased odds of lifetime abstention are the following characteristics: being male, being between 20 and 44 years of age, speaking English at home, living in British Columbia and reporting a high educational level. As mentioned above, marital-status categories are no longer significant in multivariate analysis.

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overall | 10235 | 13.7 |  |  |
| Sex |  |  |  |  |
| Male | 5030 | 11.5 | .819** | .845** |
| Female | 5206 | 15.9 | 1.191** | $1.184 * *$ |
| A ge |  |  |  |  |
| 15-17 | 559 | 13.1 | . 950 | . 715 |
| 18-19 | 322 | 8.9 | . 615 | .586* |
| 20-24 | 921 | 7.2 | 489** | .533** |
| 25-34 | 2220 | 10.6 | .747** | . 870 |
| 35-44 | 2126 | 11.8 | 843 | . 965 |
| 45-54 | 1568 | 15.0 | 1.112 | 1.244* |

Table A4
ormer drinkers vs. the rest, by sex, age, region, language, marital status, education and income, with and without other predictors taken into account

Former drinkers vs. the

## Former Drinkers

Table A5 Light/infrequent vs. the rest among current drinkers, by sex, age, region, language, marital status, education and income, with and without other predictors taken into account

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| 55-64 | 1083 | 17.5 | $1.336 * *$ | 1.325* |
| 65-74 | 963 | 21.8 | 1.756 * | 1.688** |
| $75+$ | 473 | 24.7 | $2.066 *$ | $1.919 * *$ |
| Region |  |  |  |  |
| Atlantic | 895 | 14.2 | 1.043 | . 938 |
| Quebec | 2711 | 11.1 | .787** | 673** |
| Ontario | 3595 | 13.8 | 1.008 | 1.049 |
| Prairies | 1701 | 15.2 | 1.129 | 1.127 |
| B.C. | 1333 | 16.9 | 1.281** | 1.341** |
| Language |  |  |  |  |
| English | 6714 | 14.6 | 1.077 | 1.077 |
| French | 2418 | 12.0 | . 859 | 1.207 |
| Other | 650 | 11.4 | 811 | . 862 |
| Not stated | 454 | 13.2 | 958 | 893 |
| M arital status |  |  |  |  |
| M arried/common-law | 6086 | 14.2 | 1.043 | 1.063 |
| Single/never married | 2845 | 10.8 | .763** | 1.072 |
| widowed | 591 | 20.1 | 1.585** | . 807 |
| Divorced/separated | 714 | 16.1 | 1.209 | 1.088 |
| Educational level |  |  |  |  |
| Less than secondary | 2684 | 20.5 | 1.624** | 1.535** |
| Secondary | 2451 | 13.0 | . 941 | . 964 |
| Some post-secondary | 2937 | 10.7 | .755** | . 851 |
| University degree | 1632 | 9.0 | .623** | 700** |
| Not stated | 532 | 14.0 | 1.025 | 1.135 |
| Income |  |  |  |  |
| Low | 1647 | 17.9 | 1.373** | 1.337** |
| Middle | 3546 | 13.1 | . 950 | . 986 |
| High | 1258 | 8.8 | . $608 * *$ | 740** |
| Not stated | 3784 | 14.1 | 1.034 | 1.025 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size. * $p<.01 ; * * p<.001$

There is a direct association of former drinking with age: respondents in younger age groups are less likely to be former drinkers, and those in the older groups are more likely to be former drinkers. Former drinkers, as well as lifetime abstainers, are more prevalent among women. Quebec has a lower, and B.C. a higher rate of former drinkers in comparison to the overall odds. Odds for being a former drinker are inversely related to educational level and income adequacy: odds ratios decrease as educational level and income adequacy increase. Again, the effect of marital status is no longer significant in multivariate analysis.

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overail | 7504 | 47.0 |  |  |
| Sex |  |  |  |  |
| Male | 3994 | 35.4 | .618** | .602** |
| Female | 3509 | 60.2 | 1.706** | 1.662********** |
| Age |  |  |  |  |
| 15-17 | 362 | 53.1 | 1.277 | 1.364* |


| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| 18-19 | 258 | 48.9 | 1.079 | 1.130 |
| 20-24 | 780 | 44.6 | . 908 | . 945 |
| 25-34 | 1796 | 49.2 | 1.092 | 1.088 |
| 35-44 | 1681 | 45.1 | . 926 | 899 |
| 45-54 | 1160 | 46.6 | . 984 | . 979 |
| 55-64 | 712 | 43.9 | . 882 | . 818 |
| 65-74 | 540 | 46.8 | . 992 | . 817 |
| $75+$ | 217 | 53.0 | 1.272 | 1.072 |
| Region |  |  |  |  |
| Atlantic | 628 | 50.8 | 1.164 | 1.133 |
| Quebec | 2007 | 43.6 | .872* | 753** |
| Ontario | 2577 | 46.3 | . 972 | . 969 |
| Prairies | 1279 | 51.5 | 1.197* | 1.188* |
| B.C. | 1013 | 47.4 | 1.016 | 1.018 |
| Language |  |  |  |  |
| English | 5067 | 47.5 | 1.020 | . 886 |
| French | 1791 | 44.5 | . 904 | 1.020 |
| Other | 368 | 54.4 | 1.345* | 1.382* |
| Not stated | 278 | 44.2 | 889 | 801 |
| M arital status |  |  |  |  |
| M arried/common-law | 4518 | 47.9 | 1.037 | 1.140* |
| Single/never married | 2185 | 45.0 | . 923 | . 891 |
| widowed | 292 | 57.5 | $1.526 * *$ | 1.218 |
| Divorced/separated | 508 | 42:0 | .817 | 808* |
| Educational level |  |  |  |  |
| Less than secondary | 1627 | 49.7 | 1.114 | 1.054 |
| Secondary | 1862 | 48.8 | 1.075 | 1.020 |
| Some post-secondary | 2350 | 46.6 | . 984 | . 943 |
| University degree | 1346 | 42.5 | .833** | . 913 |
| Not stated | 319 | 45.1 | 926 | 1.081 |
| Income |  |  |  |  |
| Low | 1080 | 54.4 | 1.345** | 1.313** |
| Middle | 2775 | 48.1 | 1.045 | 1.070 |
| High | 1084 | 36.1 | .637** | .702** |
| Not stated | 2564 | 47.4 | 1.016 | 1.014 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size *p<.01; **p<. 001

Light/infrequent drinkers have alcohol less than once a week, and average fewer than five drinks on days they drink. Characteristics associated with increased adjusted odds for this drinking pattern are: being female, being in the youngest age group (15 to 17), residing in the Prairie provinces, speaking a language other than English or French, being married, and/or reporting lower income. Significantly less likely to drink in this way are males, residents of Quebec, people who are separated or divorced, and/or people who report higher income.

Table A5
Light/infrequent vs.
the rest among current drinkers
(cont'd)

Light/Infrequent Drinkers

Light/frequent vs. the rest able A current drinkers by sex age region language, marital status, education and income, with and without other predictors taken into account

Light/Frequent Drinkers

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overaii | 7504 | 40.7 |  |  |
|  |  |  |  |  |
| Sex |  |  |  |  |
| Male | 3994 | 47.0 | 1.292** | 1.341** |
| Female | 3509 | 33.5 | $734 * *$ | $746 * *$ |
| Age |  |  |  |  |
| 15-17 | 362 | 11.2 | .184** | .265** |
| 18-19 | 258 | 20.4 | . 373 * | 483** |
| 20-24 | 780 | 32.0 | .686** | . 805 |
| 25-34 | 1796 | 37.6 | .878* | . 918 |
| 35-44 | 1681 | 45.3 | 1.207** | 1.260** |
| 45-54 | 1160 | 46.3 | $1.256 * *$ | 1.307 ** |
| 55-64 | 712 | 51.4 | $1.541 *$ | 1.823 ** |
| 65-74 | 540 | 50.9 | 1.510** | $2.050 * *$ |
| 75 | 217 | 45.5 | 1.216 | $1.718 *$ |
| Region |  |  |  |  |
| Atlantic | 628 | 27.1 | .542** | .632** |
| Quebec | 2007 | 45.2 | 1.202** | 1.391** |
| Ontario | 2577 | 43.2 | 1.108* | 1.197** |
| Prairies | 1279 | 34.7 | .774* | . 860 |
| B.C. | 1013 | 41.5 | 1.034 | 1.106 |
| Language |  |  |  |  |
| English | 5067 | 39.4 | . 947 | 1.049 |
| French | 1791 | 44.4 | 1.164* | 1.065 |
| Other | 368 | 37.4 | . 870 | . 832 |
| Not stated | 278 | 46.0 | 1.241 | 1.076 |
| M arital status |  |  |  |  |
| M arried/common-law | 4518 | 44.7 | 1.178** | 1.040 |
| Single/never married | 2185 | 31.5 | . 670 * | . 996 |
| widowed | 292 | 39.4 | . 947 | . 852 |
| Divorced/separated | 508 | 45.5 | 1.216 | 1.134 |
| Educational level |  |  |  |  |
| Less than secondary | 1627 | 30.9 | .652** | .707** |
| Secondary | 1862 | 37.9 | . 889 | . 892 |
| Some post-secondary | 2350 | 42.1 | 1.059 | 1.117 |
| University degree | 1346 | 52.9 | 1.636** | $1.403 * *$ |
| Not stated | 319 | 46.0 | 1.241 | 1.011 |
| Income |  |  |  |  |
| Low | 1080 | 26.8 | .533** | .642** |
| Middle | 2775 | 41.1 | 1.017 | . 998 |
| High | 1084 | 55.0 | $1.781 * *$ | $1.454 * *$ |
| Not stated | 2564 | 40.1 | . 975 | 1.074 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size

* p<.01; ** p<. 001

According to the adjusted odds ratios, characteristics associated with light/frequent drinking patterns are: being male, in age groups 35 and over, living in Quebec or Ontario, having a university degree, and/or reporting a high income. Attributes associated with decreased chances of this drinking pattern are: being female, being 19 or younger, living in Atlantic Canada, not completing secondary school, and/or having a lower income

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overail | 7504 | 4.6 |  |  |
| Sex |  |  |  |  |
| Male | 3994 | 5.8 | 1.277** | 1.354** |
| Female | 3509 | 3.3 | .708** | .738** |
| A ge |  |  |  |  |
| 15-17 | 362 | 19.8 | 5.120** | 6.323** |
| 18-19 | 258 | 13.0 | $3.099 * *$ | $4.634 * *$ |
| 20-24 | 780 | 7.8 | 1.755** | 2.986** |
| 25-34 | 1796 | 5.3 | 1.161 | 2.134** |
| 35-44 | 1681 | 3.1 | . $663 *$ | 1.166 |
| 45-54 | 1160 | 2.0 | . 423 ** | . 681 |
|  | 712 | 1.0 | .209** | .312* |
| 65-74 | 540 | . 5 | .104** | .131** |
| $75+$ | 217 | 7 | .146* | 166 |
| Region |  |  |  |  |
| Atlantic | 628 | 10.0 | 2.304** | 1.905** |
| Quebec | 2007 | 3.3 | .708* | .770 |
| Ontario | 2577 | 4.0 | . 864 | . 780 |
| Prairies | 1279 | 5.6 | 1.230 | 1.018 |
| B.C. | 1013 | 4.2 | .909 | 859 |
|  |  |  |  |  |
| Language |  |  |  |  |
| English | 5067 | 5.2 | 1.138 | . 976 |
| French | 1791 | 3.5 | . 752 | . 718 |
| Other | 368 | 4.6 | 1.000 | 1.015 |
| Not stated | 278 | 2.5 | . 532 | 1.406 |
| M arital status |  |  |  |  |
| M arried/common-law | 4518 | 3.1 | .663** | . 891 |
| Single/never married | 2185 | 8.7 | $1.976 * *$ | .884 |
| widowed | 292 | 1.1 | .231* | 1.183 |
| Divorced/separated | 508 | 3.3 | . 708 | 1.074 |
| Educational level |  |  |  |  |
| Less than secondary | 1627 | 8.4 | 1.902** | 2.178** |
| Secondary | 1862 | 5.0 | 1.092 | 1.400 |
| Some post-secondary | 2350 | 3.7 | . 797 | . 989 |
| University degree | 1346 | 1.8 | .380** | . 601 |
| Not stated | 319 | 2.1 | 445 | . 552 |
|  |  |  |  |  |
| Income |  |  |  |  |
| Low | 1080 | 7.4 | 1.657** | 1.297 |
| Middle | 2775 | 4.0 | . 864 | . 947 |
| High | 1084 | 3.1 | . 663 | . 974 |
| Not stated | 2564 | 4.8 | 1.046 | 836 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size. * p< .01; **p<. 001

Heavy/infrequent drinkers use alcohol less than once a week but have five or more drinks, when they do drink. Odds for this pattern are significantly increased, when respondents report: being male, being between 15 and 34 , with more than 6 -fold odds among those aged 15 to 19 as compared to the overall odds; living in A tlantic Canada, and/or having not completed secondary school. Attributes

Heavy/infrequent vs.
the rest among current drinkers ,by the rest among current drinkers ,by
sex, age, region, language, marital sta sex, age, region, language, marital sta
tus, education and income, with and without other predictors taken into account

Heavy/Infrequent Drinkers
associated with a reduction in the odds of drinking infrequently and heavily are being female and being between 55 and 74 years of age.

## Table A8

Heavy/frequent vs. the rest among current drinkers, by sex, age, region, language, marital status, education and income, with and without other predictors taken into account

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overail | 7504 | 7.6 |  |  |
| Sex |  |  |  |  |
| Male | 3994 | 11.8 | 1.627** | 2.145** |
| Female | 3509 | 2.9 | .363** | $466 * *$ |
| Age |  |  |  |  |
| 15-17 | 362 | 15.8 | 2.281** | 1.479 |
| 18-19 | 258 | 17.7 | 2.615** | 2.169 ** |
| 20-24 | 780 | 15.6 | 2.247 ** | $2.368^{* *}$ |
| 25-34 | 1796 | 7.9 | 1.043 | $1.728^{* *}$ |
| $35-44$ | 1681 | 6.5 | . 845 | $1.610 * *$ |
| 45.54 | 1160 | 5.1 | .653* | 1.188 |
| 55-64 | 712 | 3.7 | 467** | 769 |
| 65-74 | 540 | 1.8 | .223** | . $369 *$ |
| $75+$ | 217 | 8 | .098** | 140* |
| Region |  |  |  |  |
| Atlantic | 628 | 12.0 | 1.658** | 1.375* |
| Quebec | 2007 | 7.9 | 1.043 | 1.182 |
| Ontario | 2577 | 6.4 | 831 | .750* |
| Prairies | 1279 | 8.3 | 1.100 | . 963 |
| B.C. | 1013 | 6.9 | . 901 | 852 |
| Language |  |  |  |  |
| English | 5067 | 8.0 | 1.057 | 1.271 |
| French | 1791 | 7.6 | 1.000 | . 888 |
| Other | 368 | 3.6 | . $454 *$ | .501* |
| Not stated | 278 | 7.3 | . 957 | 1.766 |
| M arital status |  |  |  |  |
| M arried/common-law | 4518 | 4.4 | .560** | .575** |
| Single/never married | 2185 | 14.8 | $2.112 * *$ | 1.477* |
| widowed | 292 | 2.0 | .248** | . 849 |
| Divorced/separated | 508 | 9.2 | 1.232 | 1.387 |
| Educational level |  |  |  |  |
| Less than secondary | 1627 | 11.1 | 1.518** | 1.862** |
| Secondary | 1862 | 8.3 | 1.100 | 1.293 |
| Some post-secondary | 2350 | 7.6 | 1.000 | 1.041 |
| University degree | 1346 | 2.8 | . 350 * | 422*** |
| Not stated | 319 | 6.8 | 887 | 947 |
| Income |  |  |  |  |
| Low | 1080 | 11.4 | 1.564** | 1.274* |
| Middle | 2775 | 6.9 | . 901 | . 919 |
| High | 1084 | 5.8 | 749 | 936 |
| Not stated | 2564 | 7.6 | 1.000 | . 912 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size.

* p<.01; ** p<.001

The odds for heavy/frequent drinking are significantly increased for male respondents; for those belonging to age groups between 18 and 44, and particularly between 18 and 24; living in Atlantic Canada, being single or never married, having less than secondary school education, and/or a lower income. The odds for
heavy/frequent drinking are reduced for women, people in age groups over 65, people living in Ontario, those speaking a language other than French or English at home, respondents with a partner, and/or those who have completed university.

Noteworthy in these findings is a low proportion of lifetime abstainers in the youngest age group ( $22.2 \%$ for those 15 to 17) followed by a sharp decrease in the next group, where the rates are halved ( $11.3 \%$ for those aged 18 to 19). The results indicate that more than three quarters of the population begin to use alcohol before the age of 15 , and the percentage of those exposed to alcohol is high among recent cohorts. The major concentration of life abstainers is among the older members of the population. Alcohol use per se can take a variety of forms or patterns. The analyses above suggest that these patterns are not idiosyncratic; they are outcomes of complex life situations and beliefs, and they are associated with resources.

Light drinking, and especially light/frequent drinking appears to be an adult pattern that is part of a high-resource life style. Heavier drinking, on the other hand, is a young male pattern and is also associated with being less privileged. The data also suggest that higher-risk use (i.e. heavy drinking) is most prevalent in the Atlantic provinces, and that this area of the country could benefit from welldesigned prevention and intervention efforts.

Although drinking is generally a pleasant and benign activity, misuse of alcohol can harm the physical and mental well-being of the drinker and can also harm others because of the effect of alcohol on the drinker's associates. The survey asked about these two types of negative consequences from alcohol use: self-inflicted harm as a result of one's own use, and harm as a result of drinking by others.

Former and current drinkers were asked to report whether or not they experienced types of trouble due to their alcohol use in their lifetime, and current drinkers were asked about harm during the 12 months prior to the survey. Figure A 1 displays percentages of Canadians reporting different types of alcohol-related harm during their lives and during the past year.

## Harmful Consequences of

 Alcohol Use
## Harm from One's Own

 Drinking

Note: For lifetime harm, percentages are of current and former drinkers. For harm in the 12 months prior to the survey, percentages are of current drinkers. Population estimates (denominators) reflect exclusion of "not applicable" categories.
a Percentage of women only
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability
Nearly one in five (19.8\%) of former and current drinkers have had trouble with drinking during their lives, and more than one in 10 (10.5\%) have had trouble during the past year. Physical health is the most frequently reported problem for both the lifetime and twelve-month time frames ( $12.2 \%$ and $6.2 \%$, respectively).

As Table A9 indicates, harm is strongly associated with drinking patterns. A mong current drinkers, reports of having experienced one or more harmful effects from alcohol use during the past year are strongly associated with patterns of alcohol use.

The results suggest that both higher volume and higher frequency of alcohol use are important when assessing potential for alcohol-related harm. Association between heavy drinking and report of harm is also evident in the percentages reporting harm in the four categories derived from the number of drinking occasions with five or more drinks (bottom of Table A9). Percentages reporting harm is

14 times as high among current drinkers reporting 12 or more heavy drinking occasions in the past 12 months as compared to those reporting none.

Harm from drinking varies with socio-demographic characteristics. Table A9 displays the percentages of those reporting one or more harmful effects during the past year by demographic characteristics.

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Overal! | 16,652 | 10.5 |
| Sex |  |  |
| Male | 8,857 | 12.4 |
| Female | 7.794 | 8.3 |
| Age |  |  |
| 15-17 | 811 | 25.8 |
| 18.19 | 566 | 26.0 |
| $20-24$ | 1,721 | 19.8 |
| $25-34$ | 3,959 | 10.8 |
| $35-44$ | 3,731 | 8.8 |
| 45.54 | 2,579 | 7.4 |
| 55.64 | 1,588 | 3.30 |
| $65-74$ | 1,203 | 3.40 |
| $75+$ | 494 | $\cdots$ |
| Region |  |  |
| Atlantic | 1,341 | 11.8 |
| Quebec | 4,283 | 13.6 |
| Ontario | 6,015 | 5.8 |
| Prairies | 2,789 | 15.1 |
| B.C. | 2.223 | 10.5 |
| Language |  |  |
| English | 11,315 | 9.9 |
| French | 3,824 | 13.5 |
| Other | 824 | 10.30 |
| Not stated | 688 | $\cdots$ |
| Marital status |  |  |
| Married/common-law | 9,953 | 6.5 |
| Single/never married | 4,833 | 19.0 |
| Widowed | 646 | -- |
| Divorced/separated | 1,124 | 14.7 |
| Not stated | 95 | $\cdots$ |
| Educational level |  |  |
| Less than secondary | 3,607 | 14.4 |
| Secondary | 4,114 | 9.3 |
| Some post-secondary | 5,161 | 11.9 |
| University degree | 2,973 | 6.9 |
| Not stated | 796 | 3.10 |
| Income |  |  |
| Low | 2,365 | 15.7 |
| Middle | 6,061 | 10.2 |
| High | 2,390 | 10.4 |
| Not stated | 5,836 | 8.7 |

Table A9
Percentages of current drinkers eporting one or more types of harm xperienced in past 12 months from one's own alcohol use, by sex, age, egion, language, marital status, ducation, income drinking pattern and occasions with $5+$ drinks one's own alcohol use (cont'd)

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Drinking pattern |  |  |
| Light/infrequent | 7,747 | 4.6 |
| Light/frequent | 6,720 | 10.0 |
| Heavy/infrequent | 759 | 24.5 |
| Heavy/frequent | 1,253 | 40.6 |
| Not stated | 173 | $13.7{ }^{\text {Q }}$ |
| 5+drinks (past 12 months) |  |  |
| 12 times or more | 1,720 | 38.0 |
| 3-11 times | 2,637 | 17.8 |
| 1-2 times | 2,805 | 10.6 |
| Never | 8,719 | 2.7 |
| Not stated | 770 | 11.8 |

Note: Percentages are of current drinkers.
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability
A higher percentage of men than women report that drinking has caused them harm ( $12.4 \%$ vs. $8.3 \%$ ). Harm decreases with age group, dropping off sharply after age 25 . Ontario stands out as having a substantially lower percentage (5.8\%) of respondents reporting drinking-related harm; the Prairies (15.1\%) and Quebec (13.6\%) have more than double Ontario's rates.

Those whose home language is French are more likely to report harm (13.5\%) than those who speak English or another language at home. People who have never been married (19\%) or are divorced or separated (14.7\%) report higher rates of harm than do those who are married (6.5\%). Harm is greatest among respondents with less than secondary education (14.4\%) and lowest among those with a university degree (6.9\%). The rate is also higher among those in the lowest income category (15.7\%).

Multivariate analysis assesses the independent contributions of respondent characteristics to the issue of interest. In Table A10, the factors mentioned above were employed to assess their independent contribution to respondents' reports of harm from their own drinking. Two logistic regression analyses were carried out: one with demographic predictors, and another with both demographic and alcohol-use predictors. Adjusted odds ratios for each predictor category obtained from the two analyses are in the last two columns of Table A10. Comparing the two sets of adjusted odds ratios facilitates separation of the influence of drinking behaviour from other factors associated with negative consequences of alcohol use.

The relationship between gender and harm provides an example of the effect of controls on the original bivariate relationship and the effect of including
alcohol-use measures as predictors. Columns 2,3 and 4 respectively represent the percentages, unadjusted odds and adjusted odds ratios for reporting at least one type of harm from alcohol use from the analysis with demographic predictors only. When columns 3 and 4 are compared, the odds ratios for men and women are virtually unchanged; men are more likely than women to report harm from alcoho use during the past year. The 5th column presents a strikingly different picture: once alcohol-use predictors are included, odds for reporting harm are higher for women than for men ( 1.150 vs. .870 ). The finding suggests that, when drinking behaviours and demographic characteristics of men and women are the same, men are less likely and women are more likely to report being harmed by alcohol use.

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio excluding / including alcohol-use predictor |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Overail | 7428 | 10.7 |  |  |  |
| Sex |  |  |  |  |  |
| Male | 3946 | 12.7 | 1.214** | 1.257** | .870* |
| Female | 3482 | 8.5 | 775** | 796** | $1.150 *$ |
| Age |  |  |  |  |  |
| 15-17 | 355 | 26.7 | 3.040** | 2.433** | 2.115** |
| 18-19 | 256 | 25.8 | $2.902 * *$ | $2.586 * *$ | 1.832 * |
| 20-24 | 778 | 20.1 | 2.100** | $2.112 * *$ | 1.396 |
| 25-34 | 1777 | 11.0 | 1.032 | 1.312* | 1.004 |
| 35-44 | 1669 | 8.8 | . 805 | 1.045 | . 906 |
| 45-54 | 1149 | 7.6 | .686** | . 893 | . 923 |
| 55-64 | 699 | 3.4 | 294** | . 367 ** | . $419 \times *$ |
| 65-74 | 536 | 3.5 | . 303 ** | $463 * *$ | 700 |
| $75+$ | 207 | 2.6 | 223 ** | 361 | 751 |
| Region |  |  |  |  |  |
| Atlantic | 626 | 11.8 | 1.117 | 1.021 | . 917 |
| Quebec | 1989 | 13.8 | $1.336 * *$ | 1.348 | 1.307 |
| Ontario | 2535 | 5.8 | .514** | 499** | .557** |
| Prairies | 1274 | 15.1 | $1.484 * *$ | $1.458 * *$ | 1.516** |
| B.C. | 1003 | 10.7 | 1.000 | 999 | 988 |
| Language |  |  |  |  |  |
| English | 5060 | 10.0 | . 927 | 1.059 | . 919 |
| French | 1791 | 13.5 | $1.303^{* *}$ | . 994 | . 954 |
| Other | 367 | 9.7 | 897 | 1.117 | 1.771 |
| Not stated | 210 | 4.6 | 402* | . 851 | 645 |
| Marital status |  |  |  |  |  |
| M arried/common-law | 4472 | 6.6 | .590** | . 775 | . 904 |
| Single/never married | 2164 | 19.4 | 2.009** | 1.494* | 1.306 |
| Widowed | 286 | 1.8 | .153** | 419 | 440 |
| Divorced/separated | 506 | 14.9 | 1.461* | $2.060 * *$ | 1.926** |
| Educational level |  |  |  |  |  |
| Less than secondary | 1625 | 14.4 | 1.404** | 1.499* | 1.326 |
| Secondary | 1860 | 9.3 | . 856 | . 916 | . 820 |
| Some post-secondary | 2350 | 12.1 | 1.149 | 1.091 | 1.065 |
| University degree | 1344 | 7.0 | . $628 *$ | 764 | . 898 |
| Not stated | 249 | 4.5 | 393* | 877 | 962 |

## Table A10

One or more types of harm experienced in past 12 months from ne's own alcohol use, by sex, age, region, language, marital status, education and income, and alcohol-use predictors, with and without other predictors taken into account

One or more types of harm experienced in past 12 months from one's own alcohol use (cont'd)

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio excluding / including alcohol-use predictor |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Income |  |  |  |  |  |
| Low | 1080 | 15.7 | 1.554** | 1.145 | 1.163 |
| Middle | 2775 | 10.3 | . 958 | . 984 | 1.003 |
| High | 1084 | 10.5 | . 979 | 1.175 | 1.051 |
| Not stated | 2489 | 9.1 | 836* | 755** | .816* |
| Drinking pattern |  |  |  |  |  |
| Light/infrequent | 3,495 | 4.7 | .412** |  | .535** |
| Light/frequent | 3,020 | 10.3 | . 958 |  | .898** |
| Heavy/infrequent | 348 | 24.4 | $2.694 * *$ |  | 1.038 |
| Heavy/frequent | 565 | 41.6 | $5.945 * *$ |  | 2.006** |
| 5+drinks (past 12 months) |  |  |  |  |  |
| 12 times or more | 775 | 38.9 | 5.313** |  | 2.802** |
| 3-11 times | 1,194 | 18.1 | $1.844 * *$ |  | 1.345** |
| 1-2 times | 1,275 | 10.7 | 1.000 |  | . 853 |
| Never | 3,902 | 2.8 | . $240 * *$ |  | . $304 * *$ |
| Not stated | 283 | 11.9 | 1.127 |  | 1.023 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size.

* p<.01; **p<. 001

The multivariate procedures confirm the inverse relationship of age to the likelihood of reporting alcohol-related harm. They also allow some further elaboration. Without taking alcohol-use into account, adjusted odds indicate that people under 34 years old are more likely to report their alcohol-use has harmed them. Controlling for alcohol-use qualifies the findings. Odds for harm remain high and significant in the 15 to 19 year age group, indicating that drinkers under 20 are vulnerable to harm from drinking, regardless of their drinking patterns.

Regional relationships (Table A10) are maintained in both multivariate analyses: respondents residing in Ontario are less likely and those in the Prairies more likely to report harm. Adjusting for other factors including alcohol-use, the only marital-status category with significantly increased odds for harm was separated/divorced.

Drinking patterns are related to decreased and increased odds for harm. Light/infrequent drinkers have significantly decreased odds for reporting harm during the past year, while heavy/frequent drinkers have significantly increased odds. Similarly, those respondents who reported 12 or more occasions of drinking five or more drinks during the past year were most likely to report having been harmed by alcohol, while those who reported none of such occasions had significantly decreased odds for harm.


[^2]Respondents were asked to report whether they had experienced harm as a result of ercentages of respondents reporting various types of harm from others alcohol use : lifetime and past 12 months

Harm from Drinking by Others
others' drinking during their lives and during the past 12 months. Twelve kinds of trouble from others were queried. Frequencies for lifetime and the last year's experiences are presented in Figure A2.

The most frequently reported harm in the past 12 months is being disturbed by parties. More serious, however, are the high rates of negative interpersonal experiences (e.g. being insulted/humiliated, having quarrels, losing friends), which attest to the pervasiveness of serious negative social events, where alcohol is involved.

Vulnerability to harm from others appears to be strongly related to drinking patterns. With the exception of reporting being disturbed, heavier drinkers are far more likely to report a given experience than are lighter drinkers. Table A 11
presents the percentages reporting one or more types of harm, excluding experiencing "being disturbed," during the 12 month period.

Table All Percentages reporting one or more ypes of harm experienced in past 12 months from others' alcohol-use, by respondent's drinking pattern

Table A12
Percentages reporting one or more types of harm experienced in past 12 months from others' alcohol-use, by sex, age, region, language, marital status, education and income

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Overail | 23,030 | 30.1 |
| Drinking pattern |  |  |
| Lifetime abstainer | 2,957 | 16.8 |
| Former drinker | 3,098 | 22.5 |
| Light/infrequent | 7,747 | 30.8 |
| Light/frequent | 6,720 | 30.1 |
| Heavy/infrequent | 759 | 59.0 |
| Heavy/frequent | 1,253 | 65.2 |
| Not stated | 495 | 13.30 |

Note: Dependent variable excludes 'disturbed by parties'
Q Qualified release due to high sampling-variability.

The results suggest the importance of the social context involving different types of drinkers and different modes of drinking. It is not surprising that both types of abstainers had relatively low rates of problems from others. There are two probable explanations for the large differences between light and heavier drinkers: (1) heavier drinkers are more likely than others to drink with friends whose similar drinking contributes to troublesome behaviour, (2) the behaviour of heavily drinking respondents contributes to social difficulties. Some combination of the two explanations is probably the most frequent case.

Table A 12 displays the percentages of persons in each demographic category who stated that they had experienced one or more types of harm during the past year because of the drinking of others.

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Oveeraili | 23,030 | 30.1 |
| Sex |  |  |
| Male | 11,337 | 32.2 |
| Female | 11,692 | 28.0 |
| Age |  |  |
| 15-17 | 1,247 | 50.7 |
| 18.19 | 711 | 61.9 |
| 20.24 | 2,051 | 57.4 |
| 25-34 | 4,952 | 36.4 |
| 35-44 | 4.802 | 31.0 |
| 45-54 | 3,531 | 23.5 |
| 55-64 | 2,470 | 14.1 |
| 6 | 2,195 | 8.0 |
| $75+$ | 1.071 | 3.30 |
| Region |  |  |
| Atlantic | 1,907 | 33.1 |
| Quebec | 5,796 | 30.1 |
| Ontario | 8,673 | 25.0 |
| Prairies | 3,715 | 36.7 |
| B.C. | 2.939 | 34.8 |
| Language |  |  |
| English | 15,006 | 32.6 |
| French | 5,170 | 30.2 |
| Other | 1,452 | 23.8 |

(continued)

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Not stated | 1,402 | 9.5 |
| M arital status |  |  |
| M arried/common-law | 13,564 | 23.0 |
| Single/never married | 6,317 | 490 |
| widowed | 1,316 | 9.2 |
| Divorced/separated | 1,587 | 35.4 |
| Not stated | 246 | 13.1 |
| Educational level |  |  |
| Less than secondary | 5,936 | 28.8 |
| Secondary | 5.415 | 308 |
| Some post-secondary | 6,455 | 37.2 |
| University degree | 3,610 | 27.1 |
| Not stated | 1,614 | 10.5 |
| Income |  |  |
| Low | 3,612 | 37.1 |
| M̈ O ddie | 7,742 | 30.7 |
| High | 2,778 | 32.3 |
| Not stated | 8,898 | 26.0 |

Note: Dependent variable excludes "disturbed by parties".
Q Qualified release due to high sampling-variability

Reporting such harm does not differ substantially by gender. Men report harm slightly more than women do ( $32.2 \%$ versus $28 \%$ ). There are, however, large differences with age: percentages reporting harm decrease substantially with age. The age group 18 to 19 has the highest rate of reported harm from others (61.9\%). Differences between regions of the country were not large, but Ontario had the lowest proportion of those reporting negative consequences of drinking by others (25\%) and the Prairie provinces had the highest rates (36.7\%). Those who speak neither English nor French at home reported harm less often.

Almost half of single respondents (49\%), but only (9.2\%) of those who are widowed report having experienced such harm. Regarding educational level, the highest rate of reporting harm occurred among those with some post-secondary education (37.2\%) and those with university degrees had the lowest rates of reporting harm (20.1\%). The lowest income category had a higher rate of reporting harm from others' alcohol-use (37.1\%).

The analysis identifies independent associations between respondent characteristics/drinking behaviours and the odds or likelihood of reporting one or more types of harm from others' drinking. There are two sets of adjusted odds ratios: one excludes alcohol-use measures as predictors and one includes these measures. The results are presented in Table A13.

Table A12
Percentages reporting one or more types of harm experienced in past 12 months from others' alcohol use (cont'd)

Table A13
One or more types of harm experienced in past 12 months from others' alcohol use by sex, age, region, language, marital status, education, income and drinking pattern, with and without other predictors taken into account

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio excluding / including alcohol-use predictor |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Overail | 10044 | 31.2 |  |  |  |
| Sex |  |  |  |  |  |
| Male | 4927 | 33.5 | 1.111** | 1.097** | 1.009 |
| Female | 5117 | 29.0 | 901** | .912** | . 991 |
| A ge |  |  |  |  |  |
| 15-17 | 552 | 51.8 | 2.370** | 2.304** | 2.297** |
| 18-19 | 319 | 63.0 | 3.755** | 3.569** | 3.302** |
| 20-24 | 908 | 58.7 | 3.134** | 3.237** | 2.909** |
| 25-34 | 2181 | 37.7 | $1.334^{* *}$ | $1.736^{*}$ | $1.636 *$ |
| 35-44 | 2090 | 31.8 | 1.028 | 1.402** | 1.341** |
| 45-54 | 1541 | 24.6 | . $719 \times *$ | . 995 | . 987 |
| 55-64 | 1055 | 14.7 | .380** | . $513 * *$ | .533** |
| 65-74 | 936 | 8.6 | . $207 * *$ | .287** | . 316 ** |
| $75+$ | 462 | 3.5 | .080** | 106** | .124** |
| Region |  |  |  |  |  |
| Atlantic | 891 | 33.2 | 1.096 | . 986 | . 964 |
| Quebec | 2662 | 30.7 | . 977 | 1.007 | 1.001 |
| Ontario | 3489 | 26.6 | .799** | .735** | .749** |
| Prairies | 1685 | 37.2 | $1.306^{* *}$ | 1.203** | 1.210** |
| B.C. | 1318 | 35.4 | $1.208 * *$ | 1.138 | 1.143 |
| Language |  |  |  |  |  |
| English | 6693 | 32.8 | 1.076* | 1.348** | 1.276* |
| French | 2417 | 30.2 | . 954 | 1.127 | 1.092 |
| Other | 644 | 23.8 | .689** | . 825 | . 937 |
| Not stated | 291 | 19.2 | . $224 * *$ | 798 | 766 |
| Marital status |  |  |  |  |  |
| M arried/common-law | 5953 | 23.9 | .693** | .701** | .720** |
| Single/never married | 2808 | 50.1 | $2.214 * *$ | $1.203 *$ | 1.180 |
| widowed | 576 | 9.5 | .231** | . 856 | . 860 |
| Divorced/separated | 707 | 36.1 | 1246* | $1.385 * *$ | $1.370 * *$ |
| Educational level |  |  |  |  |  |
| Less than secondary | 2678 | 28.8 | .892* | 1.044 | 1.013 |
| Secondary | 2450 | 30.8 | . 981 | . 948 | . 921 |
| Some post-secondary | 2930 | 37.4 | $1.317 *$ | 1.080 | 1.062 |
| University degree | 1628 | 27.4 | .832** | .827* | . 839 |
| Not stated | 358 | 20.1 | 555** | 1.132 | 1.202 |
| Income |  |  |  |  |  |
| Low | 1647 | 37.2 | $1.306 * *$ | 1.144* | 1.165* |
| Middle | 3545 | 30.8 | . 981 | . 994 | . 984 |
| High | 1258 | 32.3 | 1.052 | 1.051 | 1.019 |
| Not stated | 3594 | 28.5 | 879** | $837 * *$ | .856** |
| Drinking pattern |  |  |  |  |  |
| Lifetime abstainer | 1294 | 17.3 | .461** |  | .486** |
| Former drinker | 1381 | 22.8 | . 651 ** |  | .667** |
| Light/infrequent | 3472 | 31.4 | 1.009 |  | .808** |
| Light/frequent | 2988 | 31.0 | 991 |  | 943 |
| Heavy/infrequent | 346 | 59.3 | $3.213 * *$ |  | 1.642 * |
| Heavy/frequent | 563 | 66.6 | 4.397************* |  | 2.469** |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size. Dependent variable excludes "disturbed by parties".
*p<.01; **p<. 001

When controlled for drinking pattern, rates of reporting harm from others do not differ significantly between men and women. There is, however, a small but significant difference (1.097 vs. .912) when drinking behaviour is excluded from the predictor variables. The apparent gender difference in reporting harm from others' drinking is primarily due to men's drinking patterns, men being more likely to be heavier drinkers.

There is an inverse relationship between age and the rate of reporting harm from others' drinking; younger people are more vulnerable to such negative consequences of alcohol-use, while chances of consequences decrease in older groups. The pattern of results is remarkably unchanged in both multivariate analyses.

The regional variability of the rates of reported harm persists in both multivariate analyses. Residents of Ontario are least likely to report having experienced harm from others' drinking during the past year, while odds of reporting harm are highest for those living in the Prairie provinces.

After controlling for all of the other predictors, the only language-group effect is for English speakers who are more likely to report being harmed as a result of drinking by others. The effect of other language groups is not significant.

Logistic regression analyses confirm that married respondents are significantly less likely, and divorced/separated people are more likely to report harm from others' drinking. Single/never-married respondents have the highest rates for reporting this outcome (50.1\%) and more than twice the likelihood of those in the other marriage categories (unadjusted odds ratio of 2.214). However, controlling for other demographic factors reduces the effect of the category upon the odds (1.203), and including drinking behaviour as a predictor results in a non-significant odds ratio. The reduction in the strength of single status as a predictor of harm probably has its source in the age effect. Younger people are more likely to report having been harmed, and younger people are more likely to be single.

Once adjusted for other factors, educational categories generally do not have a significant impact. A minor exception is the category of respondents with university degrees who have significantly lower odds of harm from others, when respondents' drinking is not taken into account. This effect is not significant once drinking pattern is included as a predictor.

These analyses confirmed that people with the lowest incomes are more likely to report having experienced alcohol-involved social consequences.

The alcohol-use relationship remains strong in the multivariate results, with a slight adjustment. Abstainers and light/infrequent drinkers are less likely to report

## Family/M arital Problems and Assaults

Table A14
Family/marital problems in the past 12 months due to others' alcohol use, by sex, age, region, language, marital status, education, income and drinking pattern, with and without other predictors taken into account
harm from others' alcohol-use in the past year, while odds are higher for heavy/infrequent and heavy/frequent drinkers. The odds ratio for light/frequent drinkers is not significant.

The above analyses addressed harm from drinking by others without specifying which types of harm were involved. The vulnerability of different subpopulations to alcohol-related social consequences differs with the type of trouble. To illustrate this point, analyses were done of two types of harm: the impact of others' alcoholuse (1) on the family and (2) on being hit, pushed or assaulted.

For brevity, discussion of these two types of harm combines tabular and multivariate analyses.


| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio excluding / including alcohol-use predictor |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| University degree | 1625 | 3.6 | $630 \times 1$ | . 542 ** | .558** |
| Not stated | 343 | 4.0 | 702 | 1.813 | 1.924 |
| Income |  |  |  |  |  |
| Low | 1646 | 7.2 | 1.308* | 1.103 | 1.091 |
| Middle | 3539 | 5.6 | 1.000 | . 949 | . 941 |
| High | 1256 | 6.2 | 1.114 | 1.209 | 1.214 |
| Not stated | 3572 | 4.6 | .813* | .790* | .803* |
| Drinking pattern |  |  |  |  |  |
| Lifetime abstainer | 1291 | 2.6 | .450** |  | .496** |
| Former drinker | 1375 | 6.2 | 1.114 |  | 1.069 |
| Light/infrequent | 3464 | 6.6 | 1.191 |  | . 995 |
| Light/frequent | 2979 | 4.2 | .739** |  | . 811 |
| Heavy/infrequent | 346 | 9.3 | 1.729 |  | 1.349 |
| Heavy/frequent | 559 | 9.9 | 1.852** |  | $1.731^{* *}$ |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size.

* p<.01; **p<. 001

Table A14 indicates that family/marital problems are not frequently reported in the sample: $5.6 \%$ of the respondents reported they experienced problems in the family or in the marriage because of someone else's drinking during the past year. Close to twice as many women as men report having had recent family or marriage problems because of someone else's drinking (3.9\% versus $7.2 \%$ ). This gender effect persists in both logistic regression analyses, with odds for women doubling the odds for men.

While percentage differences in age categories are not large, they are consistent; younger respondents are more likely to report this type of problem. It is most likely that respondents in the youngest age groups are referring to problems with parents or siblings, and not problems with a spouse. Multivariate analyses confirms such a pattern due to age groups.

Regional variations were minor, and none persisted in the multivariate analyses. People speaking English at home are most likely to report having alcohol-related family or marital problems, and separated or divorced people have increased odds of such problems.

The lower rates among university graduates observed remained intact after multivariate controls. People who have attained this level of education have decreased odds of reporting alcohol-related family problems. While higher rates were reported by people in the low-income category, the effect was no longer present in the multivariate results. Lifetime abstainers report the lowest rates of problems $(2.6 \%)$, and, among current drinkers, light/frequent drinkers report lower rates (4.2\%). The proportions for both categories of heavy drinkers are more than double the light/frequent drinkers' rates (heavy/infrequent 9.3\%; heavy/frequent 9.9\%). According to the multivariate results, the odds of heavy/frequent drinkers reporting marital/family problems (1.731) more than triple those for lifetime abstainers (.496).
mily/marital problems in the past 12 months due to others' alcohol use (cont'd)

Table A15
Physical assault in past 12 months due to others' alcohol use, by sex, age, region, language, marital status, education, income and drinking pattern, with and without other predictors taken into account

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | excluding alcohol- | ds ratio including redictor |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Overail | 9964 | 12.0 |  |  |  |
| Sex |  |  |  |  |  |
| Male | 4884 | 15.4 | 1.335** | 1.374** | 1.252** |
| Female | 5080 | 8.7 | .699** | .728** | 799** |
| A ge |  |  |  |  |  |
| 15-17 | 536 | 28.0 | 2.852** | 3.365** | 3.454** |
| 18-19 | 319 | 33.5 | 3.694** | 4.734** | 4.267** |
| 20-24 | 905 | 33.7 | 3.728** | 5.578** | 4.857** |
| 25-34 | 2169 | 14.6 | 1.254** | $2.217^{* *}$ | 2.023** |
| 35-44 | 2083 | 8.8 | .708** | 1.276 | 1.188 |
| 45-54 | 1522 | 5.9 | . 460 ** | . 820 | . 802 |
| 55-64 | 1047 | 2.9 | .219** | .374** | 392** |
| 65-74 | 927 | 1.2 | .089** | .160** | 183** |
| 75+ | 455 | 6 | .044** | .081** | 101** |
| Region |  |  |  |  |  |
| Atlantic | 899 | 13.0 | 1.096 | 1.010 | 965 |
| Quebec | 2650 | 12.1 | 1.010 | . 970 | . 990 |
| Ontario | 3440 | 10.0 | .815** | .788** | .807* |
| Prairies | 1680 | 14.4 | 1.234* | 1.162 | 1.161 |
| B.C. | 1305 | 13.3 | 1.125 | 1.115 | 1.118 |
| Language |  |  |  |  |  |
| English | 6649 | 12.5 | 1.048 | 1.238 | 1.147 |
| French | 2411 | 12.0 | 1.000 | 1.240 | 1.142 |
| Other | 642 | 9.9 | . 806 | . 971 | 1.143 |
| Not stated | 262 | 5.0 | .386** | . 672 | 668 |
| M arital status |  |  |  |  |  |
| M arried/common-law | 5909 | 6.8 | .535** | .637** | .666** |
| Single/never married | 2785 | 24.9 | 2.431** | 1.159 | 1.127 |
| Widowed | 571 | 1.5 | .112** | . 770 | .773 |
| Divorced/separated | 699 | 13.7 | 1.164 | 1.760** | $1.726 * *$ |
| Educational level |  |  |  |  |  |
| Less than secondary | 2661 | 12.2 | 1.019 | 1.252 | 1.188 |
| Secondary | 2441 | 12.6 | 1.057 | 1.089 | 1.050 |
| Some post-secondary | 2915 | 14.4 | $1.234 * *$ | 1.016 | . 988 |
| University degree | 1621 | 7.9 | .629** | .694* | 712* |
| Not stated | 327 | 5.5 | .427** | 1.040 | 1.141 |
| Income |  |  |  |  |  |
| Low | 1642 | 14.9 | 1.284** | 1.023 | 1.036 |
| Middle | 3538 | 11.4 | . 944 | 1.041 | 1.029 |
| High | 1252 | 11.6 | . 962 | 1.142 | 1.104 |
| Not stated | 3532 | 11.4 | . 944 | .822* | 849* |
| Drinking pattern |  |  |  |  |  |
| Lifetime abstainer | 1289 | 4.7 | .362** |  | 422** |
| Former drinker | 1371 | 6.3 | .493** |  | .616** |
| Light/infrequent | 3439 | 11.3 | . 934 |  | 893 |
| Light/frequent | 2965 | 11.3 | . 934 |  | 1.042 |
| Heavy/infrequent | 345 | 32.5 | 3.531** |  | 1.746** |
| Heavy/frequent | 556 | 38.2 | 4.533** |  | 2.373** |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size Variable derived from A39P "pushed/shoved" and A 39V "hit/assaulted".

* p<.01; ** p<. 001

In contrast with family/marital problems, prevalence of physical assault due to alcohol-use is high: $12 \%$ of the respondents reported being physically bothered in some way as a result of someone else's drinking (Table A15). Also in contrast with the above problems, men had a higher rate (15.4\%) than women (8.7\%), and the results are verified in the multivariate results. Younger people had dramatically higher rates of assault, especially between ages 18 and 24 .

Those living in Ontario report the lowest rates for physical harm, and the finding persists in the multivariate analysis. Cross-tabulations suggest large differences in rates of reporting physical harm from others according to marital status People with partners have lower rates (6.8\%), single people have the highest rates $(25.9 \%)$, the proportion of widows is negligible (1.5\%), while (13.7\%) of those who are divorced or separated report having had this problem during the past year. When the factors are adjusted in the multivariate analysis, marriage remains a protector against physical harm, and the status of divorced/separated is associated with increased odds of reporting such harm. Odds ratios for single status and widowhood are not significant. The only educational category with significant effects after multivariate adjustment was university completion; respondents with a university degree are significantly less likely than others to report physical harm. Higher rates are reported among people in the low-income category, but the effect does not persist in the multivariate analysis.

Rates for reporting physical harm from others increase considerably with drinker categories, and the increase is especially sharp between the light and heavy drinker categories (light/infrequent and frequent, both 11.3\%, heavy/infrequent, $32.5 \%$, heavy/frequent, $38.2 \%$ ). The multivariate analysis distinguishes between the non-drinkers, whose odds of assault are significantly decreased, and heavy drinkers, whose odds are increased.

Harm to self and harm from others due to drinking are results of a complex of characteristics, circumstances and behaviour. Certain characteristics appear to be consistently linked to an array of alcohol-related misfortune. Clearly, youth seems to be a vulnerable time for self- and other-inflicted harm. While youth is an ascribed characteristic, it is not a permanent one. People who live in Ontario can expect relatively lower rates of trouble involving alcohol. Because of the repetition of independent regional effects, some closer examination is in order. Both financial and cultural capital appear fairly consistently as protectors against trouble from alcohol. Having less income and less education is associated with more trouble. Social supports also appear to be related to the likelihood of reporting trouble with
alcohol. Marriage/having a partner seems to be a general protector against such trouble, while being separated or divorced is associated with increased harm from one's own drinking and from the drinking behaviour of others.

References
1 Yukon Bureau of Statistics. (1991). Yukon Alcohol and Drug Survey. Whitehorse: Executive Council Office.

2 Yukon Bureau of Statistics. (1994). What the Numbers Say: A Review of the Methodology and the Results of the 1993 Yukon Health Promotion Survey. Whitehorse: Executive Council Office

## CHAPTER 3

## TOBACCO

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The human cost of tobacco use is high. In 1991, an estimated 45,064 deaths in
Canada were attributable to smoking. ${ }^{1}$ Smoking is associated with numerous health problems, including cardiovascular diseases, cancer, respiratory diseases, diseases of the digestive system, health problems during pregnancy and pediatric diseases. Exposure to environmental tobacco smoke is associated with an increased risk of acute and chronic health problems. Particularly vulnerable to environmental tobacco smoke are children. Smoking is also associated with house fires and other accidents. Clearly, smoking is a major public-health problem.

Based on CADS, 1994, overall, 45.5\% of Canadians had never smoked, 26.3\% used to smoke, and $27 \%$ were current smokers. Figure T1 presents tobacco-use trends in terms of overall percentages in the Canadian population. Overall, in the past 30 years, there has been a decreasing trend in the proportions of current smokers, with a relatively stable prevalence from 1990 to $1994 .{ }^{2}$ However, among those aged 15 to 19, the prevalence of current smoking increased to $27 \%$ in 1994 from $23 \%$ in 1991. ${ }^{3}$ In 1994, the prevalence of current smoking among 15- to 19-year-olds was $30.4 \%$. The reversal of the downward trend is compatible with the appearance of low-priced contraband in Canada in the early 1990s, and the subsequent lowering of federal taxes on cigarettes in February 1994. Price as a determinant of smoking, especially among youth, has been highlighted in the literature. ${ }^{4}$ Given the likelihood of a lifetime's addiction to tobacco emanating from adolescent smoking, the increasing evidence of a reversal in prevalence of smoking among youth is cause for concern.


Figure T1
Trends in tobacco use of Canadians 1965 to 19941

Sources: ${ }^{1}$ Williams, B. Canadian Profile: Alcohol, Tobacco and Other Drugs, 1995
${ }^{2}$ Statistics Canada. Survey on Smoking in Canada, 1994
${ }^{3}$ Health Canada. Canada's Alcohol and Other Drugs Survey, 1994.

The "quitting rate" is defined as the percentage of people who have quit among all those who have ever smoked, that is the percentage of former smokers divided by the percentages of former and current smokers combined. The Canada-wide quitting rate is $49.3 \%$, indicating that about one half of those who have ever smoked are no longer smoking. The last column in Table T1 presents demographic information about quitting rates. The proportion of people quitting smoking increases with age, to more than $66 \%$ for people 65 and older and from about 1 in 4 for those 15 to 24. The higher quitting rate among the older age groups may reflect a greater need or desire to quit smoking, for example, because of medical problems. Regional differences ranging from 46.1\% in Quebec to 53.4\% in British Columbia, and gender differences ( $50.3 \%$ for men and $48.3 \%$ for women) are not statistically significant.

Single or never-married people are less likely (32.2\%), and married or widowed people are more likely ( $56.2 \%$ and $62.4 \%$ respectively) to have quit smoking as compared to the overall rate of $49.3 \%$. The association between marital status and smoking is at least partly due to the age effect observed above. The quitting rate also increases with both educational level (from $42.6 \%$ to $64.7 \%$ ) and income adequacy (from 40.5\% to 59\%).

|  | $\begin{aligned} & \text { Population } \\ & \text { estimate (000s) } \end{aligned}$ | tobacco-use Status (Percent) |  |  |  | Quitting rate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Never smoked | Former smoker | Current smoker | Not stated |  |
| Overall percentages in the population (as reported in the Preview 1995) | 23,030 | 45.5 | 26.3 | 27.0 | 1.3 | 49.3 |
| Overall percentages (excluding "not-stated" category) | 22,352 | 46.0 | 26.7 | 27.3 |  | 49.3 |
| Sex |  |  |  |  |  |  |
| Male | 10,978 | 42.0 | 29.2 | 28.8 |  | 50.3 |
| Female | 11,374 | 50.0 | 24.3 | 25.7 |  | 48.3 |
| Age |  |  |  |  |  |  |
| 15-17 | 1,221 | 62.6 | 10.6 | 26.8 |  | 28.0 |
| 18-19 | 705 | 51.7 | 11.40 | 36.9 |  | 23.4 |
| 20-24 | 2,014 | 53.0 | 12.4 | 34.6 |  | 26.6 |
| 25-34 | 4,854 | 48.7 | 20.2 | 31.1 |  | 39.3 |
| 35-44 | 4,643 | 41.3 | 28.9 | 29.9 |  | 48.6 |
| 45-54 | 3,417 | 40.2 | 33.1 | 26.7 |  | 55.3 |
| 55-64 | 2,363 | 39.6 | 37.9 | 22.6 |  | 62.9 |
| 65-74 | 2,101 | 44.2 | 37.6 | 18.2 |  | 66.3 |
| 75+ | 1,034 | 55.9 | 35.6 | 8.50 |  | 81.5 |
| Region |  |  |  |  |  |  |
| Atlantic | 1,886 | 39.0 | 29.3 | 31.7 |  | 47.9 |
| Quebec | 5,748 | 37.7 | 28.7 | 33.6 |  | 46.1 |
| Ontario | 8,165 | 53.8 | 23.3 | 22.9 |  | 50.1 |
| Prairies | 3,675 | 45.6 | 28.0 | 26.4 |  | 51.4 |
| B.C. | 2,878 | 46.1 | 28.7 | 25.2 |  | 53.4 |



Note: Quitting rate = \% former/(\% former + \% current)
Q Qualified release due to high sampling-variability

In 1994, an estimated 6.2 million (27\%) Canadians 15 years of age or older were current smokers. Among smokers, $2.8 \%$ smoke an average of fewer than one cigarette per day; $31.2 \%$ smoke from one to 10 cigarettes per day; $58.7 \%$ smoke 11 to 25 cigarettes per day; and $7.3 \%$ smoke 26 or more cigarettes per day. Among current smokers, males are more likely than females to smoke 26 or more cigarettes per day ( $9.9 \%$ vs. $4.6 \%$ ), and females are more likely than males to smoke 10 or fewer cigarettes per day ( $38.1 \mathrm{vs} .30 .1 \%$ ). The highest proportions of people smoking 26 or more cigarettes per day occurs among those aged 45 to 64 years (13.6\%).

## About 27\% of the Canadian population report being current smokers. Table T1

 shows that a higher percentage of males ( $28.8 \%$ ) than females $(25.7 \%$ ) report smoking. The highest proportion of current smokers is found among youth; about $37 \%$ of 18 - to 19 - year olds, and $35 \%$ of 20 - to 24 -year olds report being current smokers. Smoking is much less prevalent among older age groups; fewer than one in five respondents over age 65 report being current smokers.The proportions of the population who smoke vary across Canada. The highest proportions are found in Quebec and in the Atlantic region, where about one in
three respondents report being current smokers. The lowest rates are found in Ontario, where fewer than one in four respondents report being current smokers.

Smoking status also appears to vary according to language spoken at home. The lowest proportion of current smoking (21.9\%) occurs among people who speak other than English or French at home. A majority (62.3\%) in this language group have never smoked.

Table T1 also suggests that a larger proportion of francophones than of anglophones are current smokers ( $33.3 \%$ vs. $26.1 \%$ ). However, the French Ianguage is not distributed evenly across the country-most persons whose language at home is French reside in Quebec. Table T2 presents the percentages of current smokers by language within each region. Once region is taken into account, there is very little difference in the proportions of current smoking when comparing persons whose Ianguage at home is French with those whose language is English. For example, in Quebec, $34 \%$ of francophones smoke compared with $35.3 \%$ of anglophones. In Ontario, $26.8 \%$ of francophones smoke compared with $23.1 \%$ of anglophones.

| Language spoken at home |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | English |  | French |  | Other |  | Overal! |  |
|  | $\begin{aligned} & \text { Pop. est. } \\ & \text { (000s) } \end{aligned}$ | \% | Pop. est. (000s) | \% | Pop. est. (000s) | \% | $\begin{aligned} & \text { Pp. est. } \\ & \text { (000 } 10 \end{aligned}$ | \% |
| Canada | 14,779 | 26.1 | 5,138 | 33.3 | 1,437 | 21.9 | 22,352 | 27.3 |
| Atlantic | 1,646 | 32.3 | 215 | 28.00 | 14 | -- | 1,886 | 31.7 |
| Quebec | 596 | 35.3 | 4,678 | 34.0 | 268 | 24.80 | 5,748 | 33.6 |
| Ontario | 6,646 | 23.1 | 176 | 26.8 | 774 | 23.9 | 8,165 | 22.9 |
| Prairies | 3,297 | 27.4 | 49 | 22.60 | 200 | 17.00 | 3,675 | 26.4 |
| B.C. | 2,594 | 26.3 | 20 | -- | 180 | 15.80 | 2,878 | 25.2 |

Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

The proportions of the population who smoke also vary according to marital status (Table T1). A bout $39 \%$ of divorced or separated respondents and $32 \%$ of single/never-married respondents report being current smokers, whereas less than $25 \%$ of married and $19 \%$ of widowed respondents report being current smokers. The low prevalence in the last group is probably related to age, as confirmed by logistic regression.

Current smoking is inversely related to education and income, that is, the higher the educational and income levels, the lower the percentage of current smokers (Table T1). About one in three people (34.7\%) with less than secondary-
school education smoke, compared with about one in seven (14.8\%) with a university degree. Whereas $35.1 \%$ of people with a low income currently smoke, only $22 \%$ of people with a high income do.

Thus, current smoking is associated with age, gender, region, marital status, education and income. It remains to be seen whether each of these characteristics is related to smoking, when the other factors are taken into account.

Table T3 presents the results of logistic regression comparing current smokers with the rest of the respondents. Using this technique, further confirmation of the associations between current smoking and many demographic characteristics is obtained. Taking into account all other demographic characteristics, males are more likely to be current smokers than are females. People from 18 to 54 years old are more likely to be current smokers (adjusted odds ratio from 1.29 to 1.86) and those aged 15 to 17 and 65 and over are less likely (adjusted odds ratio of .73 and less than .60 ) as compared to the overall odds. People living in Quebec are more likely to be current smokers (adjusted odds ratio 1.40) and people in Ontario are less likely (adjusted odds 0.81 ) in comparison to the overall odds. Those who are divorced or separated are more likely to be current smokers (adjusted odds ratio 1.43), and married people less likely (adjusted odds ratio 0.74).

|  | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overall | 10227 | 27.4 |  |  |
| Sex |  |  |  |  |
| Male | 5024 | 28.9 | 1.077 | 1.086** |
| Female | 5204 | 25.9 | . 926 | . $921 * *$ |
| A ge |  |  |  |  |
| 15-17 | 559 | 26.9 | . 975 | .728* |
| 18-19 | 322 | 36.7 | $1.536 * *$ | $1.600 * *$ |
| 20-24 | 921 | 34.6 | $1.402 * *$ | $1.856 * *$ |
| 25-34 | 2218 | 31.2 | 1.202** | 1.785** |
| 35-44 | 2126 | 30.0 | $1.136 *$ | 1.625** |
| 45-54 | 1564 | 26.8 | . 970 | 1.290** |
| 55-64 | 1082 | 22.7 | .778** | . 864 |
| 65-74 | 962 | 18.3 | . $594 \times *$ | . $604 \times 1$ |
| $75+$ | 473 | 8.6 | 249** | . $237 \ldots$ |
| Region |  |  |  |  |
| Atlantic | 894 | 31.7 | 1.230* | 1.099 |
| Quebec | 2709 | 33.6 | 1.341 ** | 1.403 ** |
| Ontario | 3591 | 22.9 | .787** | .813** |
| Prairies | 1701 | 26.4 | . 950 | . 892 |
| B.C. | 1333 | 25.2 | 893 | 889 |
| Language |  |  |  |  |
| English | 6711 | 26.2 | . 941 | 1.218* |
| French | 2416 | 33.3 | 1.323** | 1.062 |

Table T3
Current smokers vs. current nonsmokers by sex, age, region, language, marital status, education and income, with and without other predictors taken into account

Table T3

|  | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Other | 650 | 21.9 | .743* | . 881 |
| Not Stated | 450 | 20.3 | .675** | 877 |
| M arital status |  |  |  |  |
| M arried/common-law | 6078 | 24.6 | .864** | .735** |
| Single/never married | 2845 | 32.1 | 1.253** | . 978 |
| widowed | 591 | 18.9 | .618** | . 969 |
| Divorced/separated | 714 | 39.2 | 1.708** | 1.434** |
| Educational level |  |  |  |  |
| Less than secondary | 2684 | 34.8 | 1.414** | 1.854** |
| Secondary | 2447 | 30.3 | 1.152* | 1.143 |
| Some post-secondary | 2937 | 26.3 | . 946 | .835* |
| University graduate | 1632 | 14.8 | . $460 * *$ | . $445 *$ |
| Not stated | 528 | 21.3 | .717* | 1.268 |
| Income |  |  |  |  |
| Low | 1647 | 35.2 | 1.439** | 1.221** |
| Middle | 3544 | 29.8 | 1.125 * | 1.068 |
| High | 1258 | 22.1 | .752** | . 893 |
| Not stated | 3779 | 23.5 | .814** | .859** |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size. * $p<.01$; ** $p<.001$

Taking into account all other demographic characteristics, educational level is found to be a strong predictor of current smoking. The odds are inversely related to the educational level, with the odds ratios decreasing to 0.45 (university degree) from 1.85 (less than secondary). Finally, income level is also a predictor of current smoking. People with a low income are more likely to be current smokers (adjusted odds ratio 1.22) as compared to the overall odds.

Logistic regression analysis helps clarify the relationship between current smoking and language spoken at home. Once all the other predictors including region are taken into account, people who speak French at home are not found to be at significantly higher risk. However, people who speak English at home are found to be at higher risk (adjusted odds ratio 1.22).

There is an association between smoking status and alcohol consumption (Table T4). The definitions for the categories of drinking patterns are found in the section on alcohol.

The proportion of current smoking increases according to increasing alcohol consumption, to 59.3\% current smokers among heavy/frequent drinkers from 13\% current smokers among lifetime alcohol-abstainers. The majority ( $75 \%$ ) of lifetime abstainers from alcohol have never smoked. The majority of heavy drinkers are current smokers ( $51 \%$ for the heavy/infrequent and $59.3 \%$ for the heavy/frequent).

Furthermore, the association between tobacco- and alcohol-use remains strong,
when the demographic variables are taken into account in a multivariate analysis (data not presented).

| Drinking Pattern | Pop. est. (000s) | TOBACCO-USE STATUS (PERCENT) |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Lifetime Non-smoker | Former Smoker | Current Smoker |
| Lifetime abstainer | 2,890 | 75.0 | 12.0 | 13.0 |
| Former drinker | 3,073 | 44.1 | 31.4 | 24.5 |
| Light/infrequent drinker | 7,706 | 47.1 | 26.8 | 26.0 |
| Light/frequent drinker | 6,682 | 39.0 | 33.5 | 27.4 |
| Heavy/infrequent drinker | 757 | 30.1 | 18.9 | 51.0 |
| Heavy/frequent drinker | 1,243 | 24.1 | 16.5 | 59.3 |

Note: $N=22,030,000$.

1 Ellison, L.F., Mao, Y., Gibbons, L. (1995). Projected Smoking-attributable M ortality in Canada, 1991-2000. Chronic Diseases in Canada, Spring: 84-89.

2 Williams, B., Single, E., M cKenzie, D. (1995). Canadian Profile: Alcohol, Tobacco and Other Drugs. Ottawa: Canadian Centre on Substance Abuse and Addiction Research Foundation

3 Stephens, T. (1995). Trends in the Prevalence of Smoking, 1991-1994. Chronic Diseases in Canada, Winter: 27-32.

4 Ibid.

Table T4
Tobacco-use status
and drinking pattern

References

## CHAPTER 4

## LICIT DRUGS

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CADS asked respondents about their use of five types of prescription medications. Here, use is defined as having taken prescribed pain pills, sleeping pills, tranquilizers, antidepressants or diet pills (stimulants) in the 12 months prior to the survey.

There appears to have been a slight downward trend in the use of prescription tranquilizers and sleeping pills in the past decade. Prevalence of tranquilizer-use decreased to $4 \%$ in 1994 from $6 \%$ in $1985^{1}$ and $5 \%$ in 1990. ${ }^{2}$ The prevalence of the use of sleeping pills decreased to $4.5 \%$ in 1994 from $8 \%$ in $1985^{3}$ and $7 \%$ in 1990. ${ }^{4}$ Of note is that the decreasing trend is pronounced among women. From 1985 to 1994, the proportions of women using tranquilizers decreased to $5 \%$ from $8 \%$, and the proportions of women using sleeping pills decreased to $5 \%$ from $10 \% .{ }^{5}$ No clear trend emerges regarding the use of the remaining prescription medications.

Overall, $20.8 \%$ of Canadians used at least one of the five prescription medications in 1994. A larger proportion of females than males reported using at least one of the medications ( $23.9 \%$ vs. $17.7 \%$ respectively).

The proportions of the population reporting use of each of the five medications and use of at least one of the five medications is presented in Table M1. The most commonly reported medication was prescription pain medication (13.1\%). About three to five percent of Canadians reported using sleeping pills, tranquilizers or antidepressants. Fewer than one in 100 Canadians reported using diet pills (0.9\%).

|  | $\begin{aligned} & \text { Pop. } \\ & \text { est. (000s) } \end{aligned}$ | TYPE OF PRESCRIPTION MEDICATION |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Pain pills | Sleeping pills | Tran- | Antidepressants | Diet pills | One or more |
| Overall | 23,030 | 13.1 | 4.5 | 4.3 | 3.0 | 9 | 20.8 |
| Sex |  |  |  |  |  |  |  |
| Male | 11,338 | 12.0 | 3.7 | 3.4 | 1.7 | .70 | 17.7 |
| Female | 11,692 | 14.1 | 5.4 | 5.3 | 4.2 | 1.0 | 23.9 |
| A ge |  |  |  |  |  |  |  |
| 15-24 | 4,010 | 14.0 | 3.0 | 1.30 | 1.20 | 1.90 | 18.5 |
| 25-44 | 9,754 | 13.9 | 3.5 | 2.9 | 3.0 | . 70 | 19.8 |
| 45-64 | 6,001 | 11.3 | 4.7 | 6.4 | 3.9 | . 50 | 20.4 |
| $65+$ | 3,265 | 12.6 | 9.2 | 8.4 | 3.2 | -- | 27.4 |
| Region |  |  |  |  |  |  |  |
| Atlantic | 1,907 | 13.3 | 4.3 | 4.4 | 3.3 | 1.10 | 21.5 |
| Quebec | 5,796 | 6.8 | 5.8 | 6.8 | 3.7 | . 70 | 18.5 |
| Ontario | 8,673 | 12.6 | 3.5 | 3.3 | 1.8 | .70 | 17.9 |
| Prairies | 3,715 | 17.5 | 4.5 | 3.2 | 3.7 | 1.20 | 24.9 |
| B.C. | 2,939 | 21.2 | 5.3 | 4.0 | 3.9 | . 90 | 28.4 |

Table M1
Percentages reporting prescriptionmedication use in past 12 months, by sex, age, region, language, marital status, education and income

Table M 1
Percentages reporting prescriptionmedication use in past 12 months

|  | $\begin{aligned} & \text { Pop. } \\ & \text { est. (000s) } \end{aligned}$ | TYPE OF PRESCRIPTION MEDICATION |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Pain pills | $\begin{gathered} \text { Sleeping } \\ \text { pills } \\ \text {...... } \end{gathered}$ | Tranquilizers | Antidepressants | Diet pills | One or more |
| Language |  |  |  |  |  |  |  |
| English | 15,000 | 16.6 | 4.4 | 3.8 | 3.1 | 1.1 | 23.7 |
| French | 5,170 | 7.3 | 6.4 | 7.5 | 3.9 | . 5 Q | 19.8 |
| Other | 1,452 | 6.6 | 3.40 | 2.30 | -- | -- | 10.7 |
| Not stated | 1,402 | 3.30 | -- | -- | -- | -- | 4.20 |
| M arital status |  |  |  |  |  |  |  |
| M arried/common law | 14,000 | 12.8 | 4.2 | 4.4 | 2.8 | . 60 | 20.1 |
| Single/never married | 6,317 | 13.1 | 3.9 | 2.7 | 2.0 | 1.6 | 19.2 |
| Widowed | 1,316 | 13.4 | 9.4 | 7.5 | 4.60 | -- | 27.9 |
| Divorced/separated | 1,587 | 16.4 | 6.6 | 7.9 | 7.1 | -- | 28.6 |
| Not stated | 246 | -- | -- | -- | -- | -- | 9.80 |
| Educational level |  |  |  |  |  |  |  |
| Less than secondary | 5,936 | 13.8 | 6.0 | 6.9 | 3.4 | 1.00 | 24.5 |
| Secondary | 5,415 | 14.0 | 4.1 | 4.0 | 3.0 | 1.00 | 21.3 |
| Some post-secondary | 6,455 | 14.4 | 4.3 | 3.6 | 2.8 | 1.00 | 21.6 |
| University degree | 3,610 | 12.5 | 5.1 | 3.6 | 3.5 | -- | 19.7 |
| Not stated | 1,614 | 3.60 | -- | -- | -- | -- | 5.20 |
| Income |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |
| Low | 3,612 | 14.4 | 5.9 | 5.9 | 3.7 | 1.60 | 25.1 |
| Middle | 7,742 | 14.4 | 4.9 | 4.3 | 3.7 | . 70 | 22.7 |
| High | 2,778 | 14.9 | 4.7 | 4.1 | 2.80 | -- | 22.2 |
| Not stated | 8,898 | 10.9 | 3.6 | 3.8 | 2.1 | . 7 Q | 17.0 |

Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

The results in Table M1 show a striking pattern. For all of the above types of medications, as well as for use of at least one of them, a larger proportion of females than males report use over the course of 12 months. The association between prescription drug use and gender has been observed in previous surveys in Canada. ${ }^{6,7,8}$

Use of pain medication appears to be more -or- less evenly distributed across all age groups. In the case of diet pills, young people are more likely than older people to use these stimulants. In contrast, use of prescribed tranquilizers, sleeping pills and antidepressants increases with age (Table M1). For example, about 1\% of people under 20 years of age use tranquilizers, whereas, by 65 years of age, more than $8 \%$ do. A similar pattern emerges for sleeping pills and antidepressants. Use of one or more medications increases to $27.4 \%$ in the oldest (65+) from $18.5 \%$ in the youngest age group (15-24). The finding of an association between aging and use of prescription psychotropic medications is consistent with previous surveys. ${ }^{9}, 10,11$ That older people may be on psychotropic and many other medications at the same time is a cause for concern. Because of a reduced rate of drug metabolism and
excretion in the elderly, use of psychotropic medications can result in a decreased state of alertness or in harmful drug interactions, which can further compromise the health of the elderly.

There are large regional differences in use of some types of prescription drugs. In Quebec, $6.8 \%$ of people 15 years old and older report using prescription pain medication, whereas $21.2 \%$ of people in British Columbia report use. In the other regions, prevalence of use ranges from $12.6 \%$ to $17.5 \%$. Use of tranquilizers appears to be most prevalent in Quebec, where $6.8 \%$ of the population reports use. Similar regional differences were noted in the National Alcohol and Other Drugs Survey in 1989 and the General Social Survey in 1993. ${ }^{12,13}$

Prescription-medication use is also related to the language spoken at home (Table M1). In the case of tranquilizers and sleeping pills, a larger proportion of francophones than anglophones report using those medications (for tranquilizers, $7.5 \%$ vs. $3.8 \%$; for sleeping pills, $6.4 \%$ vs. $4.4 \%$ ). However, the reverse is the case with prescription pain medication, where $16.6 \%$ of anglophones in contrast with $7.3 \%$ of francophones report use.

In the current survey, marital status appears to play a role in the use of some prescription medications. Separated, divorced or widowed people are more likely than single or married people to report using prescription sleeping pills, tranquilizers and antidepressants. For example, $9.4 \%$ of widowed people but only about 4\% of single or married people report using prescription sleeping pills.

People who used prescription pain pills, sleeping pills, tranquilizers, antidepressants or stimulants for 30 days or more in the past 12 months were asked about harmful effects arising from their use, including harm to their social life, physical health, happiness, home life, marriage, work or finances.

As presented in Figure M 1, 15.7\% of respondents who were asked these questions reported having experienced at least one harm as a result of their drug use in their lifetime. Also, $11.6 \%$ of respondents who had used at least one of the above medications 30 days or more in the 12 months prior to the survey reported having experienced at least one harm recently (in the 12 months prior to the survey).

Harmful Consequences
Arising from the Use of Prescription Medications

Figure M1 Percentages of prescriptionmedication users reporting various types of harm from medication use: lifetime and past 12 months


Note: Percentages of those who used at least one of the five prescription medications (pain pills, sleeping pills, tranquilizers, antidepressants and diet pills) for 30 days or more in the past 12 months. Population estimates reflect exclusion of "not applicable" categories.
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability
Figure M1 shows the proportions of prescription-medication users who reported various types of harm during their lifetime and in the past 12 months. The most frequently reported harm was harm to physical health, reported by $10.6 \%$ of people who used medications in their lifetime and by $7.3 \%$ of those who used medications in the 12 months prior to the survey. Least frequently reported were problems with friendships, spouse/partners or children.

| Age | Pop. est.(000s) | Lifetime | Past 12 months |
| :---: | :---: | :---: | :---: |
| Alil ages | 2,667 | 15.7 | 11.6 |
| 15-24 | 232 | 31.7 | 25.50 |
| 25-44 | 899 | 19.0 | 14.1 |
| $45-64$ | 829 | 16.7 | 12.7 |
| $65+$ | 706 | 5.20 | -- |

Note: Percentages of those who used at least one of the five prescription medications (pain pills, sleeping pills, tranquilizers, antidepressants and diet pills) for 30 days or more in the past 12 months.
Q Qualified release due to high sampling-variability
-- $\quad$ Not for release due to unacceptably high sampling-variability

The reporting of harmful consequences of medication use appears to be inversely related to age. Table M2 shows that, with increasing age, in general,
decreasing proportions of persons reported experiencing at least one harm from their medication use, both during their lifetime and recently. This finding was unexpected, especially given that the use of tranquilizers, antidepressants and sleeping pills increases with increasing age.

A possible explanation is that harm itself may be perceived, recalled or reported differently by persons in different age groups. For example, young persons may consider acute side effects from psychotropic medications as problems with physical health. Or, older persons may interpret physical harm to mean chronic health problems but ignore acute side effects. Or, older persons may not acknowledge new health problems arising from the very medications intended to treat existing health problems. Further research is needed to clarify the association between age and harmful consequences arising from the use of prescription medications.

1 Lamarche, P. and Rootman, I. (1988). Drug use. In Rootman I., Warren, R., Stephens T., Peters, L. (eds.), Canada's Health Promotion Survey: Technical' Report. Ottawa: M inister of Supply and Services.

2 Adlaf E.M. (1993). Alcohol and Other Drug Use. In Stephens, T. and Fowler-Graham, D. (eds.), Canada's Health Promotion Survey 1990: Technical Report. Ottawa: M inistry of Supply and Services Canada.

3 Lamarche, P. and Rootman, I. (1988). Canada's Health Promotion Survey: Technical Report.

4 Adlaf, E.M. (1993). Canada's Health Promotion Survey 1990: Technical Report
5 Lamarche, P. and Rootman, I. (1988). Canada's Health Promotion Survey: Technica Report.

6 Adlaf, E.M. (1993). Canada's Health Promotion Survey: Technnical Report.
7 Single, E., M acLennan, A., M acNeil, P. (1994). Horizons 1994: Alcohol and Other Drug Use in Canada. Ottawa: Health Canada and Canadian Centre on Substance Abuse.

8 Eliany, M., Giesbrecht, N., Nelson, M., Wellman, B., Wortley, S., (eds.) (1992). Alcoho and Other Drug Use by Canadians: A National Alcohol and Other Drugs Survey (1989) Technical Report. Ottawa: Supply and Services Canada.

9 Adlaf, E.M. (1993). Canada's Health Promotion Survey: Technnical Report.
10 Single, E. et al. (1994). Horizons 1994
11 Eliany, M. et al. (1992). Alcohol and Other Drug Use by Canadians
12 Ibid.
13 Single, E. et al. (1994). Horizons 1994

## CHAPTER 5

## ILLICIT DRUGS

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CADS included questions on lifetime and current use of five types of illegal substances: cannabis, cocaine or crack, LSD, amphetamines and heroin, as well as steroids and solvent sniffing. Table D1 presents the prevalence (overall and by gender) of use of each substance and combinations of drugs. There is a consistent gender difference across most of the drugs and drug combinations: a larger proportion of males than females report using these drugs.

The proportions of Canadians reporting the use of cocaine or crack, LSD, amphetamines and heroin, are small and appear to have changed little from 1989 to 1994. However, prevalence of cannabis use has fluctuated over the past five years. The proportion of Canadians reporting use was $6.5 \%$ in 1989, $5 \%$ in 1990, $4.2 \%$ in 1993 and 7.4\% in 1994.1,2,3

Cannabis is the most widely used illegal drug in Canada (Table D1). The prevalence of lifetime use is $28.2 \%$, while current use (any use in the 12 months preceding the survey) is $7.4 \%$. (Note: If 'one-time users' are excluded, the percentages are $23.1 \%$ and $7 \%$ respectively.) The rates are consistently higher among males

During the 12 months prior to the survey, $3.2 \%$ of Canadians reported use at least once a month. A gain, a gender difference exists with $4.7 \%$ for men and $1.8 \%$ for women.

Table D2 presents the percentages of respondents reporting any use of cannabis according to key demographic characteristics. In general, cannabis is used by young people. The proportion of users steadily decreases with age, to less than $1.4 \%$ among people more than 45 years of age and $10 \%$ among people 25 to 34 , from about $25 \%$ among those 15 to 19 years old. Cannabis use also varies according to region, with the highest proportion in British Columbia (11.6\%), and the lowest in Ontario (5.1\%),

|  | Lifetime |  |  | Current |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Overail | Male | Female | Overail | Male | Female |
| Pop. Est. (000s) | 23,030 | 11,337 | 11,692 | 23,030 | 11,337 | 11,692 |
| Illicit drugs |  |  |  |  |  |  |
| Cannabis ${ }^{1}$ | 28.2 | 33.5 | 23.1 | 7.4 | 10.0 | 4.9 |
| Cannabis² | 23.1 | 27.7 | 18.7 | 7.0 | 9.5 | 4.6 |
| Crack/Cocaine | 3.8 | 4.9 | 2.7 | . 7 | . 80 | . 50 |
| LSD | 5.2 | 7.2 | 3.3 | 9 | 1.3 | . 60 |
| Speed (Amphetamines) | 2.1 | 3.1 | 1.2 | . 20 | 40 | -- |
| Heroin | . 5 | 8 | --. | --..... | -- | -- |
| LSD/Speed/Heroin | 5.9 | 8.1 | 3.6 | 1.1 | 1.5 | . 70 |
| Any illicit drug use ${ }^{3}$ | 28.5 | 33.6 | 23.5 | 7.6 | 10.1 | 5.1 |

Table D1
Percentages reporting use of illicit drugs, steroids and solvents in lifetime and past 12 months, overall and by sex

|  | Lifetime |  |  | Current |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Overall | Male | Female | Overall | Male | Female |
| iilicit drug use ${ }^{4}$ | 23.9 | 28.5 | 19.4 | 7.3 | 9.7 | 4.9 |
| Steroids | 0.30 | 0.40 | -- | -- | -- | -- |
| Solvents | 0.8 | 1.2 | 0.30 | -- | -- | -- |

Includes "one-time only" use
Excludes "one-time only" use
Use of at least one of 5 illicit drugs, including "one-time only" use of cannabis
Use of at least one of 5 illicit drugs, excluding "one-time only" use of cannabis
Qualified release due to high sampling-variability

- Not for release due to unacceptably high sampling-variability

Cannabis use varies according to the language spoken at home, with the lowest proportion (2.7\%) occurring among people who speak other than English or French at home. However, there is little difference in the proportions of cannabis users between those who speak English and those who speak French (8.0\% and 8.7\% respectively). Regarding marital status, cannabis use is markedly more prevalent among the single or never-married. This finding is probably related to the young age of most cannabis users.

No clear association emerges between cannabis use and educational level. Finally, a larger proportion of people in the low-income group than the remaining income-adequacy groups use cannabis.

Table D3 presents the results of analysis comparing cannabis users with non-users, taking into account all the demographic characteristics at the same time.

Males are more than twice as likely to be cannabis users than are females (1.49 compared to 0.67 ), according to the adjusted odds ratios. Young people are more likely to be cannabis users, and the odds of being a cannabis user steadily decreases with age. The odds ratio of being a cannabis user relative to overall odds are five-to-one among people 15-17 years old, four-to-one among 18- to 24-year olds, two-and one-half-to-one among 25- to 34 year olds, and one and one-half-to-one among those 35 to 44 . Youthful age is a strong predictor of cannabis use.

Regarding region, people living in the Atlantic region and in Ontario are the least likely (adjusted odds ratios 0.72 and 0.68 ), and residents of British Columbia are the most likely to be cannabis users (adjusted odds ratio 1.78). People who speak other than English or French at home are the least likely to be cannabis users (adjusted odds ratio 0.37).

Marital status is not a statistically significant predictor of cannabis use. In particular, because most cannabis users are young, the real risk factor is youthful age. Finally, taking into account all other demographic characteristics, different educational and income levels do not appear to predict cannabis-use.

In summary, the relationship between cannabis use and all the demographic characteristics was examined using logistic regression. The following groups have increased odds (or are at increased risk) of being cannabis users: males, persons 15 to 44 years of age and residents of British Columbia. The strongest predictor of cannabis use is youthful age, particularly 15 to 24 years.

| Variable/Category | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Overail | 23,030 | 7.4 |
| Sex |  |  |
| Male | 11,337 | 10.0 |
| Female | 111,692 | 4.9 |
| Age |  |  |
| 15-17 | 1,247 | 25.4 |
| 18-19 | 711 | 23.0 |
| $20-24$ | 2,051 | 19.3 |
| 25-34 | 4,952 | 9.6 |
| 35-44 | 4,802 | 5.8 |
| 45-54 | 3,531 | 1.40 |
| 55-64 | 2,470 | -- |
| $65+$ | 3,265 | -- |
| Region |  |  |
| Atlantic | 1,907 | 6.3 |
| Quebec | 5,796 | 8.6 |
| Ontario | 8,673 | 5.1 |
| Prairies | 3,715 | 8.2 |
| B.C. | 2,939 | 11.6 |
| Language |  |  |
| English | 15,006 | 8.0 |
| French | 5,170 | 8.7 |
| Other | 1,452 | 2.70 |
| Not stated | 1.402 | $\cdots$ |
| Marital status |  |  |
| Married/common-law | 13,564 | 3.3 |
| Single/never married | 6,317 | 18.1 |
| Widowed | 1,316 | -- |
| Divorced/separated | 1,587 | 6.5 |
| Not stated | 246 | $\cdots$ |
| Educational level |  |  |
| Less than secondary | 5,936 | 8.2 |
| Secondary | 5,415 | 7.2 |
| Some post-secondary | 6,455 | 8.9 |
| University degree | 3,610 | 6.6 |
| Not stated | 1,614 | -- |
| Income |  |  |
| Low | 3,612 | 9.9 |
| Middile | 7,742 | 6.9 |
| High | 2,778 | 78 |
| Not stated | 8.898 | 6.7 |

Note: Current cannabis use = use of cannabis in past 12 months prior to the survey, including 'one-time only' use.
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

Current cannabis-users vs, non-users by sex, age region, language marita status, education and income, with and without other predictors taken into account

Cannabis use in Relation to Alcohol and Tobacco Use

|  | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overail | 9,999 | 7.7 |  |  |
| Sex |  |  |  |  |
| Male | 4,897 | 10.5 | 1.406** | 1.490** |
| Female | 5,102 | 5.0 | . 631 * | 671** |
| Age |  |  |  |  |
| 15-17 | 546 | 26.1 | 4.234** | 5.332** |
| 18-19 | 322 | 22.9 | 3.560** | 4.433** |
| 20-24 | 905 | 20.0 | 2.997 ** | 4.202** |
| 25-34 | 2,164 | 10.0 | 1.332** | 2.504** |
| 35-44 | 2,086 | 6.0 | .765* | $1.588 *$ |
| 45-54 | 1,541 | 1.4 | .170** | .355** |
| 55-64 | 1,046 | 7 | . $085 \times *$ | 175** |
| $65+$ | 1,388 | 1 | . 012 * | .041* |
| Region |  |  |  |  |
| Atlantic | 886 | 6.3 | . 806 | .723* |
| Quebec | 2,645 | 8.8 | 1.157 | 1.088 |
| Ontario | 3,480 | 5.3 | . 671 ** | .683** |
| Prairies | 1,677 | 8.3 | 1.085 | 1.048 |
| B.C. | 1,311 | 11.8 | 1.604*** | 1.776** |
| Language |  |  |  |  |
| English | 6,678 | 8.0 | 1.042 | 1.254 |
| French | 2,413 | 8.7 | 1.142 | 1.351 |
| Other | 647 | 2.8 | .345** | .373** |
| Not stated | 261 | 1.9 | .232** | 1.585 |
| M arital status |  |  |  |  |
| M arried/common law | 5,929 | 3.4 | .422** | 0.646 |
| Single/never married | 2,794 | 18.7 | 2.757** | 1.653 |
| Widowed | 572 | . 3 | .036** | . 580 |
| Divorced/separated | 703 | 6.3 | 806 | 1.615 |
| Educational level |  |  |  |  |
| Less than secondary | 2,673 | 8.0 | 1.042 | 1.456 |
| Secondary | 2,440 | 7.3 | . 944 | 1.325 |
| Some post-secondary | 2,929 | 9.0 | $1.186 *$ | 1.385 |
| University degree | 1,625 | 6.6 | . 847 | 1.453 |
| Not stated | 331 | 1.2 | .146** | . 258 |
| Income |  |  |  |  |
| Low | 1,646 | 9.9 | 1.317** | 1.127 |
| Middle | 3,538 | 6.8 | . 875 | . 996 |
| High | 1,256 | 7.8 | 1.014 | 1.084 |
| Not stated | 3,559 | 7.5 | . 972 | .822* |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size. * p > .01; ** p < . 001

Cannabis use is associated with alcohol and tobacco use. Table D4 shows the proportion of cannabis users for each alcohol- and tobacco-consumption pattern. The definitions for these consumption categories are found in the sections on alcohol and tobacco. Whereas about 4\% of lifetime non-smokers and former tobacco smokers also report having used cannabis in the 12 months prior to the survey, 16.3\%
of current tobacco smokers report being cannabis users. The proportion of cannabis users increases according to increasing alcohol consumption, to 34\% among heavy/frequent drinkers from less than $1.2 \%$ among lifetime alcohol-abstainers.

Although the results are not presented, alcohol- and tobacco-use predictors were submitted to logistic regression-cannabis-use together with all the demographic predictors presented in Table D3. When adjusted for all the other predictors, both tobacco-use and alcohol-use predictors remain significant. Furthermore, they are the second- and third-strongest predictors after age, in terms of their unique contribution to explaining cannabis-use.

|  | Pop. est. (000s) | Percent |
| :---: | :---: | :---: |
| Overall | 23,030 | 7.4 |
| Drinking pattern |  |  |
| Lifetime abstainer | 2,957 | $<1.2^{\text {a }}$ |
| Former drinker | 3,098 | 2.10 |
| Light/infrequent | 7,747 | 5.0 |
| Light/frequent | 6,720 | 9.1 |
| Heavy/infrequent | 759 | 21.9 |
| Heavy/frequent | 1,253 | 34.0 |
| Not stated | 495 | 5.79 |
| Smoker type |  |  |
| Lifetime non-smoker | 10,481 | 4.0 |
| Former smoker | 6,047 | 4.5 |
| Current smoker | 6,208 | 16.3 |
| Not stated | 294 | -- |

Note: Current cannabis use = use of cannabis in past 12 months, including 'one-time only' use.
a $\quad<1.2$ is based on the upper limit of $99 \%$ confidence interval. Actual estimate has been suppressed because of unacceptably high sampling-variability
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

Illicit drug-use is defined as use of at least one of five illicit drugs (cannabis, cocaine/crack, LSD, amphetamines and heroin), excluding 'one-time only' use of cannabis. Table D5 presents the percentages of respondents reporting illicit drug-use in their lifetime. Overall, 23.9\% of Canadians are estimated to have used illegal drugs and a higher percentage of males than females report illicit drug-use in their lifetime ( $28.5 \%$ vs. $19.4 \%$ ). Lifetime illicit drug-use is most prevalent among young persons. Whereas 30 to $38.2 \%$ of people 15 to 44 years of age report lifetime illicit drug-use, less than $14.8 \%$ of people 45 years of age and over do so.

Table D5 also presents the percentages of respondents reporting current illicit drug-use (in the 12 months prior to the survey). About 7.3\% of Canadians report such drug-use, with a higher percentage of males than females ( $9.7 \%$ vs. $4.9 \%$ ). The highest proportion of use is reported by young persons. A bout one in four people

Table D4
Current cannabis-use, and alcoholand tobacco- use patterns

15 to 19 years of age, and one in five people 20 to 24 years of age, report having used illicit substances in the previous 12 months. The prevalence of current illicit drug-use among people 45 years of age or older is less than $1.4 \%$. The region with the highest prevalence of current illicit drug-use is British Columbia (11.4\%) and the region with the lowest proportion is Ontario (4.8\%).

|  | Pop. est. (000s) | Percent |  |
| :---: | :---: | :---: | :---: |
|  |  | Lifietime | Current |
| Overall | 23,030 | 23.9 | 7.3 |
| Sex |  |  |  |
| Male | 11,337 | 28.5 | 9.7 |
| Female | 11,692 | 19.4 | 4.9 |
| Age |  |  |  |
| 15-17 | 1,247 | 30.0 | 24.0 |
| 18-19 | 711 | 32.9 | 23.8 |
| 20-24 | 2,051 | 37.7 | 19.0 |
| 25-34 | 4,952 | 38.2 | 9.6 |
| 35-44 | 4,802 | 32.9 | 5.7 |
| 45-54 | 3,531 | 14.8 | 1.40 |
| 55-64 | 2,470 | 3.70 | -- |
| $65+$ | 3,265 | 80 | -- |
| Region |  |  |  |
| Atlantic | 1,907 | 21.9 | 6.0 |
| Quebec | 5,796 | 25.3 | 8.7 |
| Ontario | 8,673 | 17.5 | 4.8 |
| Prairies | 3,715 | 27.5 | 8.2 |
| B.C. | 2.939 | 36.6 | 11.4 |
| Language |  |  |  |
| English | 15,006 | 26.4 | 7.8 |
| French | 5,170 | 26.4 | 8.9 |
| Other | 1,452 | 8.0 | 2.40 |
| Not stated | 1,402 | 3.80 | -- |
| Marital status |  |  |  |
| M arried/common-law | 13,564 | 20.2 | 3.2 |
| Single/never married | 6,317 | 35.5 | 18.0 |
| Widowed | 1,316 | 2.30 | -- |
| Divorced/separated | 1,587 | 28.0 | 6.5 |
| Not stated | 246 | 12.50 | -- |
| Educational level |  |  |  |
| Less than secondary | 5,936 | 19.0 | 8.2 |
| Secondary | 5,415 | 24.0 | 6.8 |
| Some post-secondary | 6,455 | 30.0 | 9.0 |
| University degree | 3,610 | 29.5 | 6.4 |
| Not stated | 1,614 | 4.10 | -- |
| Income |  |  |  |
| Low | 3,612 | 24.8 | 9.6 |
| Middie | 7,742 | 27.7 | 6.7 |
| High | 2,778 | 33.8 | 7.9 |
| Not stated | 8,898 | 17.1 | 6.6 |

Note: Illicit drug-use = use of at least one of 5 illicit drugs (see Table D1) excluding 'one-time only' use of cannabis.
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

An estimated 1.7 million Canadians (7.4\% of population) have used at least one of the injectable drugs (cocaine/crack, LSD, amphetamines, heroin and steroids) in their lifetime. Reflecting the general patterns of drug-use, the proportion is higher among males ( $10 \%$ vs. $4.9 \%$ ) and among those younger than 45 (greater than $10.8 \%$ ). Of those, $7.7 \%(132,000)$ report having injected drugs at some time in their lives.

Sharing needles is a special concern because of the risk of transmitting bloodborne infections such as HIV and hepatitis B. When asked whether they had shared needles with anyone, $41.4 \%$ of those who had reported injection drug-use (IDU) in their lifetime responded affirmatively. ${ }^{\text {a }}$

Less than one percent of all respondents reported sniffing solvents or glue in their lifetime and in the 12 months prior to the survey. The use of inhalants occurs mostly among young persons and other groups that may not be reached easily by a telephone survey.

Respondents who reported drug-use were asked about problems or harm caused by it. Here, lifetime drug-use is defined as having used at least one of cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids or solvents excluding 'one-time only' use of cannabis or solvents, while current drug-use does not exclude trying cannabis or solvents only once. An estimated 4,436,000 (19.3\%) Canadians are lifetime drug-users, and 1,588,000 (6.9\%) are current drug-users. The possible harmful effects about which respondents were asked included harm to their social life, physical health, happiness, home life or marriage, work or finances.

Figure D1 shows the proportions of drug-users who reported harm from their drug-use. The most frequently reported harm was harm to physical health, reported by $17.6 \%$ of lifetime users and by $12.5 \%$ of current users. The least frequently reported were problems with home life, spouse/partner or children, reported by less than $10 \%$ of lifetime or current users. One intriguing gender difference emerged in relation to harm arising from drug-use in the 12 months prior to the survey: harm to home life was reported by $5.8 \%$ of male users and $11.8 \%$ of female users. However, both of these estimates have a rather high sampling-variability, and the difference only approaches the significance level.

About $26.9 \%$ of lifetime drug-users as defined above reported having experienced at least one harm as a result of their drug-use (Table D6). A small gender

[^3]difference observed ( $28.7 \%$ of males vs. $24.2 \%$ of females) falls short of our significance level. A larger proportion of young persons than of older persons experienced harm. For example, more than $36 \%$ of persons 15 to 19 years of age reported harmful effects from their drug-use, whereas less than $24 \%$ of persons 35 years of age or older reported such harm. There is a wide regional variation in reporting harm from drug-use, with the lowest proportion being reported in Ontario (15.6\%) and the highest being reported in the Prairies (37.3\%) and in Quebec (35.2\%).

Figure D1
Percentages of lifetime and current drug-users reporting various types of harm from drug- use


Note: Lifetime drug-use = use in lifetime of at least one of cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids, or solvents, excluding 'one-time use' of cannabis or solvents.
Current drug-use = use in past 12 months of at least one of cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids or solvents.
a Discrepancies between these percentages and those reported in "Preview 1995" are due to exclusion of 'not applicable' responses
Q Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

Illicit drug-users potentially face legal repercussions associated with their druguse. A mong persons who reported having used drugs during their lifetime, 7.7\% reported having had contact with the police, with a higher proportion among males than females (10.7\% vs. 3.1\%).

|  | Lifetime |  | Current |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Pop. est. (000s) | Percent | Pop. est. (000s) | Percent |
| Overall | 4,436 | 26.9 | 1,588 | 23.8 |
| Sex |  |  |  |  |
| Male | 2,676 | 28.6 | 1,050 | 24.1 |
| Female | 1,760 | 24.2 | 537 | 23.2 |
| Age |  |  |  |  |
| 15-17 | 302 | 38.5 | 279 | 36.2 |
| 18-19 | 196 | 36.2 | 154 | 34.50 |
| 20-24 | 646 | 30.1 | 364 | 21.0 Q |
| 25-34 | 1,546 | 28.1 | 456 | 23.1 |
| 35-44 | 1,297 | 24.0 | 271 | 13.40 |
| 45-54 | 380 | 12.80 | 47 | -- |
| $55+$ | 70 | -- | 17 | -- |
| Region |  |  |  |  |
| A tlantic | 324 | 26.5 | 102 | 28.8Q |
| Quebec | 1,213 | 35.2 | 491 | 34.9 |
| Ontario | 1,268 | 15.6 | 403 | 15.10 |
| Prairies | 781 | 37.3 | 279 | 29.3 |
| B.C. | 851 | 22.3 | 312 | 11.10 |
| Language |  |  |  |  |
| English | 3,197 | 24.9 | 1,103 | 20.3 |
| French | 1,126 | 33.2 | 443 | 33.3 |
| not stated | 113 | 19.20 | 42 | -- |

Note: Lifetime drug-use = use in lifetime of at least one of cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids or solvents, excluding 'one-time only' use of cannabis or solvents.
Current drug-use = use in past 12 months of at least one of cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids or solvents.
Qualified release due to high sampling-variability
-- Not for release due to unacceptably high sampling-variability

Among the current drug-users, $23.8 \%$ reported harm arising from their druguse. (Table D6). Gender does not seem to be an issue in experiencing harm from drug-use. However, experiencing harm does appear to be related to young age. More than one in three persons 15 to 19 years of age reported harm from their drug-use, whereas by 35 years of age, less than $13 \%$ of persons reported harm from their drug-use. Reporting harm is related to the region, in that the highest proportion of harm is reported from drug-users in Quebec (34.9\%), and the lowest is from drug-users in British Columbia (11.1\%). Finally, the rate of reporting harm from current drug-use is higher among francophones than among anglophones.

As in the previous section on tobacco, logistic regression was used to help clarify the relationship between harm arising from drug-use and the demographic characteristics of drug-users. Table D7 presents the results of the analysis comparing drugusers in terms of reported harm, taking into account gender, age, region and Ianguage at the same time.

Table D6
Percentages of lifetime and current drug- users reporting one or more types of harm from drug-use, by sex age, region and language

Confirmation of the Correlates of Harm Arising from Drug-Use

One or more types of harm reported by current drug-users by sex age region and language, with and without other predictors taken into account

| Variable/Category | Weighted sample size | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| :---: | :---: | :---: | :---: | :---: |
| Overall | 713 | 23.5 |  |  |
| Sex |  |  |  |  |
| Male | 475 | 23.9 | 1.022 | 1.107 |
| Female | 238 | 22.7 | . 956 | . 904 |
| Age |  |  |  |  |
| 15-17 | 124 | 36.6 | 1.879** | 2.170* |
| 18-19 | 69 | 33.0 | 1.603 | 1.902 |
| 20-24 | 166 | 21.0 | . 865 | 1.052 |
| 25-34 | 205 | 22.2 | . 929 | 1.156 |
| 35-44 | 123 | 13.2 | . 495 | 550 |
| 45-54 | 19 | 7.9 | . 279 | 416 |
| $55+$ | 8 | 17.1 | . 672 | 872 |
| Region |  |  |  |  |
| Atlantic | 47 | 30.0 | 1.395 | 1.472 |
| Quebec | 229 | 34.3 | 1.700** | 2.460* |
| Ontario | 168 | 13.8 | .521* | .516* |
| Prairies | 127 | 29.6 | 1.369 | 1.410 |
| B.C. | 142 | 10.0 | .362** | 380** |
| Language |  |  |  |  |
| English | 488 | 19.9 | 809 | 1.340 |
| French | 207 | 32.7 | 1.582** | 820 |
| Not stated | 19 | 15.9 | . 616 | 910 |

Note: $\mathrm{N}=713$, weighted by ESSPROV. Current drug-use $=$ use of at least one of the following substances: cannabis, cocaine/crack, LSD, amphetamines, heroin, steroids or solvents in past 12 months. * $\mathrm{p}<0.01$; ** $\mathrm{p}<0.001$

Two important relationships are confirmed. First, young age, in particular, age 15 to 17, is an independent risk-factor for experiencing harm from drug-use. Drugusers in that age group have more than a two-fold risk (adjusted odds ratio $=2.17$ ) of reporting harm relative to the overall. Second, drug-users in Quebec are most likely (adjusted odds ratio $=2.46$ ) and those in British Columbia are least likely (adjusted odds ratio $=0.38$ ), to report harm. Once adjusted for regional effect, language spoken at home has no significant impact on experiencing harm from current drug-use.

1 Adlaf, E.M. (1993). Alcohol and Other Drug Use. In Health and Welfare Canada. Stephens, T. and Fowler-Graham, D. (eds.). Canada's Health Promotion Survey 1990: Technical Report. Ottawa: Ministry of Supply and Services Canada.
2 Single, E., MacLennan, A., and MacNeil, P. (1994). Horizons, 1994. Alcohol and Other Drug-Use in Canada. Ottawa: Health Canada and Canadian Centre on Substance Abuse.

3 Eliany, M., Giesbrecht, N., Nelson, M., Wellman, B., Wortley, S. (eds.) (1992). Alcohol and Other Drug Use by Canadians: A National Alcohol and Other Drugs Survey (1989) Technical Report. Ottawa: Supply and Services Canada.

## CHAPTER 6

## GAMBLING

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CADS included a series of questions on gambling behaviour. Although gambling does not entail abuse of psycho-active substances, some types of gambling behaviour can be considered to be addictive behaviour, which can cause harmful consequences to gamblers, their families and their communities. Gambling is also associated with alcohol and other drug-use. ${ }^{1}$

As seen in Figure G1, the vast majority of Canadians over the age of 15 engage in some form of gambling, the most frequent form of which is playing a lottery, betting on sports, or playing cards for money, which is reported by $60.7 \%$ of respondents in the past 12 months. Nearly half of the respondents (46.5\%) report playing a lottery, betting on sports, or playing cards for money, monthly or more. The second most common type of gambling is playing bingo, reported by $13 \%$ of respondents, with $6.4 \%$ playing monthly or more. Of those who report betting on a lottery, sports or cards, or playing bingo, 9.3\% travelled to places such as Las Vegas, Atlantic City or Canadian cities with casinos in order to gamble, with $2.2 \%$ doing so at least once a month. Furthermore, $5.5 \%$ report engaging in some other form of gambling in the past 12 months, such as video lottery machines (.8\%) or buying a raffle ticket (.7\%). Overall, $32.1 \%$ of Canadians do not gamble at all, and $16.5 \%$ are infrequent gamblers (less than monthly), 48.4\% are frequent gamblers, in that they play lotteries or cards, bingo, travel to gamble, or gamble in some other way, monthly or more often.


The two major types of gambling - betting on lotteries, cards, or sports and playing bingo - are related to one another, but not as strongly as might be

Prevalence of Gambling

## Fiqure G1

Percentages reporting two forms of gambling in past 12 months, at least once, and monthly or more often

Table G1
Gambling in past 12 months, by sex, age, region, language, marital status, education, income and drinking pattern, with and without other pattern, with and without other predictors taken into account
expected. Lottery players are more than twice as likely as non-lottery players to play bingo (17.1\% vs. 7.4\%).

The relationship between the two major types of gambling and several key sociodemographic variables was examined. People who play bingo differ from people who bet on lotteries, sports or cards, in terms of their sex, age, region, language, marital status, education, income and drinking pattern. Logistic regression was used to examine the relationship between gambling and particular characteristics, taking into account all the information on demographic characteristics at the same time.

Table G1 presents the results of logistic regression comparing gamblers with those who do not gamble. The first column gives the percent of respondents with a particular characteristic who engage in that type of gambling; the second column gives the same result in terms of an odds ratio, indicating how their odds for engaging in gambling compares with the overall or average odds for the entire sample. If the odds ratio is greater than 1.0, their odds are higher than the average, and if the number is less than 1.0, their odds are lower than the average. The third column presents the "adjusted" odds ratios obtained through a series of logistic regression analyses. The "adjusted" odds ratio may be interpreted as a ratio indicating how the odds for that particular group of people would compare to the average odds, controlling for the confounding influence of the other predictors. The table also shows the statistical significance of the relationship between predictor categories and the two types of gambling behaviour.

| Variable/Category | Weighted sample size | Bets on lotteries, cards, sports |  |  | Plays bingo |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Percent | $\begin{aligned} & \text { Unadjusted } \\ & \text { odds } \\ & \text { ratio } \end{aligned}$ | $\begin{aligned} & \text { Adjusted } \\ & \text { odds } \\ & \text { ratio } \end{aligned}$ | Percent | Unadjusted odds ratio | Adjusted odds ratio |
| Overail | 10,105 | 62.7 |  |  | 13.4 |  |  |
| Sex |  |  |  |  |  |  |  |
| Male | 4,959 | 66.1 | 1.160** | 1.141** | 8.2 | .577** | .649** |
| Female | 5,146 | 59.4 | 870** | 876** | 18.5 | $1.467 \times 1$ | 1.541 ** |
| Age |  |  |  |  |  |  |  |
| 15-17 | 547 | 28.6 | .238** | .287** | 11.6 | . 848 | .607** |
| 18-19 | 320 | 40.9 | 412** | .504** | 11.0 | . 799 | . 808 |
| 20-24 | 913 | 55.4 | . $739 \times *$ | . 995 | 15.4 | 1.176 | $1.531 *$ |
| 25-34 | 2,193 | 67.0 | $1.208 * *$ | $1.576 * *$ | 13.9 | 1.043 | $1.353 * *$ |
| 35-44 | 2,108 | 67.7 | $1.247 \times *$ | $1.541 * *$ | 11.8 | 865 | 1.062 |
| 45-54 | 1,545 | 70.0 | $1.388 * *$ | $1.743 * *$ | 11.6 | 848 | . 965 |
| 55-64 | 1,062 | 70.6 | $1.429 * *$ | 1.769** | 13.9 | 1.043 | . 987 |
| 65-74 | 954 | 62.1 | 975 | $1.232 *$ | 16.1 | 1.240 | 996 |
| $75+$ | 463 | 48.4 | . $558 \%$ | .755* | 18.0 | 1.419* | 976 |
| Region |  |  |  |  |  |  |  |
| Atlantic | 893 | 60.6 | . 915 | . 881 | 17.8 | 1.399** | 1.164 |
| Quebec | 2,696 | 72.3 | $1.553 * *$ | $1.382 * *$ | 12.2 | 898 | .740* |



| Language |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| English | 6,691 | 59.9 | .889** | . 867 | 13.9 | 1.043 | . 937 |
| French | 2,418 | 72.8 | 1.592 * | 1.050 | 13.4 | 1.000 | 1.113 |
| Other | 644 | 58.4 | 835 | . 903 | 8.5 | . $600^{* *}$ | 595** |
| Not stated | 352 | 55.5 | 742* | 1.217 | 12.7 | 940 | 1.611 |


| M arital status |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| M arried/common-law | 6,008 | 67.3 | 1.224** | . 995 | 13.1 | . 974 | 991 |
| Single/never married | 2,807 | 52.0 | .644** | $861 *$ | 12.1 | 890 | 886 |
| Widowed | 580 | 58.5 | 839 | 1.071 | 21.6 | $1.781 * *$ | 1.158 |
| Divorced/separated | 710 | 70.0 | 1.388** | 1.091 | 14.5 | 1.096 | 984 |
| Educational level |  |  |  |  |  |  |  |
| Less than secondary | 2,676 | 62.5 | 991 | 1.421** | 20.0 | 1.616** | 2.042** |
| Secondary | 2,447 | 67.1 | $1.213 *$ | 1.259** | 14.1 | $1.061 *$ | 1.245 |
| Some post-secondary | 2,935 | 64.8 | 1.095 | 1.109 | 11.0 | 799** | . 908 |
| University degree | 1,628 | 54.9 | 724** | $641 * *$ | 6.5 | 449** | .612** |
| Not stated | 418 | 53.9 | 696** | 786 | 11.5 | 840 | 708 |


| Income |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Low | 1,645 | 62.1 | . 975 | . 933 | 19.7 | 1.585** | 1.308** |
| Middie | 3,546 | 70.7 | 1.435** | $1.174 * *$ | 12.9 | 957 | . 985 |
| High | 1,258 | 66.8 | 1.197* | 1.121 | 7.4 | . $516 \times *$ | 768* |
| Not stated | 3,656 | 53.8 | $693 *$ | $814 * *$ | 132 | 983 | 1.011 |


| Drinking pattern |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Lifetime abstainer | 1,298 | 49.1 | .574** | .592** | 15.2 | 1.158 | . 922 |
| Former drinker | 1,397 | 60.2 | 900 | 838* | 19.2 | $1.536 * *$ | 1.200 |
| Light/infrequent | 3,486 | 64.7 | 1.090 | 1.077 | 15.0 | 1.140* | 1.002 |
| Light/frequent | 3,015 | 66.1 | $1.160^{*}$ | 996 | 8.0 | . $562^{* *}$ | $667 \times$ |
| Heavy/infrequent | 344 | 63.5 | 1.035 | 1.366* | 17.3 | 1.352 | 1.347 |
| Heavy/frequent | 565 | 69.4 | $1.349 \times 1$ | 1.375** | 12.0 | 881 | 1.004 |

Note: Weighted by ESSPROV - the weighting variable scaled down to produce the effective sample size *p < . $01 ; *$ p < 001

It can be seen that males are more likely than females to bet on lotteries, cards or sports ( $66.1 \%$ vs. $59.4 \%$ ) and that this relationship persists, when the other variables are taken into account. On the other hand, women are more likely than men to play bingo ( $18.5 \%$ vs. $8.2 \%$ ). The relationship between gender and playing bingo is relatively unaffected, when other variables are controlled.

Age is also significantly related to both types of gambling. Betting on lotteries, cards or sports increases among those aged 15 to 17 (28.6\%) to age 55 to 64 (70.6\%) and then declines. The magnitude and direction of this pattern generally remains the same, when the other predictors are taken into account. Playing bingo is most common among those aged 20 to 24 (15.4\%), 65 to 74 (16.1\%), and 75 or more (18\%). Indeed multivariate analysis indicates that it is people aged 20 to 29 years who are most likely to play bingo, once the other variables are taken into account.

Betting on lotteries, cards or sports is most commonly reported in Quebec (72.3\%) and least commonly in Ontario (55.3\%). This pattern persists in multivariate analysis. On the other hand, bingo is most commonly played in the Atlantic provinces (17.8\%) and British Columbia (15.3\%), and least often in Ontario (12.1\%) and Quebec ( $12.2 \%$ ). However, the relationship between region and playing bingo is much weaker, when other predictors are controlled. Then residents of Quebec have the lowest odds of playing bingo and residents of British Columbia have the highest.

Multivariate analysis clarifies the relationship between gambling and language spoken at home. Although a higher proportion of francophones appear to engage in lotteries, sports or cards, this relationship is no longer significant, once all other demographic variables are taken into account. As for bingo, anglophones and francophones do not appear to differ greatly. However, people who speak other than English or French at home are less likely to play bingo.

Respondents who have never married are somewhat less likely to bet on lotteries, cards or sports than those who are married or were once married. Widowed people are particularly likely to play bingo.

University graduates are less likely than others to bet on lotteries, cards or sports, or play bingo, and those without a post-secondary school education are more likely to do both. Indeed, a stronger and consistent relationship between education and gambling emerges, once the other factors are taken into account. It can be seen that the higher the respondent's educational level, the less likely he or she reports betting on lotteries, cards or sports, or playing bingo.

Middle income is associated with a higher probability of betting on lotteries, cards or sports. This association persists in the multivariate analysis. In contrast, income is inversely related to playing bingo; lower-income Canadians are much more likely to play bingo and those with high income are much less likely to do so.

Drinking patterns are associated with gambling but not in a simple manner Lifetime abstainers are the least likely to bet on lotteries, sports or cards, while heavy/frequent are most likely to do so. The association between heavy drinking and betting on lotteries, sports or cards continues to be significant after multivariate analysis. Although former drinkers are the most likely to play bingo (19.2\%), once age and other factors are taken into account in the multivariate analysis, this association is no longer significant. Light/frequent drinkers are the least likely to play bingo, a finding that persists when the other variables are controlled.

## CHAPTER 7

## PUBLIC OPINION

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Because public opinion is an important factor in policy formulation, CADS included questions about alcohol and drug-policy issues, which had been asked in the 1989 National Alcohol and Other Drugs Survey (NADS), as well as several additional questions on new or emerging issues.

Table P1 shows responses to a set of questions about alcohol-policy issues. On questions about the perceived seriousness of various alcohol-issues, it can be seen that impaired driving is most often perceived as a serious concern. A bout one fourth of respondents $(25.8 \%)$ view drinking and driving in their neighbourhoods as serious or very serious, another one fourth (24.7\%) view it as a problem but not a serious one, and $41.7 \%$ see it as not a problem. Alcohol-related public fights in their neighbourhood is seen as a problem by $39.8 \%$ of respondents, domestic vioIence by $37.8 \%$, and alcohol problems in the workplace by $21.3 \%$.

| Importance of Alcohol-related problems | Percent |
| :---: | :---: |
| Drinking and driving in the neighbourhood |  |
| Serious or very serious | 25.8 |
| A problem but not very serious | 24.7 |
| Not a problem | 41.7 |
| Not stated | 7.8 |
| A lcohol-related domestic violence |  |
| Serious or very serious | 18.0 |
| A problem but not very serious | 19.8 |
| Not a problem | 48.0 |
| Not stated | 14.2 |
| Alcohol-related public fights in the neighbourhood |  |
| Serious or very serious | 16.7 |
| A problem but not very serious | 23.1 |
| Not a problem | 53.4 |
| Not stated | 6.9 |
| Alcohol-related problems in the workplace |  |
| Serious or very serious | 7.8 |
| A problem but not very serious | 13.5 |
| Not a problem | 47.8 |
| Not applicable (e.g. homemaker) | 26.7 |
| Not stated | 4.2 |
| Access controls |  |
| Taxes on alcoholic beverages should be |  |
| Increased | 25.4 |
| Unchanged | 44.8 |
| Decreased | 25.4 |
| Not stated | 4.5 |

Alcohol Issues

Table P1
Public opinion on alcohol issues, importance of alcohol-related problems, access control, promotions and counter- promotions, and interventions

favouring an increase. Two thirds of Canadians (66.6\%) are also satisfied with the operating hours of beer and liquor outlets, with $16 \%$ favouring a decrease in hours as compared to only $10.9 \%$ favouring an increase. A similar majority ( $66.8 \%$ ) believes alcohol should not be sold in convenience stores.

There is continuing support, however, on a need for increased counteradvertising by government and for warning labels on alcoholic beverages ( $48.8 \%$ and $69.5 \%$ respectively), and strong support for more alcohol and drug-prevention programs ( $74.4 \%$ ), for preventing people who are drunk from being served ( $75.5 \%$ ), and for more treatment programs (64.6\%).

Although Canadians remain generally supportive of current alcohol-control policies and of increased prevention and treatment efforts, they are now somewhat less supportive both of control measures, and of more prevention and treatment than they were in 1989 as measured by the NADS.

|  | $19890 . . . . . . . . . . . . . . . . .$ |  | Difference |
| :---: | :---: | :---: | :---: |
| ACCESS CONTROLS |  |  |  |
| Taxes on alcoholic beverages should be |  |  |  |
| Increased | 27.0 | 25.4 | -1.6 |
| Unchanged | 46.1 | 44.8 | -1.3 |
| Decreased | 18.1 | 25.4 | +7.3 |
| Legal drinking age should be |  |  |  |
| Increased | 49.7 | 38.3 | -11.4 |
| Unchanged | 44.9 | 54.7 | +9.8 |
| Decreased | 2.8 | 4.1 | +1.3 |
| Alcohol outlet hours should be |  |  |  |
| Increased | 7.2 | 10.9 | +3.7 |
| Unchanged | 69.9 | 66.6 | -3.3 |
| Decreased | 17.3 | 16.0 | -1.3 |
| Alcohol should be sold in convenience stores |  |  |  |
| Yes | 23.4 | 30.0 | +6.6 |
| No | 73.6 | 66.8 | -6.8 |


| Increased | 61.1 | 48.8 | -12.3 |
| :---: | :---: | :---: | :---: |
| Unchanged | 28.0 | 34.4 | +6.4 |
| Decreased | 6.4 | 12.9 | +6.5 |
| Warning labels on alcoholic beverages |  |  |  |
| Yes | 74.4 | 69.5 | -4.9 |
| No | 22.5 | 27.6 | +5.1 |


| Interventions |  |  |  |
| :---: | :---: | :---: | :---: |
| Alcohol and drug-prevention programs should be |  |  |  |
| Increased | 81.0 | 74.4 | -6.6 |
| Unchanged | 12.8 | 18.0 | +5.2 |
| Decreased | 1.1 | 2.6 | +1.5 |

Trends in Public Opinion on Alcohol Issues

Trend in public opinion on alcohol issues: NADS, 1989 and CADS, 1994
Table P2
Trend in public opinion on alcohol
issues (cont'd)

Correlates of Opinions on Alcohol Issues

|  | Percent |  |  |
| :---: | :---: | :---: | :---: |
| Interventions |  |  |  |
| Preventing drunken people from being served should be |  |  |  |
| Increased | 82.1 | 75.5 | -6.6 |
| Unchanged | 10.1 | 15.3 | +5.2 |
| Decreased | 3.1 | 5.2 | +2.1 |
| Treatment programs should be |  |  |  |
| Increased | 74.1 | 64.6 | -9.5 |
| Unchanged | 13.6 | 24.2 | +10.6 |
| Decreased | 0.8 | 2.3 | +1.5 |

Note: Weighted by FINWGHT in CADS, and WEIGT in NADS.

Table P2 shows that, although support for lower alcohol taxes increased by $7.3 \%$ between 1989 and 1994, support for an increase in the legal drinking age declined by $\mathbf{1 1 . 4} \%$. Although most Canadians favour no change to the number of hours that alcohol outlets are permitted to be open, support for increasing the number of hours open for service rose by $3.7 \%$. Similarly, while the majority of Canadians remain opposed, support for the sale of alcohol in convenience stores has increased $6.6 \%$ while support for increased government advertising against alcohol declined by $12.3 \%$ and for warning labels on alcoholic beverages by $4.9 \%$. While most Canadians still favour increases in prevention and treatment programs and measures to prevent people who are drunk from being served, support for increasing these measures declined, while support for maintaining the current level rose.

Canadians who favour more restrictive alcohol policies, and increased government funding for prevention and treatment, tend to be females or older Canadians. As Table P3 illustrates, women are more likely than men to favour a higher drinking age, decreases in alcohol-store hours, more government advertising against drinking, more prevention and treatment programs, increased efforts to prevent service to people who are intoxicated, increases in alcohol taxes and warning labels on alcoholic beverages. Older people are more likely to favour increasing the drinking age, decreasing store hours, increasing alcohol taxes, and they are generally more often against alcohol sales in convenience stores. However, age is not clearly related to the other alcohol-policy issues, and is indeed, negatively related to support for increasing treatment programs; support for warning labels is highest in the youngest age category (15-17 years).

| Variable | Percent |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Favours higher drinking age | Favours fewer store hours | Favours more ads against drinking | Favours more alcohol prevention | Favours more to prevent drunks served | Favours more alcohol treatment | Favours higher alcohol taxes | Favours alcohol warning labels | A gainst alcohol in corner stores |
| Overall | 38.3 | 16.0 | 48.8 | 74.4 | 75.5 | 64.6 | 25.4 | 69.5 | 66.8 |
| Sex |  |  |  |  |  |  |  |  |  |
| Male | 33.3 | 12.4 | 42.2 | 71.3 | 72.1 | 59.7 | 20.8 | 63.8 | 57.1 |
| Female | 43.2 | 19.5 | 55.2 | 77.4 | 78.8 | 69.3 | 29.8 | 75.1 | 76.2 |
| Age |  |  |  |  |  |  |  |  |  |
| 15-17 | 10.9 | 12.6 | 46.3 | 66.1 | 66.2 | 68.4 | 22.6 | 78.9 | 62.5 |
| 18-19 | 14.8 | 10.00 | 40.8 | 75.7 | 71.4 | 72.7 | 19.1 | 68.1 | 69.1 |
| 20-24 | 24.1 | 9.6 | 51.9 | 80.5 | 77.0 | 68.6 | 21.4 | 66.6 | 67.3 |
| 25-34 | 35.6 | 12.5 | 53.0 | 78.4 | 78.3 | 66.1 | 21.6 | 68.6 | 65.7 |
| 35-44 | 41.7 | 16.2 | 51.6 | 77.0 | 77.9 | 65.3 | 26.9 | 67.8 | 66.3 |
| 45-54 | 46.0 | 18.4 | 49.3 | 74.2 | 78.0 | 66.1 | 28.8 | 68.3 | 64.6 |
| 55-64 | 46.1 | 18.9 | 43.4 | 69.7 | 71.5 | 60.8 | 26.1 | 70.5 | 64.2 |
| 65-74 | 47.2 | 20.8 | 42.8 | 69.7 | 72.4 | 57.7 | 26.7 | 74.7 | 72.6 |
| $75+$ | 49.0 | 27.5 | 42.0 | 62.6 | 69.0 | 54.2 | 34.8 | 70.6 | 77.4 |
| Province |  |  |  |  |  |  |  |  |  |
| Newfoundland | 42.8 | 17.3 | 57.2 | 84.9 | 83.9 | 76.6 | 31.3 | 87.4 | 64.4 |
| P.E.I. | 33.5 | 15.4 | 56.2 | 83.8 | 82.5 | 60.3 | 22.5 | 78.1 | 77.0 |
| Nova Scotia | 36.7 | 12.7 | 52.2 | 81.9 | 82.8 | 69.7 | 24.0 | 79.9 | 72.5 |
| New Brunswick | 34.5 | 19.6 | 55.3 | 79.4 | 76.3 | 69.7 | 27.1 | 83.6 | 60.7 |
| Quebec | 37.2 | 16.0 | 53.1 | 75.3 | 76.5 | 65.7 | 21.7 | 68.3 | 52.3 |
| Ontario | 37.4 | 11.2 | 44.5 | 69.3 | 69.0 | 62.1 | 23.0 | 65.5 | 69.1 |
| Manitoba | 45.9 | 18.6 | 45.8 | 76.5 | 82.1 | 66.0 | 28.6 | 69.9 | 82.3 |
| Saskatchewan | 39.0 | 17.0 | 46.0 | 76.4 | 79.2 | 62.8 | 34.4 | 75.3 | 78.7 |
| Alberta | 46.8 | 33.0 | 46.6 | 76.7 | 82.3 | 59.0 | 31.6 | 71.2 | 75.5 |
| B.C. | 35.5 | 17.0 | 52.3 | 80.2 | 81.4 | 69.4 | 31.0 | 72.5 | 74.3 |

Note: Percentages are based on the denominator of $23,030,000$, weighted by FINWGHT (expansion weight). Q Qualified release due to high sampling-variability

There are several notable variations among the provinces in public opinion on these issues. Support for a higher drinking age is highest in Alberta (46.8\%), Manitoba (45.9\%) and Newfoundland (42.8\%), and lowest in Prince Edward Island (33.5\%), New Brunswick (34.5\%) and British Columbia (35.5\%). Residents of Alberta are also the most likely to favour reduced hours for alcohol sales (33\%), and people in the Atlantic provinces are most likely to favour advertising against excessive drinking. However, most Canadians support prevention and treatment programs, increased efforts to prevent serving intoxicated people, and alcohol-warning labels. Respondents in Newfoundland and the western provinces were most likely to support higher alcohol taxes, but there is variation concerning alcohol sales in convenience stores. Although a majority of respondents in all provinces are against such sales, opposition is much lower in Quebec (52.3\%) than elsewhere, possibly reflecting the fact that beer and wine have long been available in convenience stores in that province.

Table P4 Public opinion on cannabis policy, by sex, age, and province

| Variable/category | Percent |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Possession of cannabis should not be against the law | Possession of small amounts of cannabis should not be subject to a jail sentence | Possession of small amounts of cannabis should be subject to a potential jail sentence | No opinion |
| Overall | 27.0 | 42.1 | 16.8 | 14.0 |
| Sex |  |  |  |  |
| Male | 33.4 | 38.7 | 14.7 | 13.1 |
| Female | 20.8 | 45.4 | 18.9 | 14.9 |
| A ge |  |  |  |  |
| 15-17 | 31.6 | 39.7 | 21.4 | 7.30 |
| 18-19 | 36.0 | 36.9 | 21.7 | 5.40 |
| 20-24 | 31.0 | 42.1 | 20.7 | 6.2 |
| 25-34 | 29.8 | 42.6 | 17.5 | 10.1 |
| 35-44 | 32.2 | 42.0 | 14.5 | 11.4 |
| 45-54 | 26.3 | 41.6 | 16.0 | 16.1 |
| 55-64 | 20.5 | 44.1 | 15.2 | 20.3 |
| 65-74 | 14.6 | 43.8 | 17.5 | 24.1 |
| 75+ | 14.5 | 40.7 | 13.5 | 31.4 |
| Province |  |  |  |  |
| Newfoundland | 16.3 | 44.2 | 30.6 | 8.9 |
| P.E.I. | 16.9 | 47.8 | 23.6 | 11.8 |
| Nova Scotia | 23.2 | 46.6 | 19.6 | 10.6 |
| New Brunswick | 21.3 | 46.3 | 24.0 | 8.4 |
| Quebec | 28.9 | 45.8 | 15.6 | 9.7 |
| Ontario | 24.7 | 38.2 | 17.1 | 20.0 |
| Manitoba | 23.2 | 47.2 | 19.2 | 10.4 |
| Saskatchewan | 18.1 | 52.9 | 20.8 | 8.2 |
| Alberta | 28.2 | 44.9 | 18.0 | 8.9 |
| B.C. | 36.9 | 37.5 | 11.5 | 14.1 |

Note: Percentages are based on the denominator of $23,030,000$, weighted by FINWGHT (expansion weight). Q Qualified release due to high sampling-variability

Table P4 shows the relationships between opinion on cannabis policy and gender, age and province. Overall, $27 \%$ of respondents believe possession of cannabis should be legal, $42.1 \%$ believe it should be against the law but subject to either no penalty or a fine only for a first offence, and $16.8 \%$ that it should be subject to a potential jail sentence for a first offence. The remaining 14\% express no opinion.

Males are more likely than females to believe possession of cannabis for personal use should not be against the law (33.4\% vs. $20.8 \%$ ), and females are more likely than males to favour potential jail sentences for possession (18.9\% vs. $14.7 \%)$. This finding is not surprising, as males are more likely than females to use cannabis.

Unlike gender, age does not relate to support for more liberal cannabis policy in quite the same manner as it does to rates of use. Just as younger Canadians are more likely to use cannabis, Canadians under the age of 45 are more likely to sup-
port the position that possession of cannabis should not be against the law.
However, the proportion who favour potential jail sentences for cannabis possession also decreases with age. Perhaps the most noteworthy pattern with regard to age is that, in every age group, most people with any opinion on cannabis policy favour the middle option, where possession is against the law but not subject to a jail sentence.

There is considerable provincial variation in attitudes. Support for the position that cannabis possession should not be against the law is highest in British Columbia (36.9\%), and lowest in Newfoundland (16.3\%), Prince Edward Island (16.9\%), and Saskatchewan (18.1\%). Similarly, support for the policy that possession of cannabis be subject to a potential jail sentence is strongest in Newfoundland (30.6\%) and weakest in British Columbia (11.5\%).

APPENDICES

The following table summarizes the hit rates, response rates and the final sample size for CADS

| Province | Strata | Telephone Numbers Generated | Supplement Numbers Generated | Total <br> Telephone <br> Numbers <br> Generated | Hit Rate | Response Rate | $\begin{aligned} & \text { Sample } \\ & \text { Size } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Newfoundland | CMA | 540 | 0 | 540 | 44.2\% | 86.6\% | 206 |
|  | Non-CMA | 1,246 | 0 | 1,246 | 42.0\% | 87.9\% | 459 |
| Prince Edward Island |  | 855 | 0 | 855 | 44.0\% | 83.2\% | 313 |
| Nova Scotia | CMA | 948 | 0 | 948 | 48.6\% | 74.8\% | 344 |
|  | Non-CMA | 1,243 | 0 | 1,243 | 50.0\% | 85.8\% | 533 |
| New Brunswick | CMA | 393 | 0 | 393 | 34.1\% | 88.8\% | 119 |
|  | Non-CMA | 1,530 | 170 | 1,700 | 44.4\% | 81.9\% | 617 |
| Quebec | M ontreal | 2,307 | 193 | 2,500 | 55.2\% | 72.8\% | 1,000 |
|  | Other CMA | 784 | 60 | 844 | 55.0\% | 75.6\% | 351 |
|  | Non-CM A | 4,686 | 0 | 4,686 | 22.8\% | 81.8\% | 874 |
| Ontario | Toronto | 2,588 | 830 | 3,418 | 48.4\% | 60.4\% | 997 |
|  | Other CMA | 1,872 | 170 | 2,042 | 57.5\% | 72.5\% | 851 |
|  | Non-CMA | 2,788 | 302 | 3,090 | 41.7\% | 70.3\% | 905 |
| M anitoba | CMA | 1,262 | 200 | 1,462 | 49.8\% | 72.4\% | 527 |
|  | Non-CMA | 968 | 0 | 968 | 47.0\% | 75.4\% | 343 |
| Saskatchewan | CMA | 683 | 65 | 748 | 51.3\% | 81.3\% | 312 |
|  | Non-CMA | 1,492 | 0 | 1,492 | 44.4\% | 79.9\% | 529 |
| Alberta | CMA | 1,784 | 0 | 1,784 | 57.5\% | 76.7\% | 786 |
|  | Non-CMA | 1,433 | 0 | 1,433 | 49.6\% | 80.2\% | 570 |
| British Columbia | CMA | 1,976 | 190 | 2,166 | 55.9\% | 72.8\% | 881 |
|  | Non-CM A | 1,356 | 0 | 1,356 | 58.0\% | 81.3\% | 638 |
| TOTAL |  | 32,734 | 2,180 | 34,914 | 46.1\% | 75.6\% | 12,155 |

Source: Statistics Canada (1994). Microdata User's Guide: Canada's Alcohol and Other Drugs Survey.

The table below shows the design effects, sample sizes and population counts by province, which were used to produce the approximate sampling-variability tables.

| Province | Design effect | Sample size | Population |
| :---: | :---: | :---: | :---: |
| Newfoundland | 1.15 | 665 | 457,961 |
| Prince Edward Island | 1.02 | 313 | 103,920 |
| Nova Scotia | 1.10 | 877 | 742,975 |
| New Brunswick | 1.12 | 736 | 602,504 |
| Quebec | 1.12 | 2,225 | 5,795,927 |
| Ontario | 1.20 | 2,753 | 8,672,981 |
| Manitoba | 1.11 | 870 | 874,366 |
| Saskatchewan | 1.15 | 841 | 767,332 |
| A lberta | 1.15 | 1,356 | 2,073,112 |
| British Columbia | 1.14 | 1,519 | 2,938,661 |
| Atlantic Provinces | 1.17 | 2,591 | 1,907,360 |
| Prairies | 1.2 | 3,067 | 3,714,810 |
| Canada | 1.43 | 12,155 | 23,029,739 |

Source: Statistics Canada (1994). Microdata User's Guide: Canada's Alcohol and Other Drugs Survey.

Appendix A
Sample sizes and response rates

Appendix B Design effects

| Appendix C C.V. release guidelines | Type of estimate | C.V. (in\%) | Guidelines |
| :---: | :---: | :---: | :---: |
|  | Unqualified | 0.0-16.5 | Estimates can be considered for general unrestricted release. Requires no special notation. |
|  | Qualified | 16.6-25.0 | Estimates can be considered for general unrestricted release but should be accompanied by a warning cautioning subsequent users of the high sampling-variability associated with the estimates. Such estimates should be identified by the letter Q (or in some other similar fashion). |
|  | Confidential | 25.1-33.3 | Estimates can be considered for general unrestricted release, only when sampling variabilities are obtained using an exact variance calculation procedure. <br> Unless exact variances are obtained, such estimates should be deleted and replaced by dashes in statistical tables. |
|  | Not for release | 33.4 or greater | Estimates cannot be released in any form under any release OR circumstances. In statistical tables, such estimates must be deleted and replaced by dashes. |

Source: Statistics Canada (1994). Microdata User's Guide: Canada's Alcohol and Other Drugs Survey.


[^0]:    a The SPSS command for computing WT12155 is: compute WT12155 = FINWGHT/1894.7.
    b The SPSS command for computing ESSPROV for Ontario with an average design effect of 1.20, for example, is: if DVPROV $=5$ ESSPROV $=$ WT12155/1.20.

[^1]:    ${ }^{\text {a }}$ Classification was originally developed by the Yukon Bureau of Statistics.
    ${ }^{\text {b }}$ Survey questions were not identical. In NADS: "Did you ever drink alcohol beverages regularly?" and if "No", followed by "Does this mean you never had a drink?" In CADS "Have you ever had any alcohol?"

[^2]:    Note: Weighted by FINWGHT (expansion weight).
    a Excludes "disturbed by loud parties or behaviour of people drinking"
    Q Qualified release due to high sampling-variability

[^3]:    ${ }^{\text {a }}$ The rate appeared to be much higher for women than men. However, such estimates are based on unacceptably low numbers of respondents according to Statistics Canada's guidelines. Further investigations are warranted.

