

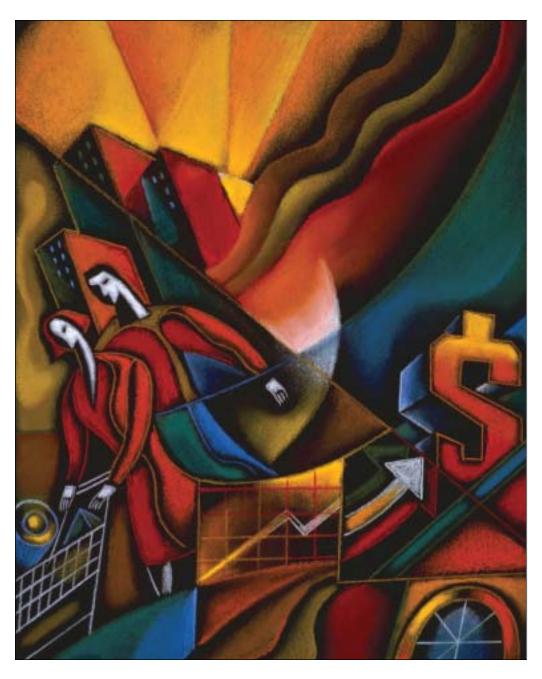
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# PERSPECTIVES

ON LABOUR AND INCOME

# **DECEMBER 2001** Vol. 2, No. 12

- THE MALE-FEMALE WAGE GAP
- PRIVATE PENSION SAVINGS, 1999
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# **Highlights**

## In this issue

## The male-female wage gap

- In 1997, the average annual earnings of women working full-year, full-time were 73% of men's. Using the average hourly wages of all employees, the ratio was 80%. This sizeable difference results mainly from differences in the study population and differences in work volume.
- Male-female differences in labour market experience, education and major field of study, occupation, job responsibilities, and industry can explain half of the wage gap. Actual labour market experience and major field of study, rather than proxy measures of experience and education, further help to understand the issue.
- Most studies examine the *average* pay differential and assume that the size and components of the gap are constant along the whole wage distribution. However, about 47% of the difference in pay at the 90<sup>th</sup> percentile and 57% at the 25<sup>th</sup> percentile are explained by differences in observable characteristics. This suggests that the 50% of the wage gap due to men and women possessing different wage characteristics at the mean fails to accurately represent the differences encountered along the wage distribution.
- Questions related to pay differentials are often framed in a manner that examines the extent to which women are paid the same as comparable men. Adopting alternative comparative wage structures can lead to quite different interpretations of the components of male-female pay differentials. For example, the portion of the gap attributable to differences in worker characteristics can vary from 6% to 61%.

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# The male-female wage gap

## Marie Drolet

HE ISSUE OF male-female wage inequality is complex and requires analysis from a number of different perspectives. The goal of this article is *not* to provide a single, definitive estimate of the wage gap, but to demonstrate the importance of measurement, decomposition techniques, and differences in the gap within the distribution of wages.

#### **Measures of earnings**

For three decades, the Survey of Consumer Finances (SCF) was the primary source of data on the earnings of Canadians. As of 1997, the Survey of Labour and Income Dynamics (SLID) replaced the SCF. According to the SCF, among full-year, full-time (FYFT) workers, women working full-year, fulltime earned 72.5% of what men earned (Table 1). Using the 1997 SLID, the ratios varied from 61.6% (using annual wages and salaries for all employees in their main job) to 68.3% (using annual wages and salaries for full-year, full-time employees in their main job) to 80.4% (using the hourly wage rate for all employees in all jobs).

# Why the large differences in earnings ratios?

The SLID ratio, based on the hourly wages of all employees, was eight percentage points greater than

Table 1: Female-male earnings ratios, 1997

			Ea	arnings	Female-
Source	Measure	Coverage	Men	Women	male ratio
				\$	%
SCF	Annual	FYFT workers	42,600	30,900	72.5
SLID	Annual	FYFT employees	44,800	30,600	68.3
SCF	Annual	All workers	33,200	21,200	63.8
SLID	Annual	All employees	38,900	24,000	61.6
SLID	Hourly	All employees	18.81	15.12	80.4
SLID	Hourly	FT employees	18.92	15.31	80.9

Sources: Survey of Consumer Finances; Survey of Labour and Income Dynamics

the SCF ratio for full-year, full-time workers. This sizeable disparity arises from several conceptual differences between SLID and the SCF:

First, the ratios are for different populations of workers. The SCF ratio considers both employees and the self-employed, while the SLID ratio reflects only employees. Analyses based solely on FYFT workers tend to exclude a significant portion of the labour force—especially women. According to 1997 SLID data, roughly 76% of men and 60% of women worked full year, full time. Ratios based on hourly wage rates avoid these definition problems and are representative of female-male pay differences.

Second, the definition of earnings is different for each survey. SCF annual earnings include wages and salaries from all jobs, and net income from self-employment. SLID annual earnings include only wages and salaries. Hourly wage rates are job-specific, thus facilitating comparisons between the wages of men and women in comparable jobs.

Third, pay ratios based on annual earnings do not accurately account for differences in work volume. Even among men and women working full year, full time, the number of hours worked per week varies considerably. According to the Labour Force Survey, men employed full time in 1997 worked 43.1 hours per week, while women worked 39.0 hours per week. Ratios based on hourly wage rates overcome this problem.

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#### Data source

The data are from the 1997 Survey of Labour and Income Dynamics (SLID). SLID is designed to capture the nature and the pattern of individuals' labour market experiences from the types of jobs they hold, their earnings and selected employer characteristics. However, to get a full picture of economic well-being, SLID also collects information about income, using categories similar to those used by the Survey of Consumer Finances.

Since January 1993, SLID has conducted annual cycles of labour and income interviews with 15,000 to 20,000 households across the country. A new panel is introduced

every three years (the second, in 1996) and remains in the survey for six years.

This study looks at employees, aged 18 to 64, who are not enrolled as full-time students during the year. Self-employed individuals, with or without an incorporated business; and individuals with missing data on hourly wages, full-year full-time work experience, years of schooling, or total usual hours worked in the job during the year are excluded. The analysis is based on 28,303 jobs—14,604 held by men and 13,699 held by women.

The remainder of this article focuses on 1997, when women earned on average \$15.12 per hour and men earned \$18.81. In other words, women earned about 80% of the average male hourly wage.<sup>1</sup>

#### Measures of experience

A major limitation of previous studies is the lack of sufficient information on the determinants of the wage gap. One such determinant is the quantity of lifetime work experience. Often age or the Mincer measure of experience (age – schooling – 6) is used as a proxy for the acquisition of general human capital skills or for potential work experience. However, proxy measures tend to overstate the *actual* work experience of women by not accounting for interruptions related to parenting (that is, complete withdrawals from the labour market) or for any restrictions on the number of hours worked per week or the number of weeks worked per year.

SLID collects information on work experience starting with the first full-time paid job. It asks respondents the years in which they worked at least six months (full year), as well as the years in which they did not work at all, since starting to work full time. The remaining years are coded as working part year. Respondents are then asked the years in which they worked six months or more, and if they worked full time (30 or more hours per week), part time, or both. Part-time or summer jobs while in school are not included. From the information provided, a measure of full-year, full-time equivalent (FYFTE) work experience is calculated:

FYFTE = years worked full year, full time

- + 0.5 (years worked full year, part time)
- + 0.5 (years worked full year; some full time, some part time)
- + 0.25 (years worked part year, part time *or* part year, full time)

Men and women bring different work experience to the labour market (Table 2). The Mincer proxy for potential work experience shows little difference in the work experience of men and women (19.5 and 19.1 years respectively). A different story emerges when actual labour market experience is applied. The average FYFTE work experience of men is 18.3 years

Table 2: Differences in work experience, by sex

	Men	Women
	`	/ears
Average age	39.2	39.0
Average years of schooling	13.6	13.8
Mincer proxy for experience: Average potential years of work experience (age – schooling – 6)	19.5	19.1
Actual labour market experience: Average years of full-year, full-time equivalent work experience (FYFTE)	18.3	14.4
Proportion of potential years of work experience spent working full year,		%
full time	93.8	75.4 ———

Source: Survey of Labour and Income Dynamics, 1997

compared with 14.4 for women. Men also spend a greater proportion of their potential years of work experience working full year, full time (94% versus 75% for women).

In absolute terms, the actual experience of young women is similar to that of their male counterparts but, with age, the gap widens (Table 3). This may be partially because older women (aged 55 to 64 in 1997) were less inclined to combine work and family than later cohorts (age 25 to 34 in 1997). As well, young men and women are new to the labour market and have not yet experienced the effects of career interruptions.

The SLID measure of experience, although a welcome addition to the study of pay differentials, is far from perfect. Information is missing on the *continuity* of work experience, the *duration* of labour force withdrawals, and the *frequency* and *timing* of withdrawal. These factors may influence the pay women receive in several ways (Corcoran and Duncan, 1979).

First, men and women differ considerably in the amount of time they work and in the continuity of their work experience. Women are more likely to combine periods of paid work with periods of labour force withdrawal for family-related reasons. This affects job tenure—a factor that influences wages.

Second, human capital skills may depreciate during long periods of labour force withdrawal. Women returning to the same employer after an interruption in employment may be less likely to be promoted. Or, women not returning to the same employer may have to accept lower wages than they received prior to their withdrawal.

Third, women expecting several withdrawals from the labour force may postpone training, or may decide to accept low-wage jobs in industries or occupations that are easy to enter and exit.

Fourth, the timing of labour force withdrawals may affect wages. Job-related skills are usually acquired at the start of careers—which generally coincides with

decisions regarding children. A significant portion of real lifetime earnings growth has been found to occur during the first years after graduation (Murphy and Welch, 1990). If so, the timing of labour force withdrawals may have important long-term implications for future earnings patterns.

Using definitions of full year, full time to calculate FYFTE work experience may be problematic. Full year refers to working at least six months in a calendar year and full time refers to working at least 30 hours per week. As noted earlier, among full-time workers, the number of hours usually worked per week varies considerably. For example, Person A works 45 hours per week for 12 months, while Person B works 32 hours per week for 7 months. Under the SLID measure of FYFTE, each person would have one year of FYFTE experience despite the significant variation in work hours. As well, the measure of FYFTE work experience does not account for short-term (less than six months) labour force withdrawals in a given year. Almost 40% of working women take less than six months off work after giving birth (Marshall, 1999). If these unrecorded work interruptions and hours of work become more important as experience increases, then the measure of FYFTE work experience may become less accurate in reflecting the relative amount of work done by men and women.

The unadjusted wage gap is small for workers with less than two years of experience (4%) but grows larger as years of work experience increase.<sup>2</sup> A study of recent post-secondary graduates from the

Table 3: Actual versus potential work experience

		Men				Women		Female/ Male
	Potential	Actual	Ratio	Ī	Potential	Actual	Ratio	Actual experience
Age	Y	ears	%		Y	ears	%	%
18 to 64 18 to 24 25 to 34 35 to 44 45 to 54 55 to 64	19.5 3.4 10.0 19.4 29.6 40.2	18.3 2.9 9.2 18.2 27.9 37.1	93.8 85.3 92.0 93.8 94.3 92.3		19.1 2.8 9.4 19.6 29.2 40.2	14.4 2.4 7.9 15.6 21.5 23.8	75.4 85.7 84.0 79.6 73.6 59.2	78.7 82.8 85.9 85.7 77.1 64.2

Source: Survey of Labour and Income Dynamics, 1997

Note: Ratio refers to the proportion of potential work experience spent working full year, full time.

National Graduates Survey found that the gender pay gap was relatively small two years after graduation (7%) but widened two to five years after graduation (16%) (Finnie and Wannell, 1999): "These findings have interesting implications for the longer-term earnings profiles of graduates as they suggest that longer run (permanent) reductions in the earnings gap amongst cohorts of graduates might not be nearly as great as immediate post-graduation records suggest."

#### **Education and major field of study**

Canadian women have increased their educational attainment, matching and at some levels surpassing men. The importance of human capital characteristics—notably education—in the wage determination process has been firmly established in the economics literature. SLID contains information on educational attainment as well as major field of study. Not surprisingly, many fields of study are dominated by one of the sexes. For instance, graduates of engineering, applied sciences, technologies, and trades fields are mostly men. On the other hand, women are overrepresented among graduates with a commerce or business administration degree, and in the health and education fields. Since wages differ by major field of study, the choices made by men and women may account for some of the pay gap.

#### **Explaining the gap**

In describing the causes of labour market differentials, economists look first at the attributes different workers bring to the workplace. There is no universally accepted set of conditioning variables that should be included. However, the consensus is that controls for productivity-related factors such as education (level and major field of study), experience, experience squared, job tenure, marital status, presence of children, part-time status, union status, firm size, region, and urban size should be included. It is debatable whether occupation and industry should be included. If employers differentiate between men and women through their tendency to hire into certain occupations, then occupational assignment is an outcome of employer practices rather than an outcome of individual choice (Altonji and Blank, 1999). Analyses that omit occupation and industry may overlook the importance of background and choice-based characteristics on wage outcomes, while analyses that fully control for these variables may undervalue the significance of labour market constraints on wage outcomes. (see Accounting for male-female pay differences).

Using actual rather than potential experience allows *more* of the wage gap to be explained (Table 4). The explained component is about 9.2% to 37.4% in the conventional model (using potential work experience) and 29.2% to 49.3% in the augmented model (using actual work experience). This finding is novel for Canada but similar to findings in other countries (for example, Wright and Ermisch, 1991, for the United Kingdom; O'Neill and Polachek, 1993, for the United States).

The portion of the wage gap attributable to differences in work experience between men and women is severely underestimated when potential instead of actual labour market experience is used. Differences in actual work experience account for at most 12% of the gap, while potential work experience explains less than 1%. This may be explained as follows: First, as stated earlier, men and women differ little in the mean characteristics of potential experience but they differ

Table 4: Accounting for male-female differences in hourly earnings

	Fraction of the gap explained by				
	Conver	ntional	Augm	ented	
	A*	B**	A*	B**	
		%	·		
Mincer proxy Education level	0.5 -4.9	0.6 -3.3			
Actual experience Education level/major field of study			11.9 5.4	10.8 4.1	
Job tenure Part-time status Union status Firm size Marital status Age of youngest child Region Urban size	3.3 8.4 0.5 1.0 1.0 0.0 -0.3 -0.3	2.6 2.9 0.8 0.8 0.9 0.3 -0.2	2.7 6.8 0.5 1.1 1.1 0.4 -0.3	2.2 2.2 0.8 0.8 1.0 0.5 -0.2	
Occupation Job responsibilities Industry		9.2 6.2 17.0		6.8 5.7 15.1	
Explained Unexplained	9.2 90.8	37.4 62.6	29.2 70.8	49.3 50.7	

Source: Survey of Labour and Income Dynamics, 1997

<sup>\*</sup> Ignoring industry and occupation.

<sup>\*\*</sup> Controlling for industry and occupation.

significantly in actual FYFTE work experience.<sup>3</sup> Second, although potential and actual work experience are highly correlated, an additional year of FYFTE experience gives greater returns than a year of potential work experience. So, when FYFTE experience is used, both the difference in means and the difference in returns produce a greater explained component than when potential experience is used.

Differences in men's and women's educational attainment reduce the explained component by at most 5%, while differences in the major field of study by educational attainment account for no more than 5% of the wage gap. The contribution of each major field of study to the wage gap varies considerably. About 15% can be explained because men are more likely to graduate from engineering and applied sciences programs, and degrees in these programs yield high returns. However, the prevalence of women graduating from health science and education programs—occupations associated with high earnings—reduces the explained component by 5% to 9%.

Despite the long list of productivity-related factors used in this study, a substantial portion of the wage gap cannot be explained. The same results were found in other notable Canadian studies (Baker et al., 1995; Gunderson, 1998; Christofides and Swidinsky, 1994). Such large, unexplained differences may be related to productivity-related factors, labour market decisions, or skill measures that are not captured by SLID.

# The choice of non-discriminatory wage structure matters

Questions related to pay differentials are often framed in a manner that examines the extent to which women are paid the same as comparable men. For this reason, the male wage structure is often considered the 'competitive' or the 'non-discriminatory' wage structure. Alternative competitive wage structures may be employed. If the female pay structure is used, the question would be: What is the hypothetical wage men would receive if they were paid according to the female wage structure? Other approaches hint that the competitive norm falls somewhere between the male and female wage structure. One suggestion is that the competitive wage structure should be more similar to the structure for the larger group (Cotton, 1988). That is, the returns are estimated as the weighted average of the male and female coefficients, where the weights are simply the proportion of the total population that are men and women. Another suggestion is that the least squares estimates from a combined or *pooled* model be used as the estimate of the competitive wage structure (Neumark, 1988).

When the male wage structure is adopted as the competitive standard, 51% of pay differentials result from differences in the returns to wage-determining characteristics, while 49% result from differences in the endowments of wage-determining characteristics. When the female wage structure is adopted, malefemale differences in characteristics explain 6%, while differences in the returns to wage-determining characteristics account for 94% of the pay differential.

By construction, the weighted average method yields decomposition techniques that are bounded by those obtained from the separate wage structures of men and women. This method yields an explained component of 30% and an unexplained component of 70%. Finally, the pooled method attributes a smaller proportion (39%) of the wage differential to differences in the returns to wage-determining characteristics, and a larger proportion to gender differences in characteristics, than the other three competitive wage structures.

These results suggest that adopting alternative comparative pay structures can lead to quite different interpretations of the components of male-female pay differentials.

#### The pay gap differs within the wage distribution

A major drawback in the preceding methodology is that it considers only the information of the *average* pay differential, and it assumes that the size of the wage gap and its components are constant along the whole wage distribution.

In relation to the absolute predicted wage gap based on the attributes of the typical worker, about 47% of the difference in pay at the 90th percentile compared with 57% at the 25th percentile was explained by malefemale differences in observable characteristics. This suggests that the 50% of the wage gap due to men and women possessing different wage-determining characteristics at the mean fails to accurately represent the differences encountered along the wage distribution. As well, the contribution of relevant factors in explaining pay differentials varies throughout the wage distribution. Work experience, job responsibilities and industry contribute more to understanding pay differentials at the upper end of the distribution relative to the

lower end, while occupation, union status and parttime status explain more at the lower end of the distribution relative to the upper end.<sup>5</sup>

#### Other factors may contribute

The regression models used in the decomposition analysis account for no more than half of the variation in the hourly wages of men and women. The fit of the model to the SLID data can be improved by including other variables deemed to influence wages. The data typically used come from household surveys. Researchers have been unable to document the potential effect of firm characteristics—other than industry and firm size—on the wages of men and women. The new Workplace and Employee Survey will allow researchers to move beyond the individual worker to consider the importance of the workplace. But as noted earlier, pay differentials can be explained only if the factor in question influences wages and differences exist in the distribution of the factor among men and women.

People differ in their preference for particular types of work (that is, paid work or self-employment, hours, location, responsibilities). Differences between men and women in the labour market may reflect genuine differences in preferences, pre-labour market experiences, expectations, or opportunities. It is therefore difficult to distinguish between choice-based decisions and differential treatment based on sex. For example, pre-market experiences—which are related to expectations based on sex, both at home and in the education system—may influence the level of educational attainment and the choice of major field of study, labour force participation, job selection, and work habits.

Much of the literature has emphasized the importance of imperfect information about worker attributes. Employers are constantly making decisions regarding hiring and promotion and may use sex to predict future work commitment. Some firms may hesitate to hire women because women have, on average, more career interruptions and more absences for family reasons than men. Consequently, statistical discrimination may result (Arrow, 1973). This would be especially true in firms with substantial hiring or training costs, or where wages are higher. Insofar as employer behaviour changes as a result of women's increased

participation in the labour force, women's access to high-paying jobs will improve due to decreases in statistical discrimination.

#### **Summary**

While earnings differences between men and women have narrowed since the 1970s, they continue to be remarkably persistent. Measurement and methodological issues play important roles in studying these differences.

Estimates of the pay ratio are sensitive to the measurements used. Conceptual differences in data sources and earnings measures that do not fully account for differences in work volume lead to different estimates of the ratio. The portion of the wage gap attributable to differences in work experience between men and women is severely underestimated when proxies for experience are used instead of actual labour market experience. Education level plays little role in explaining wage differentials, but choice of major field of study furthers our understanding of the issue.

Conclusions are often sensitive to the methodology chosen by the researcher. Questions are often framed in a manner that examines the extent to which women are paid the same as comparable men. Most studies use the male pay structure to make comparisons. However, alternative pay structures can be employed, with estimates of the unexplained portion of the differential ranging from 39% to 94%.

One drawback is that only the average difference is considered. When the size of the wage gap and its components along the whole wage distribution are examined, a different story emerges.

#### Perspectives

#### Notes

- 1 In terms of mean-log wages, the female-male hourly wage differential is 0.220.
- 2 The female-male hourly wage ratios are 0.96 for 0 to 2 years of actual FYFTE work experience, 0.87 for 3 to 5 years, 0.85 for 6 to 9 years, 0.82 for 10 to 19 and 20 to 29 years, and 0.80 for 30 or more years.
- 3 Men have 0.4 years more potential and 3.9 years more actual years of work experience than women.

- 4 The contribution of each major field of study to the female-male wage gap is calculated using results from the augmented model including detailed occupation and industry.
- 5 Work experience, job responsibilities and industry explain 14%, 13% and 29% of pay differences between men and women at the 90th percentile compared with 12%, 4% and 15% at the 25th percentile. Occupation, union status, and part-time status contribute 6%, 5% and 7% to pay differentials at the 25th percentile compared with the 90th percentile, where these factors tend to reduce the explained component by -15%, -3% and -4% respectively. The experiment is best framed using the characteristics of the typical worker around the percentile since the characteristics of workers at the upper end of the distribution differ from those of the average worker or workers at the lower end of the wage distribution.

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## Appendix: Accounting for male-female pay differences

The wage structures of men and women were examined by estimating the relationship between hourly wage rates and observed characteristics in semi-logarithmic form:

$$\ln w_i = X_i \beta + u_i \quad i = m, f \tag{1}$$

where the natural logarithm of wages is used as the dependent variable,  $X_i$  is a vector of worker and employer characteristics,  $\beta$  is a vector of coefficients measuring the returns to those characteristics, and u is the error term. Each coefficient is the percentage change in hourly wage rates associated with a one-unit change in the explanatory variable.

From the estimated regressions, the difference in the mean log wages between men and women is decomposed into three terms (Table A1):

$$\overline{\ln w_m} - \ln \overline{w}_f = (\overline{X}_m - \overline{X}_f) \beta^* + \overline{X}_m (\beta_m - \beta^*) + \overline{X}_f (\beta^* - \beta_f)$$
 (2)

where  $\hat{\boldsymbol{\beta}}^*$  is the choice of comparative wage structure. The choice of comparative wage structure is examined in the section *The choice of non-discriminatory wage structure matters*. The first term represents the explained portion, which includes male-female differences in worker characteristics  $\overline{X}_m - \overline{X}_f$  evaluated at the competitive wage structure  $\hat{\boldsymbol{\beta}}^*$ . The residual or unexplained component is the proportion of the gap due to differences in the returns to

Table A1: Decomposition of male-female wage differences

	$oldsymbol{eta}^*$ value		wage ial = 0.220
Structure	(Equation 2)	Explained <sup>†</sup>	Unexplained <sup>††</sup>
Male	$oldsymbol{eta}_m$	0.108 49.3%	0.112 50.7%
Female	$oldsymbol{eta}_{\scriptscriptstyle f}$	0.013 5.9%	0.206 93.6%
Weighted	$\beta_m p_m + \beta_f p_f$	0.067 30.5%	0.153 69.5%
Pooled	$\widetilde{oldsymbol{eta}}$	0.135 61.1%	0.086 39.1%

Source: Survey of Labour and Income Dynamics, 1997.

wage-determining characteristics and consists of a male advantage (second term) and a female disadvantage (third term). This decomposition is made possible by the ordinary least squares (OLS) property that the sample average wage,  $\overline{w}$ , is equal to the product of the average vector of characteristics  $\overline{X}$  and the estimated regression coefficients  $\hat{\beta}$ .

Previous work has shown that men and women differ in ways that may affect their productivity, but these differences do not necessarily explain the pay gap. It can only be explained if the differences in the factors in question are themselves important determinants of the pay received. The wage gap between men and women may be the result of either differences in their productivity-related characteristics or differences in the compensations they receive for the same productivity-related characteristics.

Quantile regression methods reveal dispersions in the wage differential that cannot be captured by OLS models. The properties of the OLS model used above (Equation 1) ensure that the average wage,  $\overline{w}$ , is equal to the sample average characteristics  $\overline{X}$  evaluated at the estimated regression coefficients  $\hat{\beta}$ . However, the quantile regression model does not have a comparable property and as a result, no exact decomposition is possible. The difference in the log wage between men and women can be formulated (following the notation of Mueller, 1998) as:

$$\ln w_m^q - \ln w_i^q = \beta_m^q (\overline{X}_m - \overline{X}_f) + \overline{X}_f (\beta_m^q - \beta_f^q)$$
(3)

where  $\ln w_i^q$  is the natural logarithm of hourly wages for sex i evaluated at quantile q,  $\beta_i^q$  is a vector of estimated coefficients for sex i evaluated at quantile q,  $\overline{X_i}$  is a vector of average characteristics of worker i (i = m, f denotes male and female respectively and q = 0.10, 0.25, 0.50, 0.75 and 0.90).

Two sets of predicted values can be calculated from the coefficients of the quantile regression model of the log hourly wages. The first is conditioned on a vector of characteristics of men and women around the percentile, while the second is conditioned on the vector of average characteristics of all men and women (Table A2). To calculate the predicted wages of men and women at the 25th percentile, the first set of conditioning variables includes the characteristics of individuals falling within the 20th – 30th percentile. The second set of conditioning variables used to calculate the predicted wages is simply the average characteristics of all men and women respectively.

For each percentile, both the actual and predicted wages are always greater for men than for women. The predicted wage gap can deviate from the actual wage gap. For instance, the actual wage gap at the 10th percentile is about

<sup>&</sup>lt;sup>†</sup> Due to differences in wage-determining characteristics.

<sup>&</sup>lt;sup>††</sup> Due to differences in returns to wage-determining characteristics.

## Appendix: Accounting for male-female pay differences (concluded)

23%. But the predicted wage gap is 18% when the characteristics of the *typical* worker are used, and 19% when the average characteristics are used. Using the characteristics of the *average* worker, the predicted wage gap

increases along the wage distribution. However, a similar conclusion cannot be reached when the predicted gap is based on the characteristics of a typical worker in the percentile or when the actual wage gap is computed.

Table A2: Wage ratios and decomposition across the wage distribution

		Log wage gaps			position† of
	Actual	Predicted-1* $= \beta_m^q X_m^q - \beta_f^q X_f^q$	Predicted-2** $= \beta_m^q \overline{X}_m - \beta_f^q \overline{X}_f$	Quantile log explained-1	Wage gap explained-2
					%
10th percentile	0.230	0.182	0.192	51.8	55.6
25th percentile	0.237	0.237	0.210	57.3	55.0
Median	0.226	0.193	0.231	46.0	46.4
75th percentile	0.215	0.209	0.238	54.1	49.8
90th percentile	0.209	0.214	0.258	47.0	40.6

Source: Survey of Labour and Income Dynamics, 1997

The log wage gap is calculated using the coefficients of the quantile regression model of log hourly wages conditioned on a vector of characteristics of the typical worker falling between +/- 5 of the selected percentile.

<sup>\*\*</sup> The log wage gap is calculated using the coefficients of the quantile regression model of log hourly wages conditioned on a vector of the mean characteristics of workers.

<sup>&</sup>lt;sup>†</sup> The characteristics of the typical worker falling between +/- 5 of the selected percentile is used in the calculation of explained-1 while the characteristics of the average worker are used in the calculation of explained-2.

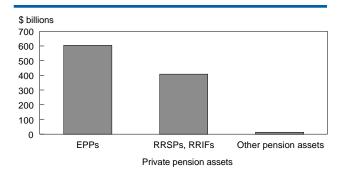
# Private pension savings, 1999

# Karen Maser and Thomas Dufour

As defined by the Survey of Financial Security, private pension assets include individual savings in registered retirement savings plans (RRSPs) and registered retirement income funds (RRIFs), the value of pension plan benefits 'earned' through participation in an employer pension plan (EPP), and other pension savings held in vehicles such as annuities and deferred profitsharing plans. Together, they are referred to as *private* pension savings to indicate that they do not include the value of the income to be received from the Old Age Security/Guaranteed Income Supplement (OAS/GIS) program, and the Canada and Quebec Pension Plans (C/QPP).

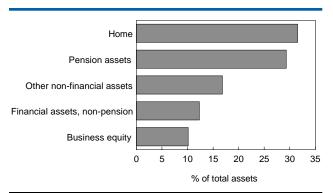
Private pension assets are a major component of the assets of Canadian families, accounting for close to 29% of the value of all assets. Despite the size and importance of these assets, they are still second to the most valuable asset: the home. It too must be

Chart B: Employer pension plans account for the largest amount of private pension assets.



Source: Survey of Financial Security

Chart A: Private pension savings and homes are the major assets of most Canadians.



Source: Survey of Financial Security

considered in any discussion of readiness for retirement, since it can also serve a very important role in providing for a more financially secure future.

The total estimated value of private pension assets is just over \$1 trillion. This money will be essential in providing a large part of the income of seniors. It also, however, plays a very important role in the current financial markets as one of the largest pools of investment capital in the country.

The value of the benefits to be paid from employer pension plans (\$604 billion) is well over half the total amount in private pension savings. This is considerably more than the amount saved in RRSPs and RRIFs (\$408 billion). Close to 60% of family units have RRSPs or RRIFs, with a median value of \$20,000. Although fewer family units have EPP assets (47%), their median value is much larger (\$49,300). The value of employer pension plan benefits includes that of current plan members as well as those receiving this income.

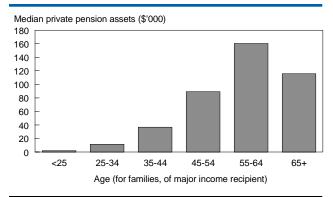
Karen Maser and Thomas Dufour are with the Income Statistics Division. Karen Maser can be reached at (613) 951-0793 or karen.maser@statcan.ca. Thomas Dufour can be reached at (613) 951-2088 or thomas.dufour@statcan.ca.

The amount held by family units as private pension assets increased significantly with the age of the individual or, in the case of families, the person with the highest pre-tax income. Pension assets peaked for family units with a major income recipient between 55 and 64 years of age. These family units, who would have been approaching retirement or just recently retired, had median pension assets of \$160,300. It is not surprising that these families led the way in terms of pension holdings, given that the value of employer pension plan benefits increases with the number of years of membership in the plan. As well, these people had a longer period in which to accumulate RRSP assets.

Senior family units (those in which the person with the highest pre-tax income was 65 years and older) also had significant private pension assets, second only to those with a major income recipient aged 55 to 64. Median pension assets for these family units was about \$115,700. Most of the major income recipients in these families were retired and would already have been drawing on these assets, reducing the amount from a pre-retirement peak.

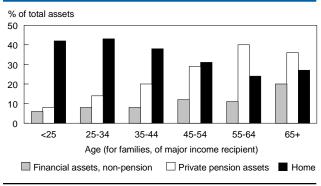
As the age of the major income recipient increased, private pension assets became a much larger proportion of total assets. For family units with a major income recipient aged 55 to 64, pension assets were by far the most significant component of total assets (40%). On the other hand, pension assets represented only about 20% of the assets of family units with a major income recipient between 35 and 44; for these families the home was the most valuable asset (38% of total assets). Financial assets held outside private pension plans were a more important asset for the older age groups—for those family units with a major income recipient 65 years and older, these assets accounted for 20% of the total. These assets are important since they generate income for this age group.

Chart C: Median private pension assets are greatest for those 55 to 64.



Source: Survey of Financial Security

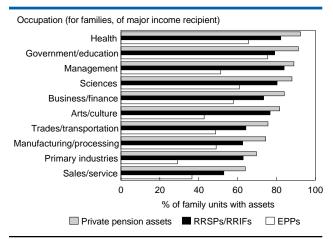
Chart D: Private pensions are the largest asset for those 55 and over.



Families with a major income recipient employed in the public sector (with the government or in education) were not only among the most likely to have private pension assets (91%), they also had the highest median pension value (\$84,400). Almost all public-sector employers offer an employer pension plan.

Family units with a major income recipient whose occupation was classified as "management" also had relatively large private pension assets; the median value was \$74,300. In addition, this group had the highest median RRSP/RRIF holdings (\$35,000). This is likely because they had the highest median after-tax family income (\$56,100) and therefore were in a better position to save.

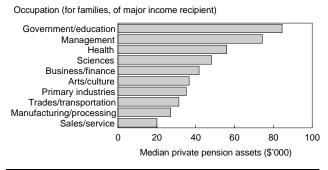
# Chart E: Government, education and health employees are the most likely to have private pension assets.



Source: Survey of Financial Security

Those employed in certain occupations in the private sector were not only less likely to have pension assets, they also had assets with a lower median value. Some of these occupations account for a large proportion of the labour force. For example, 13% of individuals or major income recipients were employed in sales and service. A much lower proportion of the family units in this occupation had pension assets (64%), and the median value of those assets (\$20,000) was lower than for any other occupation. Family units with a major income recipient also had the lowest median after-tax family income, providing them with less opportunity to contribute to RRSPs.

# Chart F: Those employed in government and education have the highest median private pension assets.



Source: Survey of Financial Security

#### Data source

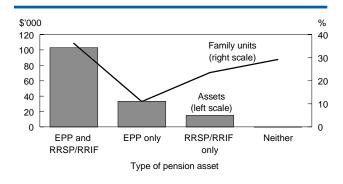
The 1999 Survey of Financial Security, which covered about 16,000 responding households, collected information from May to July 1999 on the assets and debts of families and unattached individuals. It captured the value of all major financial and non-financial assets, as well as money owing on mortgages, vehicles, credit cards, student loans and other debts. The survey was developed with the support of Human Resources Development Canada, Canada Mortgage and Housing Corporation, Industry Canada and the Policy Research Initiative.

This note analyzes Canadians' retirement savings. It makes a number of assumptions such as the age of retirement, income required, assets available to generate that income, and the expected earnings on those assets. These assumptions and the methodology used are described in more detail in Survey of Financial Security: Methodology for estimating employer pension plans benefits (Catalogue no. 13F0026MIE), free on Statistics Canada's Web site (www.statcan.ca).

Families with both employer pension plan assets and RRSP/RRIF assets had significantly higher pension assets than those holding only one or the other. The 36% of families with both types had a median pension value of \$102,900, compared with \$33,300 for those with only EPP assets and \$15,000 for those with only RRSP/RRIF assets. Interestingly, people who belong to EPPs tend to contribute more to their RRSP than those who do not. This is largely related to income. Half of the family units with EPP and RRSP assets had after-tax family incomes of \$40,000 or more, compared with 17% of those family units that had only EPP assets.

Significantly more family units had only RRSP assets than had only EPP assets (24% compared with 11%). This is largely because RRSPs are more widely available. Any person with earned income (largely employment income) could have contributed to an RRSP, while only those working for an employer that provided a pension plan would have EPP assets.

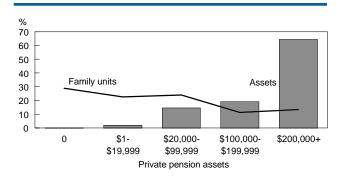
Chart G: Largest median private pension assets held by family units with both EPPs and RRSPs/RRIFs.



Source: Survey of Financial Security

Family units holding only EPP assets had much higher median pension assets than those with only RRSP/RRIF assets. Again, this is because EPPs require that regular contributions be made, while RRSPs do not.

Chart H: About 25% of family units had 84% of private pension assets.



Source: Survey of Financial Security

Private pension assets were concentrated in a relatively small percentage of family units. The 25% of family units with \$100,000 or more in private pension savings held 84% of these assets. About half of these (13%) had at least \$200,000 in pension assets—this

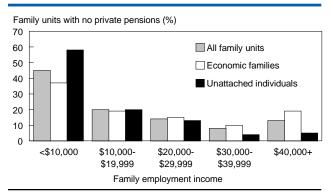
group alone held close to two thirds (64%) of the total. However, almost 29% of all family units did not have any private pension assets in 1999.

Close to half (49%) of the family units in which the major income recipient was between 55 and 64 had at least \$100,000 in private pension assets. This age group also had the lowest percentage of family units with no pension assets (21%). A much smaller percentage (36%) of family units in which the major income recipient was likely to be retired (being 65 or older) had pension assets of \$100,000 or more. Many of these family units would already have been drawing on their pension assets, thereby reducing the amount held.

Notably, 34% of the family units with a major income recipient 65 or older had no pension assets. These families may not necessarily be less well off than in pre-retirement years since their income from government programs (OAS/GIS and C/QPP) may be sufficient to maintain their former standard of living.

The large majority of family units with no private pension assets had lower employment incomes. Considering only those economic family units with a major income recipient between 25 and 64, just over 70% of families of two or more with no pension savings had earnings of less than \$30,000. Approximately 78% of unattached individuals had earnings of under \$20,000. Even though these families and individuals have little saved privately, public plans such as OAS/ GIS and C/QPP will provide them with a minimum income in retirement. This income would replace a substantial portion of their pre-retirement earnings. Using 1998 rates, a single person with no other income would receive from OAS/GIS an annual income of just under \$11,000 at 65; a couple, both 65, would get about \$17,800.

Chart I: Most family units aged 25 to 64 with no private pension assets had earnings of less than \$30,000.

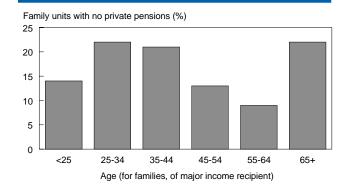


Source: Survey of Financial Security

Most of those with no private pension assets were relatively young. This puts them further from retirement and means they still have a number of years in which to accumulate assets. More than half (57%) of families with no private pension assets had a major income recipient under age 45.

However, the remaining 43% of family units (1.5 million in total) without private pension assets had a major income recipient 45 or older. For these family units, the government-sponsored programs (OAS/GIS and C/QPP) will, or do now, constitute an essential source of income in retirement.

Chart J: The majority of family units with no private pension assets were under 45.



How much income does a family require in retirement? Because certain work-related expenses (for example, contributions to C/QPP, Employment Insurance, and employer pension plans) cease at retirement, it is not necessary to have the same gross income after retirement in order to receive the same net income. Also, because expenditures for consumer goods often decrease while mortgage payments and child-rearing costs are eliminated or reduced, a similar standard of living after retirement is possible with less net income than was earned through employment.

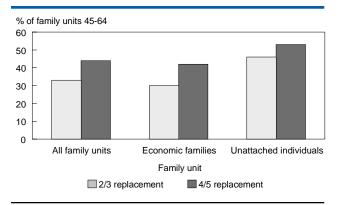
An estimated 33% of family units with a major income recipient aged 45 to 64 may not, given their current asset situation, have saved enough to replace two-thirds of their earnings, or to generate an income in retirement that is likely to be above the low income cutoff (LICO). This increases to 44% if four-fifths of the pre-retirement earnings is to be replaced. The proportion is much higher for unattached individuals because of the number that may not be able to generate an income above the LICO.

The amount of income that a family must replace from private sources increases with their preretirement earnings. It is therefore not surprising that the percentage of family units with high employment incomes (of \$75,000 or more) that might not be able to replace two-thirds of their earnings is relatively high, at 41%. Family units earning at least \$75,000 represent over one-third of all family units that might not have saved enough; their median net worth was \$235,300. By comparison, those with earnings of \$75,000 or more that appear to have sufficient savings had a median net worth of \$628,400.

A lower proportion (just under one-quarter) of those with employment incomes of between \$20,000 and \$40,000 may not be able to replace two-thirds of their earnings. The income this group will receive from public pension programs (OAS/GIS and C/QPP) will help most of them to maintain a similar standard of living in retirement.

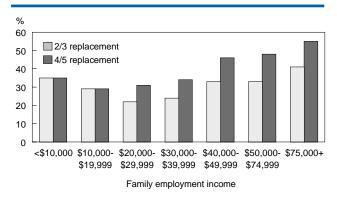
Just over 35% of those with employment incomes under \$10,000 may not have saved enough to generate an income in retirement above the low income cutoff. For many in this group, the result may not be a drop in standard of living but a continuation of one that is restricted.

Chart K: One-third of family units aged 45 to 64 may not have sufficient retirement savings.



Source: Survey of Financial Security

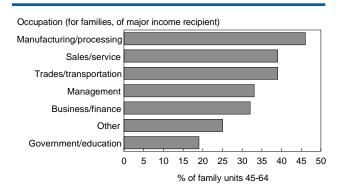
Chart L: Family units with lower and higher employment incomes are the least likely to have saved enough.



Family units with a major income recipient employed in the public sector—with some level of government or in an education-related occupation—were most likely to have saved enough for retirement. A relatively small 19% of this group did not appear to have sufficient assets. This group also had the largest proportion of family units with both employer pension plan assets and RRSPs (63%). Only 9% had no savings in either of these private pension programs.

The family units that may have the most difficulty replacing two-thirds of their earnings in retirement have major income recipients working in occupations associated with processing and manufacturing (46%). About 26% did not have any assets in employer pension plans or RRSPs.

Chart M: Those working in manufacturing and processing are more likely to have difficulty replacing at least two-thirds of their earnings in retirement.

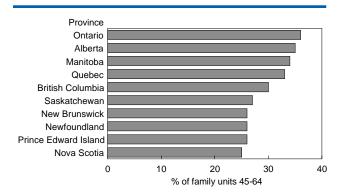


Source: Survey of Financial Security

Alberta and Ontario had the largest proportion of family units aged 45 to 64 that may not have saved enough to replace two-thirds of their earnings in retirement (just over 35%). This is true even though the median net worth of all family units 45 to 64 in Alberta and Ontario was higher than in any other province except British Columbia (\$255,000 in Alberta and \$267,000 in Ontario). However, median employment income for this same population was also among the highest in these provinces and, as was seen earlier, the proportion who may have difficulty replacing their earnings when they retire rises with income.

The Atlantic provinces and Saskatchewan had lower proportions of family units that may not have saved enough. For the most part these provinces have lower median employment income; income from OAS/GIS and the C/QPP will help many family units in these provinces to maintain their standard of living when they retire.

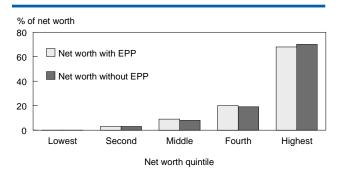
Chart N: The proportion of family units aged 45 to 64 that may not be able to replace at least two-thirds of their earnings in retirement is highest in Alberta and Ontario.



The net worth of Canadians, excluding the value of employer pension plan benefits, is about \$81,000. Including EPP benefits increases net worth by almost 35% to approximately \$109,200. This indicates the importance of considering these retirement savings when examining the financial situation of Canadians. Although EPP benefits are not savings that can be withdrawn and used for other purposes prior to retirement, they nevertheless constitute an essential part of the financial picture of Canadians.

Including the value of employer pension plan benefits in the net worth of Canadians changed the distribution of net worth only very slightly. Because a relatively large proportion of Canadians belong to employer pension plans (47% of family units had at least one member who belonged to an EPP), it might be expected that adding the value of the EPP benefit would result in a more even distribution of net worth. However, using the estimate of net worth including the value of EPP benefits, the wealthiest 20% of family units continued to hold the largest percentage of

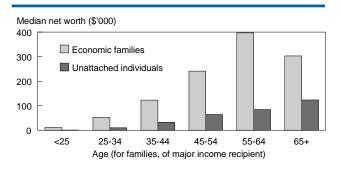
Chart O: Including the value of employer pension plan benefits changed the distritution of the net worth of Canadians only slightly.



Source: Survey of Financial Security

personal wealth: 68%. This is slightly less than the 70% excluding EPP. The proportion of the net worth held by the 40% of Canadians with the lowest personal wealth did not change.

Chart P: Median net worth is highest for families with the major income recipient aged 55 to 64.



Source: Survey of Financial Security

Families and individuals generally acquire their assets over their lifetime, so it is not surprising that net worth increases with age. Median net worth was highest for those families of two or more in which the major

income recipient was 55 to 64 years of age (\$397,000). It dropped for those 65 and over to \$302,800. This is to be expected since many elderly families may be required to use some of their assets to generate or supplement their income.

The net worth of unattached individuals was well below that of economic families for every age group. Although net worth increased with age for the unattached, the median net worth of all age groups under 65 was substantially lower than for those 65 and older. Many of the unattached 65 and older may have spent a large part of their lives as part of a family—their higher net worth may be a reflection of this.

Those under 25, who have had less time to accumulate savings and purchase assets, had the lowest median net worth. Unattached persons under 25 had a median net worth of \$1,000. Families in which the major income recipient was under 25 had a median net worth of \$11,400.

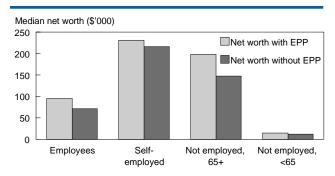
Adding the value of employer pension plan benefits to net worth increased the median net worth of employees by 33%, compared with 7% for the self-employed. Despite this, the median net worth of family units in which the major income recipient was self-employed was still almost two and a half times greater than for an employee (\$231,100 compared with \$95,200). This reflects the effect of business equity on the net worth of the self-employed; for that group business equity represented 31% of total assets, compared with 6% for employees.

The net worth of those not working and at least 65 years of age also increased substantially (to \$198,300) with the addition of the value of employer pension plan benefits. Many in that group were retired and collecting pensions from previous employment. The net worth of those not working and under 65 also increased 22% because of those receiving pensions or having employer pension plan assets from previous employment.

The highest level of education makes a significant difference to the financial situation of the family unit. It is one of the most important determinants of net worth, since it affects both income and occupation. Median net worth for family units in which the individual or major income recipient had not graduated from high school was \$79,600, compared with \$419,600 for those with a professional degree in law, medicine, dentistry, veterinary medicine or optometry.

Adding the value of employer pension plan benefits to net worth had the largest effect on those with doctorates. This group had the largest percentage of family units with EPP assets (71%). Although those with a professional degree in law, medicine, dentistry, veterinary medicine or optometry had the highest median net worth, they were less affected by the addition of the value of EPP benefits since a much smaller percentage of family units in that group belonged to EPPs (37%). Many are self-employed and would therefore not be eligible to participate in an EPP. This group would depend more heavily on RRSPs.

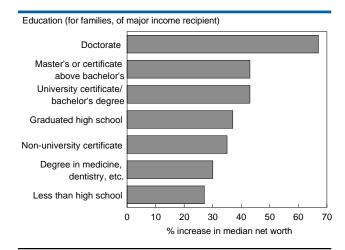
Chart Q: Median net worth of the self-employed is higher than for employees.



Employment status (for families, of major income recipient)

Source: Survey of Financial Security

Chart R: Increase in net worth with addition of the value of EPP is greatest for those with doctorates.



1989-2001

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employer needs Provincial earnings differences	Summer 2000	Liberal arts degrees and the labour market (July 2001)	Autumn 2001

Employment and earnings of	Autumn 2001	The many faces of unemployment	Spring 1996
postsecondary graduates	Samin ~ 2001	Women as main wage-earners	Winter 1995
Overqualified? Recent graduates,	Spring 2001	Families and moonlighting	Summer 1995
employer needs	Smin ~ 2000	Hours of working couples Work and low income	Summer 1995
The school-to-work transition	Spring 2000		Summer 1995
Paying off student loans	Spring 1999	Adults living solo	Winter 1994
Facing the future: Adults who go	Autumn 1997	High income families	Winter 1994
back to school	Summer 1997	Left behind: Lone mothers in the labour market	Summer 1994
After high school	Autumn 1995		Summer 1994
Employment prospects for high school graduates		Spending patterns of couples without children	
Labour market outcomes for university co-op graduates	Autumn 1995	Balancing work and family responsibilities	Spring 1994
Work experience	Summer 1995	Family facts (charts)	Spring 1994
Youths—waiting it out	Spring 1994	Employed parents and the division	Autumn 1993
Labour market outcomes for high	Winter 1993	of housework	
school leavers		Female lone parents in the labour	Spring 1993
School, work and dropping out	Summer 1993	market	
Women in academia—a growing	Spring 1993	Alimony and child support	Summer 1992
minority		The changing profile of dual-earner	Summer 1992
A degree of change	Winter 1992	families	
Juggling school and work	Spring 1992	Marriage, money and retirement	Winter 1991
Lifelong learning: Who goes back	Winter 1991	Family income inequality in the 1980s	Autumn 1991
to school?	*****	Who's looking after the kids? Child	Summer 1991
Overview of literacy skills in Canada	Winter 1990	care arrangements of working mother	
Male-female earnings gap among	Summer 1990	Women's earnings and family incomes	Summer 1991
recent university graduates	0 : 4000	Tracking down discretionary income	Spring 1991
The graduates of '82: Where are they?	Spring 1990	Government transfer payments and family income	Autumn 1990
		Where the money goes: Spending	Autumn 1990
FAMILIES		patterns in Canada and the U.S.	6 1000
r	A	Work and relative poverty	Summer 1990
Low income intensity: urban and rural families (June 2001)	Autumn 2001	Wives as primary breadwinners	Spring 1990
Experiencing low income for	Summer 2001		
several years (March 2001)	0 : 0004	GOVERNMENT TRANSFER PA	YMENTS
Part-time by choice	Spring 2001		
(November 2000)	W. 4000	Low income intensity: urban and rural	Autumn 2001
Working together—self-employed	Winter 1999	families (June 2001)	0.004
couples	1000	Repeat users of employment insurance	Summer 2001
Employment after childbirth	Autumn 1999	(April 2001)	W/ 1000
Family income: 25 years of	Spring 1999	Family income inequality, 1970-1995	Winter 1998
stability and change	W/: 1000	Intergenerational equity in Canada	Autumn 1997
Family income inequality, 1970-1995	Winter 1998	Transfer payments to families	Autumn 1996
Couples working shift  Patierment patterns of working couples	Autumn 1998	with children	Winter 1005
Retirement patterns of working couples	Summer 1998	Men retiring early: How are they doing?	Winter 1995 Summer 1994
Income after separation—people without children	Summer 1996	Who gets UI? Family income inequality in the 1980s	Autumn 1991
Stay-at-home dads	Spring 1998	Dependence on government transfer	Summer 1991
Intergenerational equity in Canada	Autumn 1997	payments, 1971-1989	Julillier 1771
Family income after separation	zididili 1///		Winter 1990
i aminy income arter separation	Summer 1997	Laxes fransfers and regional disparities	
	Summer 1997 Autumn 1996	Taxes, transfers and regional disparities Government transfer payments and	
Transfer payments to families	Summer 1997 Autumn 1996	Government transfer payments and	Autumn 1990

HEALTH		Incomes of seniors	Winter 2000
	2000	Income inequality within provinces	Winter 2000
Long working hours and health	Spring 2000	(November 2000)	
You wear it well: Health of older workers	Autumn 1996	Family income: 25 years of stability and change	Spring 1999
A job to die for	Summer 1996	Income transition upon retirement	Winter 1998
Tired workers	Summer 1995	Family income inequality, 1970-1995	Winter 1998
Perceptions of workplace hazards Defining and measuring employment	Spring 1994 Winter 1993	Income after separation—people without children	Summer 1998
equity Back injuries at work, 1982-1990	A t 1002	Family income after separation	Summer 1997
Under the influence	Autumn 1992 Autumn 1990	Northern earnings and income	Spring 1997
Disabled workers	Winter 1989	Transfer payments to families with children	Autumn 1996
		Dual-pensioner families	Autumn 1996
*************		Men retiring early: How are they doing?	Winter 1995
HIGH TECHNOLOGY		Work and low income	Summer 1995
Warling with computars	Summar 2001	High income families	Winter 1994
Working with computers	Summer 2001	Who gets UI?	Summer 1994 Autumn 1993
(May 2001) Computer programmers	Autumn 1998	Income facts (charts) Investment income of Canadians	Summer 1993
The booming market for programmers	Summer 1998	Facing retirement	Spring 1993
Computers in the workplace	Summer 1997	Alimony and child support	Summer 1992
The future of data dissemination	Summer 1996	Hard at work	Spring 1992
Computers in the workplace	Summer 1991	Family income inequality in the 1980s	Autumn 1991
Measuring Canada's international	Summer 1990	Dependence on government transfer	Summer 1991
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High technology at work	Spring 1990	Women's earnings and family incomes	Summer 1991
<i>.</i>	1 0	Work and relative poverty	Summer 1990
IMMIGRANTS			
IMMIGRANTS		INDUSTRIES	
IMMIGRANTS  Knowledge workers on the move	Summer 2000	INDUSTRIES	
	Summer 2000 Spring 1995	INDUSTRIES Fact-sheet on gambling	Summer 2001
Knowledge workers on the move			Summer 2001 Spring 2000
Knowledge workers on the move Canada's newest workers Defining and measuring employment equity	Spring 1995 Winter 1993	Fact-sheet on gambling Update on gambling Missing work in 1998—industry	
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Are single industry towns diversifying?	Spring 1992	Demography and the labour market	Spring 2001
A look at fishing, mining and		(February 2001)	0.000
wood-based communities	W/inton 1000	Help-wanted Index	Summer 2000
Immigrants in product fabricating The Canadian auto industry,	Winter 1989 Autumn 1989	Labour force participation in the 1990s Working overtime in today's labour	Autumn 1998 Winter 1997
1978-1986	Autumm 1909	market	
		The labour market: Mid-year review	Every Autumn (1989 to 1997)
INTERNATIONAL COMPARIS	ONS	The labour market: Year-end review	Every Spring
INTERNATIONAL COMPARIS	ONS		0 to 1997, 2001)
Recent trends in taxes internationally (January 2001)	Spring 2001	Employment and industrial development in the North	Spring 1997
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by industry—a note		Hiring difficulties in manufacturing	Summer 1995
Gail Cook Johnson speaks out on	Spring 1991	Job-related moves	Winter 1992
human resource issues		Staying put: Job tenure among paid	Winter 1992
Labour force participation: An	Winter 1990	workers	
international comparison	****	Workers on the move: Permanent	Autumn 1992
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The distribution of wealth in Canada	Spring 1990		
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INTERVIEWS		A AMERICA CAY	
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		Gail Cook Johnson speaks out on	Spring 1991
		human resource issues	
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771 111 .1 .1 .1 .C	W/: 4 2004	Training the work force: A challenge	Winter 1990
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September 11 (October 2001) Liberal arts degrees and the labour market (July 2001)	Autumn 2001		

OCCUPATIONS  OCCUPATIONS  Earnings of lawyers Earnings of physicians  Winter 1999 Work patterns of truck drivers Computer programmers The booming market for programmers Work patterns of managers Work patterns of managers Winter 1999 Work absence rates, 1995 Women in non-traditional occupations Imployer-supported training—it varies by occupation Unemployment—occupation makes a difference Trading places: Men and women in non-traditional occupations, 1971-86  PENSIONS  Who contributes to RRSPs? Autumn 1996 A re-examination (July 2001) Pension coverage and retirement savings Incomes of younger retired women: the past 30 years (December 2000) In for the long term: pension plans offered by employers (October 2000) In for the long term: pension plans offered by employers (October 2000) In for the long term: pension plans offered senior savings and RPPs The RRSP Home Buyers' Plan RRSPs—mew rules, new growth C/QPP costs and private pensions Spring 1993 Spring 1993 Spring 1994 Spring 1999 Earnings of lawyers Spring 2000 Savings for retirement Summer 1999 Winter 1999 Winter 1999 Winter 1999 Women and RRSP ocntributions And padate  An interview with Laurence E. Coward RRSPs—new rules, new growth Spring 1999 Earing retirement Spring 1993 And payouts  Employer-sponsored pension Winter 1992 RRSPS—not just for retirement Winter 1992 Women and RRSP Winter 1999 Women and RRSP Winter 1999 Women and RRSP Autumn 1996 Autumn 1996 Women and RRSP Retirement attitudes, plans and behaviour The pension carrot: Incentives to early retirement Dependence on government transfer payments and family income The performance of trusteed pension Traditional occupations with the past 30 years (December 2000) In for the long term: pension plans offered by employers (October 2000) In for the long term: pension plans offered by employers (October 2000) In for the long term: pension plans offered by employers (October 2000) RRSPs in the 1990s Spring 2000 Savings for retirement Winter 1990  Employer-sponsored pension Winter 1990 Women and RRSP Women anp	MEN		RRSP rollovers	Winter 1996
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Perspectives